

# On the optimality of kernels for high dimensional clustering - supplementary

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## 1 Notation and Preliminaries

For any  $d \in \mathbb{N}$ , Let  $g : \{(i, d)\}_{i \in [m], d \in [p]} \rightarrow [mp]$  be an injective mapping. For any  $(i, d)$ , for ease of notation, we simply refer to  $g(i, d)$  as  $id$  when it occurs as an index. For any  $i \in [m]$ ,  $d \in [p]$ ,  $x_{id}$  refers to the  $d^{\text{th}}$  component of  $x_i$ . For any two tensors  $x, y$   $x \otimes y$  refers to the outer product between  $x, y$ .

## 2 Definitions

### 2.1 $\alpha$ - sub-Exponential random variables

**Definition 1** ( $\alpha$ - sub-Exponential random variables (Götze, Sambale, and Sinulis 2019)). A centered random variable  $X$  is said to be  $\alpha$ - sub-Exponential if there exist two constants  $c, C$  and some  $\alpha > 0$  such that for all  $t \geq 0$ ,

$$\Pr(|X| \geq t) \leq c \exp\left(-\frac{t^\alpha}{C}\right)$$

The corresponding  $\alpha$ - sub-Exponential norm of  $X$  is given by:

$$\|X\|_{\psi_\alpha} = \inf \left\{ t > 0 : \mathbb{E} \exp\left(\frac{|X|^\alpha}{t^\alpha}\right) \leq 2 \right\}$$

$\alpha$ - sub-Exponential random variables with  $\alpha = 2$  are referred to as sub-Gaussian random variables. Random variables with  $\alpha = 1$  sub-Exponential decay are referred to simply as sub-Exponential random variables.

**Definition 2** (Tensor norms (Götze, Sambale, and Sinulis 2019)). For any  $d^{\text{th}}$  order, symmetric tensor  $A \in \mathbb{R}^{n^d}$ , let  $\mathcal{J} = \{J_1, J_2, \dots, J_k\}$  be any partition of  $[d]$ . Then for any  $x = x^1 \otimes x^2 \otimes \dots \otimes x^k$ , where  $x^i \in \mathbb{R}^{n^{J_i}}$ :

$$\|A\|_{\mathcal{J}} := \sup \left\{ \sum_{i_1, \dots, i_d} a_{i_1 \dots i_d} \prod_{j=1}^k x_{i_{J_j}}^j : \|x^j\|_2 \leq 1 \right\}. \quad (1)$$

### 2.1.1 Properties of sub-Gaussian random variables:

**Proposition 1 (Sums of sub-Gaussian random variables (Vershynin 2018)).**

Let  $\{X_1, X_2, \dots, X_m\}$  be  $m$  independent, centered, sub-Gaussian random variables. Then  $\sum_{i \in [m]} X_i$  is a sub-Gaussian random variable and,

$$\left\| \sum_{i \in [m]} X_i \right\|_{\psi_2}^2 \leq C \sum_{i \in [m]} \|X_i\|_{\psi_2}^2.$$

**Proposition 2 (Products of sub-Gaussian random variables (Vershynin 2018)).** Let  $X_1$  and  $X_2$  be sub-Gaussian random variables. Then  $X_1 \cdot X_2$  is a sub-Exponential random variable and,

$$\|X_1 \cdot X_2\|_{\psi_1} \leq \|X_1\|_{\psi_2} \cdot \|X_2\|_{\psi_2}.$$

**Proposition 3 (Squares of sub-Gaussian random variables (Vershynin 2018)).** Let  $X$  be sub-Gaussian random variables. Then  $X_1^2$  is a sub-Exponential random variable and,

$$\|X_1^2\|_{\psi_1} = \|X_1\|_{\psi_2}^2.$$

## 3 Useful concentration results

**Proposition 4 (Bernstein's inequality (Vershynin 2018)).** Let  $\{X_1, X_2, \dots, X_m\}$  be a set of independent, centered, sub-Exponential random variables. Then for any  $t > 0$ , we have:

$$\Pr \left( \left| \sum_{i \in [m]} X_i \right| \geq t \right) \leq 2 \exp \left( -C \min \left( \frac{t^2}{\sum_{i \in [m]} \|X_i\|_{\psi_1}^2}, \frac{t}{\max_{i \in [m]} \|X_i\|_{\psi_1}} \right) \right)$$

for some fixed constant  $C > 0$ .

