
Supplementary Material for: Differentiable Causal Backdoor Discovery

A Proofs

Lemma 1. If $W \perp\!\!\!\perp Y \mid Z^* \cup \{X\}$, then there exists some scalar $\phi_X(Z^*)$ such that $W \perp\!\!\!\perp Y \mid \{\phi_X(Z^*), X\}$.

Proof. We will discuss the case for binary Y , as the proof for linear-Gaussian models follows the same idea. Let the structural equation for Y be given by $f_y(x, \pi_Z, \pi_U)$, where π_Z and π_U are the observed and unobserved parents of Y in the corresponding causal graph. The conditional distribution of Y is given by

$$\begin{aligned} p(y \mid x, \mathbf{z}^*) &= \\ p(f_y(x, \pi_Z, \pi_U) = 1 \mid x, \mathbf{z}^*)^y &\times \\ (1 - p(f_y(x, \pi_Z, \pi_U) = 1 \mid x, \mathbf{z}^*))^{1-y}. & \end{aligned}$$

By assumption, $f_y(\cdot)$ is functionally independent of W . Now we just have to show that the random variable $f_y(x, \pi_Z, \pi_U)$ is conditionally independent of W given X and Z^* . Since $W \perp\!\!\!\perp Y \mid Z^* \cup \{X\}$, it cannot be the case that W and $\pi_{Z^*, X}$, the parents of Y not in $Z^* \cup \{X\}$, are conditionally dependent given $Z^* \cup \{X\}$. We define $\phi_X(\mathbf{z}^*)$ as $p(f_y(x, \pi_Z, \pi_U) = 1 \mid x, \mathbf{z}^*)$ for each possible realization of X . Given X , we can fully reconstruct from $\phi_X(\mathbf{z}^*)$ a conditional distribution of Y that makes information about W irrelevant. \square

Theorem 1. If $W \not\perp\!\!\!\perp Y \mid Z^* \cup \{X\}$, and

$$\begin{aligned} \sum_{\mathbf{z}^* \in \Phi_{x\mathbf{z}^*}^f} p(y \mid w, x, \mathbf{z}^*) \frac{p(\mathbf{z}^* \mid w, x)}{\Pr(Z^* \in \Phi_{x\mathbf{z}^*}^f \mid w, x)} &\neq \\ \sum_{\mathbf{z}^* \in \Phi_{x\mathbf{z}^*}^f} p(y \mid x, \mathbf{z}^*) \frac{p(\mathbf{z}^* \mid x)}{\Pr(Z^* \in \Phi_{x\mathbf{z}^*}^f \mid x)}, & \end{aligned} \quad (6)$$

for some value f in the range of $\phi_x(\cdot)$, then $W \not\perp\!\!\!\perp Y \mid \{\phi_X(Z^*), X\}$.

Proof. Assume, contrary to the hypothesis, that $W \perp\!\!\!\perp Y \mid \{\phi_X(Z^*), X\}$. Then

$$\begin{aligned} p(y \mid w, x, \phi_x(Z^*) = f) &= p(y \mid x, \phi_x(Z^*) = f) \Rightarrow \\ \sum_{\mathbf{z}^*} p(y \mid w, x, \phi_x(Z^*) = f, \mathbf{z}^*) p(\mathbf{z}^* \mid w, x, \phi_x(Z^*) = f) &= \\ \sum_{\mathbf{z}^*} p(y \mid x, \phi_x(Z^*) = f, \mathbf{z}^*) p(\mathbf{z}^* \mid x, \phi_x(Z^*) = f) &\Rightarrow \\ \sum_{\mathbf{z}^*} p(y \mid w, x, \mathbf{z}^*) p(\mathbf{z}^* \mid w, x, \phi_x(Z^*) = f) &= \\ \sum_{\mathbf{z}^*} p(y \mid x, \mathbf{z}^*) p(\mathbf{z}^* \mid x, \phi_x(Z^*) = f) &\Rightarrow \\ \sum_{\mathbf{z}^* \in \Phi_{x\mathbf{z}^*}^f} p(y \mid w, x, \mathbf{z}^*) \frac{p(\mathbf{z}^* \mid w, x)}{\Pr(Z^* \in \Phi_{x\mathbf{z}^*}^f \mid w, x)} &= \\ \sum_{\mathbf{z}^* \in \Phi_{x\mathbf{z}^*}^f} p(y \mid x, \mathbf{z}^*) \frac{p(\mathbf{z}^* \mid x)}{\Pr(Z^* \in \Phi_{x\mathbf{z}^*}^f \mid x)}, & \end{aligned}$$

which contradicts the hypothesis. \square