

## A Proving lemma 5, concentration result

**Lemma 10.** *Let  $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, \Pr)$  be a filtered space, with  $\mathcal{F}_0$  the trivial sigma algebra. Let  $(x_t)_{t \geq 1}$  be a previsible sequence  $x_t: \Omega \mapsto \mathbb{R}^d$  and let  $(\epsilon_t)_{t \geq 1}$  with  $\epsilon_t: \Omega \mapsto \mathbb{R}$  be a sequence of random variables adapted to the filtration, with  $\epsilon_t$  1-subGaussian conditionally on  $\mathcal{F}_{t-1}$  for all  $t$ . Let  $(N_t)_{t \geq 1}$  be a non-decreasing sequence of integers. Let  $(\mathcal{A}_t)_{t \geq 0}$  be a sequence of random sets  $\mathcal{A}_t: \Omega \mapsto 2^{2^{\mathcal{X}}}$ , such that  $\mathcal{A}_0$  is  $\mathcal{F}_0$ -measurable,  $(\mathcal{A}_t)_{t \geq 1}$  previsible and  $1 \leq |\mathcal{A}_t| \leq N_t$  almost surely for all  $t \geq 0$ . Let  $k: \mathcal{X} \times \mathcal{X} \mapsto \mathbb{R}$  be a symmetric, positive-semidefinite kernel. Then for any given  $\delta \in (0, 1)$  and  $\eta > 0$ , for all  $t \geq 0$  and all  $A \in \mathcal{A}_t$  we have*

$$\|\epsilon_{1:t}^A\|_{(I + (K_t^A + \eta I)^{-1})^{-1}} \leq 2 \log \left( \det(K_t^A + I + \eta I)^{\frac{1}{2}} N_t / \delta \right),$$

with probability  $1 - \delta$ , where  $\epsilon_{1:t}^A$  for the random vector that is the concatenation of  $(\epsilon_z: x_z \in A)_{z=1}^t$ .

**Proof.** For a function  $g: \mathcal{X} \mapsto \mathbb{R}$  and a sequence of real numbers  $(a_t)_{t \geq 1}$ , define

$$\Delta_t^{g,n} = \exp \left\{ (g(x_t) + a_t)\epsilon_t - \frac{1}{2}(g(x_t) + a_t)^2 \right\},$$

with  $\Delta_0^{g,n}$  defined as equal to 1 almost surely. Then  $\Delta_t^{g,n}$  is  $\mathcal{F}_t$  measurable for all  $t \geq 0$ . By the conditional subGaussianity of  $\epsilon_t$ , we have that  $\mathbb{E}[\Delta_t^{g,n} | \mathcal{F}_{t-1}] \leq 1$  for all  $t \geq 0$  almost surely. For a set  $A \in 2^{\mathcal{X}}$ , define

$$\mathcal{M}_t^{g,n}(A) = \Delta_0^{g,n} \prod_{z=1}^t (\Delta_z^{g,n})^{1_{\{x_z \in A\}}}.$$

Then, for any  $A \in 2^{\mathcal{X}}$  and all  $t \geq 1$ ,  $\mathbb{E}[\mathcal{M}_t^{g,n}(A) | \mathcal{F}_{t-1}] \leq \mathcal{M}_{t-1}^{g,n}(A)$  and  $\mathbb{E}[\mathcal{M}_t^{g,n}(A)] \leq 1$ .

Let  $\zeta = (\zeta_t)_{t \geq 1}$  be a sequence of independent and identically distributed Gaussian random variables with mean 0 and variance  $\eta > 0$ , independent of  $\mathcal{F}_\infty = \bigcup_{t \geq 0} \mathcal{F}_t$ . Let  $h$  be a random real valued function on  $A$  distributed according to the Gaussian process measure  $\mathcal{GP}(0, k|_A)$ , where  $k|_A$  is the restriction of  $k$  to  $A$ . Define

$$M_t^A = \mathbb{E}[\mathcal{M}_t^{h,\zeta}(A) | \mathcal{F}_\infty].$$

Then  $M_t^A$  is itself a non-negative supermartingale bounded in expectation by 1. Define  $\widetilde{M}_t^A = M_t^A / N_t$ . Since  $N_t \geq 1$  for all  $t \geq 0$  and is non-decreasing,  $\widetilde{M}_t^A$  is a non-negative supermartingale bounded in expectation by  $1/N_t$ .