**Proposition 5 (Tail bounds for chi-squared distributions (Birgé 2001)).**

The following lower and upper tail bounds hold for non-central chi-squared distributions : For any  $t > 0$ ,

$$\Pr \left( \chi_d^2(\mu^2) < d + \mu^2 - 2\sqrt{(d + 2\mu^2)t} \right) < \exp(-t) \quad (2)$$

$$\Pr \left( \chi_d^2(\mu^2) > d + \mu^2 + 2\sqrt{(d + 2\mu^2)t} + 2t \right) < \exp(-t) \quad (3)$$

**Proposition 6 (Polynomials of  $\alpha$ - sub-Exponentials (Götze, Sambale, and Sinulis 2019)).** Let  $X_1, \dots, X_n$  be a set of independent random variables satisfying  $\|X_i\|_{\psi_2} \leq b$  for some  $b > 0$ . Let  $f: \mathbb{R}^n \rightarrow \mathbb{R}$  be a polynomial of total degree  $D \in \mathbb{N}$ . Then, for any  $t > 0$ ,

$$\mathbb{P}(|f(X) - \mathbb{E}f(X)| \geq t) \leq 2 \exp \left( -\frac{1}{C_D} \min_{1 \leq s \leq D} \min_{\mathcal{J} \in \mathcal{P}_s} \left( \frac{t}{b^s \|\mathbb{E}f^{(s)}(X)\|_{\mathcal{J}}} \right)^{\frac{2}{|s|}} \right).$$

where, for any for any  $s \in D$ ,  $f^{(s)}$  denotes the symmetric  $s^{\text{th}}$  order tensor of its  $s^{\text{th}}$  order partial derivatives and  $\mathcal{P}_s$  denotes the set of all possible partitions of  $[s]$ .

## 4 Other useful results

**Proposition 7 (Grothendieck's inequality (Grothendieck 1956)).** For any matrix  $A \in \mathbb{R}^{m \times m}$ ,

$$\sup_{\substack{X \succeq 0 \\ \text{diag}(\bar{X}) \leq 1}} |\langle X, A \rangle| \leq K_G \|A\|_{\infty \rightarrow 1}.$$

where  $K_G \approx 1.783$  is the Grothendieck's constant.

## 5 Proofs of lemmas

By an application of Bernstein's inequality for sub-exponential random variables, followed by an union bound over all  $s, s' \in [k]$ , it can be verified that, with high probability,

$$\min_{s \in [k]} \|\mu_s\|^2 = p + O(\sqrt{p \log p}); \quad \min_{s \neq s' \in [k]} \langle \mu_s, \mu_{s'} \rangle = \frac{-p}{k-1} + O(\sqrt{p \log p}) \quad (4)$$

**Proof of Lemma 9.** Let  $X = \{x_{id}\}_{i \in [m], d \in [p]}$  and let  $f(X) =$

$$\sum_{i \neq j \in [m]} y_i z_j (\langle x_i, x_j \rangle^2 - \frac{\rho^2}{p^2} \langle \mu_i, \mu_j \rangle^2 - p) + \sum_{i \in [m]} y_i z_i (\|x_i\|^2 - (p + \frac{\rho}{p} \|\mu_i\|^2)^2 - p).$$

Since  $f$  is a  $4^{th}$  order polynomial in  $X$ , from Proposition 6, for any  $t > 0$ :

$$\mathbb{P}(|f(X) - \mathbb{E}f(X)| \geq t) \leq 2 \exp\left(-\frac{1}{C_4} \min_{1 \leq s \leq 4} \min_{\mathcal{J} \in P_s} \left(\frac{t}{b^s \|\mathbb{E}f^{(s)}(X)\|_{\mathcal{J}}}\right)^{\frac{2}{|\mathcal{J}|}}\right).$$

We have  $\mathbb{E}f(X) = C_f(m^2\rho + m)$  for some constant  $C_f > 0$ .

We need the following tensors and their corresponding tensor norms to establish the results of Lemma 9 via an application of Proposition 6.

