

# Sequential no-Substitution $k$ -Median-Clustering

Tom Hess and Sivan Sabato

## A Bernstein and empirical Bernstein inequalities

*Proof of Lemma 3.2.* We use the Empirical Bernstein inequality of (Maurer and Pontil, 2009). This inequality states that for  $\hat{\sigma}^2 := \frac{1}{2n(n-1)} \sum_{i,j \in [n], i \neq j} (Y_i - Y_j)^2$ , with a probability at least  $1 - \delta$ , we have

$$\hat{\mu} - \mu \leq \frac{7 \ln(\frac{2}{\delta})}{3(n-1)} + \sqrt{\frac{2\hat{\sigma}^2 \ln(\frac{2}{\delta})}{n}}.$$

We have  $\hat{\sigma}^2 \leq \frac{n}{2(n-1)} \frac{1}{n^2} \sum_{i,j \in [n]} (Y_i - Y_j)^2 = \frac{n}{2(n-1)} \mathbb{E}[(Y - Y')^2]$ , where  $Y, Y'$  are drawn independently and uniformly from the fixed sample  $Y_1, \dots, Y_n$ . Since  $\mathbb{E}[(Y - Y')^2] \leq 2\mathbb{E}[Y^2]$ , and  $Y \in [0, 1]$ , we have  $\hat{\sigma}^2 \leq \frac{n}{n-1} \mathbb{E}[Y] \equiv \frac{n}{n-1} \hat{\mu}$ . Therefore,

$$\hat{\mu} - \mu \leq \frac{7 \ln(\frac{2}{\delta})}{3(n-1)} + \sqrt{\frac{2\hat{\mu} \ln(\frac{2}{\delta})}{n-1}}.$$

If  $\hat{\mu} = a \ln(\frac{2}{\delta})/(n-1)$  for  $a \geq 16$ , then the RHS is at most

$$(7/3 + \sqrt{2a}) \ln(\frac{2}{\delta})/(n-1) \leq a/2 \cdot \ln(\frac{2}{\delta})/(n-1) \leq \hat{\mu}/2.$$

□

*Proof of Lemma 4.3.* Let  $\sigma^2 = \text{Var}[Y_i]$ . By Bernstein's inequality (Hoeffding, 1963) (See, e.g., Maurer and Pontil 2009 for the formulation below),

$$\mu - \hat{\mu} \leq \frac{\ln(\frac{1}{\delta})}{3n} + \sqrt{\frac{2\sigma^2 \ln(\frac{1}{\delta})}{n}}.$$

Since  $Y_i$  are supported on  $[0, 1]$ , we have  $\sigma^2 \leq \mu$ . Since  $\mu = a \ln(\frac{1}{\delta})/n$  for  $a \geq 10$ , we have that the RHS is equal to  $(1/3 + \sqrt{2a}) \ln(\frac{1}{\delta})/n \leq a/2 \cdot \ln(\frac{1}{\delta})/n \leq \mu/2$ . The statement of the lemma follows. □

## B Tightness of multiplicative factor of SKM

*Proof of Theorem 3.6.* We define a weighted undirected graph  $G = (V, E, W)$ , and let  $(\mathcal{X}, \rho)$  be a metric space such that  $\mathcal{X} = V$  and  $\rho(u, v)$  is the length of the shortest path in the graph between  $u$  and  $v$ .  $G$ , which

is illustrated in Figure 2, is formally defined as follows. The set of nodes is  $V := U \cup Y \cup \{o, v\}$ , where  $U := [0, 1]$  and  $Y := [3, 4]$ . The set of edges is

$$E := \{\{u, o\} \mid u \in U\} \cup \{\{v, y\} \mid y \in Y\} \cup \{\{o, v\}\}.$$

Denote  $m_1 := m/2$ , and let  $\eta := 1/(4m_1)$ . The weight function  $W$  assigns a weight of 1 to all edges except for those that have a node in  $Y$  as an endpoint, which are assigned a weight of  $2 - \eta$ .

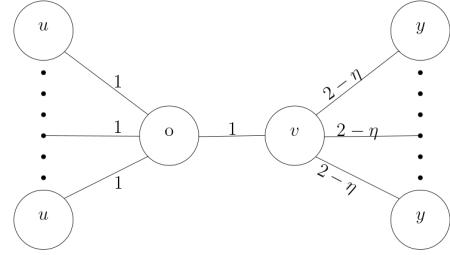


Figure 2: Illustration of the graph  $G$  which defines the metric space.