For  $A \in \mathcal{B}(\mathcal{X})$ , let  $B_t^A = \{\omega: \widetilde{M}_t^A > 1/\delta\}$  and  $B_t = \bigcup_{A \in \mathcal{A}_t} B_t^A$ . Define the stopping time  $\tau(\omega) = \inf\{t: \omega \in B_t\}$ . Then

$$\Pr[B_\tau^A | \mathcal{F}_{\tau-1}] \leq \delta \mathbb{E}[\widetilde{M}_\tau^A | \mathcal{F}_{\tau-1}] = \delta \mathbb{E}[M_\tau^A | \mathcal{F}_{\tau-1}] / N_\tau \leq \delta / N_\tau M_{\tau-1}^A \quad \text{a.s.}$$

We now examine the probability of  $B_\tau$ . We have

$$\Pr[B_\tau] = \mathbb{E}[\Pr[B_\tau | \mathcal{F}_{\tau-1}]] \leq \sum_{A \in \mathcal{B}(\mathcal{X})} \mathbb{E}[1\{A \in \mathcal{A}_\tau\} \Pr[B_\tau^A | \mathcal{F}_{\tau-1}]] \leq \delta / N_\tau \sum_{A \in \mathcal{B}(\mathcal{X})} \mathbb{E}[1\{A \in \mathcal{A}_\tau\} M_{\tau-1}^A].$$

The final expectation is complicated by the fact that the event  $\{A \in \mathcal{A}_t\}$  is not independent of  $M_{t-1}^A$ . However,

$$\{A \in \mathcal{A}_t\} \subset \{A \in \mathcal{Z}: \mathcal{Z} \subset \mathcal{B}(\mathcal{X}), |\mathcal{Z}| \leq N_t\}.$$

The latter event holds with probability 1 for all  $t \geq 1$ , and is therefore independent of  $M_t^A$ . This gives,

$$\Pr[B_\tau] \leq \delta / N_\tau \sum_{A \in \mathcal{B}(\mathcal{X})} \mathbb{E}[1\{A \in \mathcal{A}_\tau\} M_{\tau-1}^A] \leq \delta / N_\tau \sum_{A \in \mathcal{B}(\mathcal{X})} \mathbb{E}[1\{A \in \mathcal{Z}: |\mathcal{Z}| \leq N_t\} M_{\tau-1}^A] \quad (5)$$

$$= \delta / N_\tau \mathbb{E}[M_{\tau-1}^A] \sum_{A \in \mathcal{B}(\mathcal{X})} \mathbb{E}[1\{A \in \mathcal{Z}: |\mathcal{Z}| \leq N_t\}] \leq \delta, \quad (6)$$

and consequently

$$\Pr[\bigcup_{t \geq 0} B_t] = \Pr[\tau < \infty] = \Pr[B_\tau, \tau < \infty] \leq \Pr[B_\tau] \leq \delta. \quad (7)$$

Finally, by comparing with the proof of Theorem 1 in Chowdhury and Gopalan (2017), it can be verified that

$$M_t^A = \det(K_t^A + I + \eta I)^{-\frac{1}{2}} \exp \left\{ \frac{1}{2} \|\epsilon_{1:t}^A\|_{(I + (K_t^A + \eta I)^{-1})^{-1}} \right\}.$$

The statement of the lemma follows from using this expression with equation (7), and noting that logarithms preserve order.  $\square$

**Proof of lemma 5.** To prove lemma 5, first since  $|\mathcal{A}_t| \leq |\tilde{\mathcal{A}}_t| \leq \tilde{N}_t$ , we can use  $\tilde{N}_t$  from lemma 6 as the bound  $N_t$  required for lemma 10. Then the proof of lemma 5 follows the proof of theorem 2 in Chowdhury and Gopalan (2017), with our concentration inequality, lemma 10, used instead of their theorem 1.  $\square$