- For each  $s \in [4]$ ,  $s^{th}$  order tensors  $A_s$  of expectations of  $s^{th}$  order derivatives of  $f$  with respect to each  $\{x_{i_1, d_1}, \dots, x_{i_s, d_s}\}_{i_j \in [m], d_j \in [p]}$ .
- Tensor norms for  $A_s$  with respect to each  $\mathcal{J}$  in  $P_s$  - the set of all possible partitions of  $[s]$ .

**Computing  $A^1$ :** The first order derivative of  $f$  with respect to  $x_{id}$  for any  $i \in [m]$  and  $d \in [p]$ :  $\frac{\partial f(X)}{\partial x_{id}} =$

$$4x_{id}^3 y_i z_i + \sum_{d' \neq d} 2x_{id} x_{id'}^2 y_i z_i + \sum_{j \neq i} 2x_{id} x_{jd}^2 y_i z_j + \sum_{d' \neq d} \sum_{j \neq i} x_{id'} x_{jd} x_{j d'} y_i z_j. \quad (5)$$

$$\mathbb{E}\left(\frac{\partial f(X)}{\partial x_{id}}\right) = O(\sqrt{p \log p}). \quad (6)$$

Therefore,  $A^1 = O(\sqrt{p \log p}) \mathbb{J}_{mp}$ , where  $\mathbb{J}_{mp} \in \mathbb{R}^{mp}$  denotes the vector of ones.

**Tensor norms of  $A^1$ .**  $P_1 = \{1\}$  and

$$\|A^1\|_{\{1\}} = \sup \left\{ \sum_{1 \in [m], d \in [p]} A_{id}^1 x_{id}^1 : \|x^1\|_2 \leq 1 \right\} = \|A^1\|_2 = O(p\sqrt{m \log p}).$$

All the inequalities in this proof are obtained from multiple applications of Hölder's inequality with  $p = 1$  and  $q = \infty$ , CauchySchwarz inequality and the inequality: for any  $x \in \mathbb{R}^n$ ,  $\|x\|_1 \leq \sqrt{n} \|x\|_2$ .

**Computing  $A^2$ :** The second order derivative of  $f$  with respect to  $x_{id}, x_{k\beta}$  for any  $i, k \in [m]$  and  $d, \beta \in [p]$ :  $\frac{\partial^2 f(X)}{\partial x_{id} \partial x_{k\beta}} =$

$$\begin{cases} 12x_{id}^2 y_i z_i + \sum_{d' \neq d} 2x_{id'}^2 y_i z_i + \sum_{j \neq i} 2x_{jd}^2 y_i z_j & \text{if } k = i; \beta = d, \\ 4x_{id} x_{i\beta} y_i z_i + \sum_{j \neq i} 2x_{jd} x_{j\beta} y_i z_j & \text{if } k = i; \beta \neq d, \\ 4x_{id} x_{kd} y_i z_k + \sum_{d' \neq d} x_{id'} x_{kd'} y_i z_k & \text{if } k \neq i; \beta = d, \\ x_{i\beta} x_{kd} y_i z_k & \text{otherwise .} \end{cases} \quad (7)$$

Then  $A^2(i, j) =$

$$\begin{cases} O(p) & \text{if } k = i; \beta = d, \\ O(\log p) & \text{if } k = i; \beta \neq d, \\ O(1) & \text{if } k \neq i; \beta = d, \\ O(\log p/p) & \text{otherwise .} \end{cases} \quad (8)$$

**Tensor norms of  $A^2$**

$P_2 = \{\{1, 2\}, \{\{1\} \{2\}\}\}$ .

From the definition, its clear that  $A_{\{1,2\}}^2 = \|A\|_2 = O(p^2)$ .