Define the distribution  $P$  over  $\mathcal{X}$  such that  $P(o) = 0$ ,  $P(v) = 1/m_1$ ,  $P(Y) = 2q$ , with a uniform conditional distribution over  $Y$ . Lastly,  $P(U) = 1 - 2q - \frac{1}{m_1}$ , with a uniform conditional distribution over  $U$ . Note that the latter is positive for a large enough  $m_1$ , since  $q(m_1) \rightarrow 0$ .

Let  $S \sim P^m$  be the i.i.d. sample used as an input sequence to SKM, and set  $k = 1$ . Let  $S_1$  be the sample observed in the first phase of SKM, of size  $m_1$ . Define the following events:

1.  $E_1 := \{o \notin S_1\}$ .
2.  $E_2 := \{v \text{ appears in } S_1\}$ .
3.  $E_3 := \{\text{at least } qm_1 \text{ of the samples in } S_1 \text{ are from } Y\}$ .

First, observe that all these events occur together with a positive probability, as follows.  $\mathbb{P}[E_1] = 1$  since  $P(o) = 0$ . For  $E_2$ , we have

$$\mathbb{P}[E_2] > 1 - (1 - \frac{1}{m_1})^{m_1} > \frac{1}{2}.$$

For  $E_3$ , note that the probability mass of  $Y$  is  $2q$ , Apply Lemma 4.3 with  $\mu = 2q$ ,  $n = m_1$  and a confidence

value of  $1/4$ . By the assumption of the theorem, for sufficiently small  $\delta$ , we have  $q \geq 5 \log(4)/m_1$ . Therefore, Lemma 4.3 implies that  $P[E_3] \geq 3/4$ . It follows that  $\mathbb{P}[E_1 \wedge E_2 \wedge E_3] \geq 1/4$ .

Now, assume that all the events above hold. By  $E_1$ ,  $o$  does not appear in  $S_1$ , and by  $E_2$ ,  $v$  appears in  $S_1$ . We show that out of the points in  $S_1$ , the 1-clustering  $\{v\}$  has the best empirical risk. The only other options in  $S_1$  are centers from  $Y$  or from  $U$ . For a center  $u \in U$  from  $S_1$ , note that with a probability 1, it does not have additional copies in  $S_1$ . Its distance from all other  $u' \in U$  is the same as that of  $v$ , while its distance from points in  $Y$  and from  $v$  is larger. Thus,  $R(S_1, \{u\}) > R(S_1, \{v\})$ . For a center  $y \in Y$ , it too does not have additional copies in  $S_1$ . Its distance to all other points is larger than that of  $v$ . Thus,  $R(S_1, \{y\}) > R(S_1, \{v\})$ . Therefore,  $v$  has the best empirical risk on  $S_1$ . Thus,  $\mathcal{A}(S_1)$  returns the 1-clustering  $\{v\}$ .

By  $E_3$ , the number of instances of vertices from  $Y$  is at least  $qm_1$ . Since the points in  $Y$  are the closest to  $v$  in  $S_1$ , we have  $y' := \text{qp}_{S_1}(v, q) \in Y$ . Therefore,  $\text{qball}(v, y') = \{v\} \cup Y$ . It follows that **SKM** selects as a center the first element from  $\{v\} \cup Y$  that it observes in the second phase. With a probability  $\frac{2q}{2q + \frac{1}{m_1}}$ , the first element that **SKM** observes from  $\{v\} \cup Y$  is in  $Y$ . Since  $q \geq 1/m_1$ , this probability is at least  $2/3$ . Thus, the output center of **SKM** is from  $Y$  with a constant probability.

However, the risk of this clustering is large:

$$\begin{aligned} R(P, \{y\}) \\ = (4 - \eta)(1 - 2q - \frac{1}{m_1}) + (2 - \eta)\frac{1}{m_1} + (4 - \eta)2q. \end{aligned}$$

For large  $m$ , we have  $m_1 \rightarrow \infty$ . In addition,  $q, \eta \rightarrow 0$ . Hence,  $R(P, \{y\}) \rightarrow 4$ . In contrast, the risk using  $o$  as a center is small:

$$R(P, \{o\}) = (1 - 2q - \frac{1}{m_1}) + \frac{1}{m_1} + (3 - \eta)2q.$$

This approaches 1 for large  $m$ . Therefore, for  $m \rightarrow \infty$ ,  $R(P, \{y\})/R(P, \{o\}) \rightarrow 4$ . Since  $\{y\}$  is the output of **SKM** with a constant probability, the multiplicative factor obtained by **SKM** cannot be smaller than  $4 = 2\beta$  in this case.  $\square$

## C Full results of experiments

The results of the experiments for large stream sizes with the  $k$ -medoids as the black box are reported in Figure 3. The results for the BIRCH black-box are reported in Figure 4 and in Figure 5. For the  $k$ -medoids black box, the risk ratios for large stream sizes are

in the following ranges: **MNIST** 1.02 – 1.04, **Covtype** 1.04 – 1.08, **Census** 1 – 1.04. For the BIRCH black box, the risk ratios for large stream sizes are in the following ranges: **MNIST** 1.03 – 1.04, **Covtype** 1.05 – 1.1, **Census** 1 – 1.02. Thus, the risk ratio converges to a ratio very close to 1.

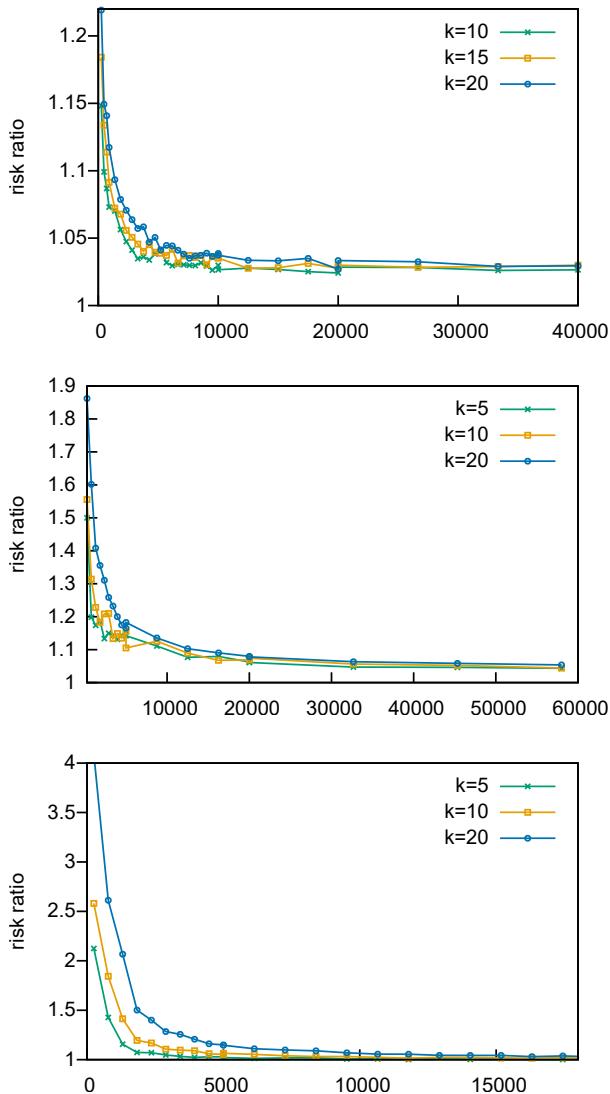


Figure 3: Risk ratio between SKM with  $k$ -medoids and offline  $k$ -medoids for large stream sizes, as a function of the stream size, for various values of  $k$ . Top to bottom: MNIST, Covertype, Census.