$$\begin{aligned} A_{\{\{1\} \{2\}\}}^2 &= \sup \left\{ \sum_{i,j \in [m], d,d' \in [p]} A_{id,jd'}^2 x_{id}^1 x_{jd'}^2 : \|x^1\|_2, \|x^2\|_2 \leq 1 \right\} \\ &\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i \in [m]} \sum_{d \in [p]} O(p) |x_{id}^1 x_{id}^2| + \sum_{i \in [m]} \sum_{d \neq d' \in [p]} O(\log p) |x_{id}^1 x_{id'}^2| \right. \\ &\quad \left. + \sum_{i \neq j \in [m]} \sum_{d \in [p]} O(1) |x_{id}^1 x_{jd}^2| + \sum_{i \neq j \in [m]} \sum_{d \neq d' \in [p]} \rho O\left(\frac{\log p}{p}\right) |x_{id}^1 x_{jd'}^2| \right\} \\ &\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ O(p) \|x^1\|_2 \|x^2\|_2 + O(\log p) \sum_{d \neq d' \in [p]} \sqrt{\sum_{i \in [m]} (x_{id}^1)^2 \sum_{i \in [m]} (x_{id'}^2)^2} \right. \\ &\quad \left. + \sum_{i \neq j \in [m]} \sum_{d \in [p]} O(1) \sqrt{\sum_{d \in [p]} (x_{id}^1)^2 \sum_{d \in [p]} (x_{jd}^2)^2} + \rho O\left(\frac{\log p}{p}\right) (\sqrt{mp} \|x^1\|_2) (\sqrt{mp} \|x^2\|_2) \right\} \\ &\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ O(p) \|x^1\|_2 \|x^2\|_2 + O(\log p) \sqrt{p} \|x^1\|_2 \sqrt{p} \|x^2\|_2 \right. \\ &\quad \left. + O(1) \sqrt{m} \|x^1\|_2 \sqrt{m} \|x^2\|_2 + \rho O\left(\frac{\log p}{p}\right) (\sqrt{mp} \|x^1\|_2) (\sqrt{mp} \|x^2\|_2) \right\} \\ &\leq O(p \log p). \end{aligned}$$

**Computing  $A^3$ :** The third order derivative of  $f$  with respect to  $x_{id}, x_{k\beta}, x_{\alpha l}$  for any  $i, k \in [m]$  and  $d, \beta \in [p]$ :  $\frac{\partial f(X)}{\partial x_{id} \partial x_{k\beta} \partial x_{\alpha l}} =$

$$\begin{cases} 24x_{id}y_i z_i & \text{if } \alpha = k = i; l = \beta = d, \\ 4x_{il}y_i z_i & \text{if } \alpha = k = i; l \neq \beta = d, \\ 4x_{\alpha d}y_i z_\alpha & \text{if } \alpha \neq k = i; l = \beta = d, \\ x_{\alpha\beta}y_i z_\alpha & \text{if } \alpha \neq k = i; l = d \neq \beta, \\ 0 & \text{otherwise .} \end{cases} \quad (9)$$

Then  $A^3_{id,k\beta,\alpha l} =$

$$\begin{cases} O\left(\frac{\log p}{p}\right)y_i z_i & \text{if } \alpha = k = i; l = \beta = d, \\ O\left(\frac{\log p}{p}\right)y_i z_i & \text{if } \alpha = k = i; l \neq \beta = d, \\ O\left(\frac{\log p}{p}\right)y_i z_\alpha & \text{if } \alpha \neq k = i; l = \beta = d, \\ O\left(\frac{\log p}{p}\right)y_i z_\alpha & \text{if } \alpha \neq k = i; l = d \neq \beta, \\ 0 & \text{otherwise .} \end{cases} \quad (10)$$

**Tensor norms of  $A^3$ :** The set of all possible partitions of  $[3]$  up to symmetries is  $P_3 =$

$$\{\{\{1, 2, 3\}\}, \{\{1\}, \{2\}, \{3\}\}, \{\{1\}, \{2, 3\}\}\}$$

**Computing  $\|A^3\|_{\{\{1,2,3\}\}}$ :** It follows from the definition that  $\|A^3\|_{\{\{1,2,3\}\}} = \|A^3\|_2 = O(m\sqrt{p \log p})$ .

**Computing  $\|A^3\|_{\{\{1\},\{2\},\{3\}\}}$ :**

$$\begin{aligned} &= \sup \left\{ \sum_{i_1, i_2, i_3} a_{i_1, i_2, i_3} x_{i_1}^1 x_{i_2}^2 x_{i_3}^3 : \forall l, \|x^l\|_2 \leq 1 \right\} \\ &\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sum_{d, d' \in [p]} |x_{id}^1| |x_{id'}^2| |x_{jd}^3| \right\} \cdot O\left(\sqrt{\frac{\log p}{p}}\right) \\ &\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sqrt{\sum_{d \in [p]} (x_{id}^1)^2} \sqrt{\sum_{d \in [p]} (x_{jd}^2)^2} \sqrt{p} \sqrt{\sum_{d \in [p]} (x_{id'}^3)^2} \right\} \cdot O\left(\sqrt{\frac{\log p}{p}}\right) \\ &\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sqrt{mp} \|x^1\|_2 \|x^2\|_2 \|x^3\|_2 \right\} \cdot O\left(\sqrt{\frac{\log p}{p}}\right) \\ &\leq O(\sqrt{p \log p}). \end{aligned}$$