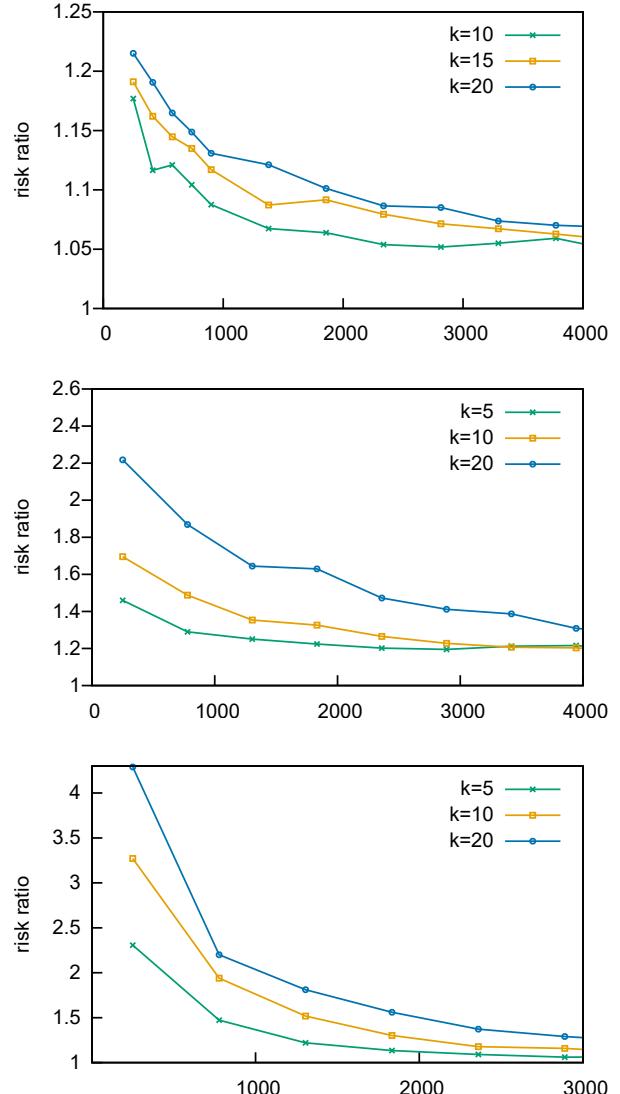


Figure 4: Risk ratio between SKM with BIRCH and offline BIRCH as a function of the stream size, for various values of  $k$ . Top to bottom: MNIST, Covertype, Census.

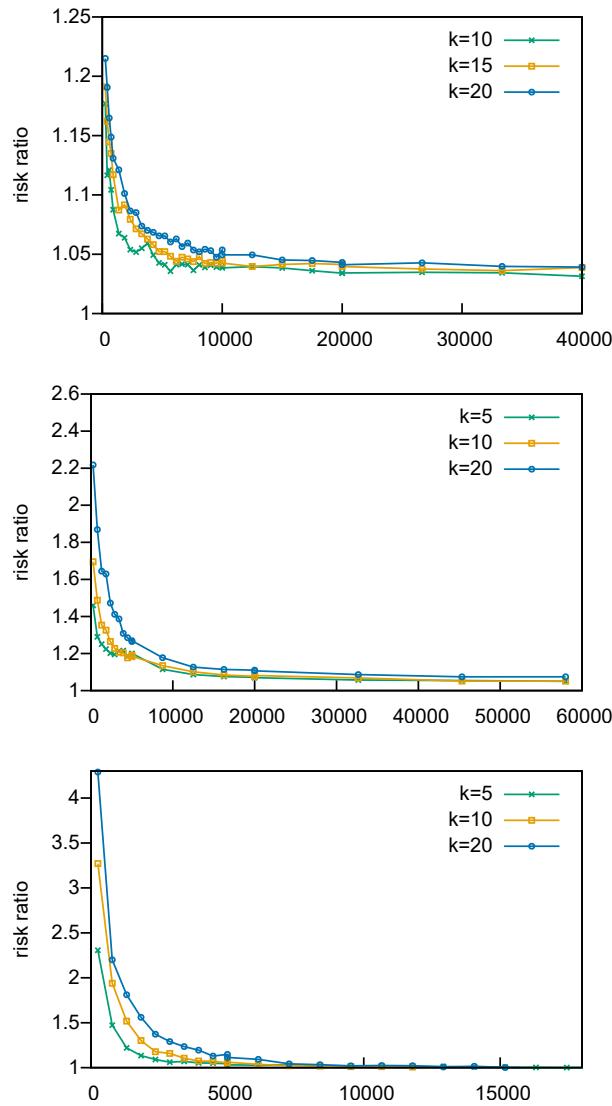


Figure 5: Risk ratio between SKM with BIRCH and offline BIRCH for large stream sizes, as a function of the stream size, for various values of  $k$ . Top to bottom: MNIST, Covertype, Census.