**Computing**  $\|A^3\|_{\{\{1,2\},\{3\}\}}$ :

$$\begin{aligned}
&= \sup \left\{ \sum_{i_1, i_2, i_3} a_{i_1, i_2, i_3} x_{i_1}^1 x_{i_2}^2 : \forall l, \|x^l\|_2 \leq 1 \right\} \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sum_{d, d' \in [p]} |x_{id}^1| |x_{id', jd}^2| \right\} \cdot \|A^3\|_\infty \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sqrt{\sum_{d \in [p]} (x_{id}^1)^2} \sqrt{p} \sqrt{\sum_{d' \in [p]} (x_{id', jd}^2)^2} \right\} \cdot O\left(\sqrt{\frac{\log p}{p}}\right) \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sqrt{mp} \|x^1\|_2 \|x^2\|_2 \right\} \cdot O\left(\sqrt{\frac{\log p}{p}}\right) \\
&\leq O(\sqrt{p \log p}).
\end{aligned}$$

**Computing**  $A^4$ : The fourth order derivative of  $f$  with respect to  $x_{id}, x_{k\beta}, x_{\alpha l}, x_{q\gamma}$  for any  $i, k, \alpha, q \in [m]$  and  $d, \beta, l, \gamma \in [p]$ :  $\frac{\partial^4 f(X)}{\partial x_{id} \partial x_{k\beta} \partial x_{\alpha l} \partial x_{q\gamma}} =$

$$\begin{cases}
24y_i z_i & \text{if } q = \alpha = k = i; \gamma = l = \beta = d, \\
4y_i z_i & \text{if } q = \alpha = k = i; \gamma = l \neq \beta = d, \\
4y_i z_\alpha & \text{if } q = \alpha \neq k = i; \gamma = l = \beta = d, \\
y_i z_\alpha & \text{if } q = \alpha \neq k = i; l = d \neq \beta = \gamma, \\
0 & \text{otherwise .}
\end{cases} \quad (11)$$

**Tensor norms of**  $A^4$ : The list of all possible partitions of  $[4]$  is the following (up to symmetries).  $P_{[4]} = \{\{\{1, 2, 3, 4\}\}, \{\{1\}, \{2\}, \{3\}, \{4\}\}, \{\{1, 2\}, \{3, 4\}\}, \{\{1\}, \{2, 3, 4\}\}, \{\{1\}, \{2\}, \{3, 4\}\}\}$ .

**Computation of**  $\|A^4\|_{\{\{1,2,3,4\}\}}$ :

It's clear from the definition that  $\|A^4\|_{\{\{1,2,3,4\}\}} = \|A\|_2 \leq 24mp$ .

**Computation of**  $\|A^4\|_{\{\{1\},\{2\},\{3\},\{4\}\}}$ :

$$\begin{aligned}
&= \sup \left\{ \sum_{i_1, i_2, i_3, i_4} a_{i_1, i_2, i_3, i_4} x_{i_1}^1 x_{i_2}^2 x_{i_3}^3 x_{i_4}^4 : \forall l, \|x^l\|_2 \leq 1 \right\} \\
&= \sup \left\{ \sum_{i, j \in [m]} \sum_{d, d' \in [p]} A_{id, id', jd, jd'}^4 x_{id}^1 x_{id'}^2 x_{jd}^3 x_{jd'}^4 : \forall l, \|x^l\|_2 \leq 1 \right\} \\
&\leq \sup \left\{ \sum_{i \in [m]} \left( \sqrt{\sum_{d \in [p]} (x_{id}^1)^2} \sqrt{\sum_{d' \in [p]} (x_{id'}^2)^2} \right) \sum_{j \in [m]} \left( \sqrt{\sum_{d \in [p]} (x_{jd}^3)^2} \sqrt{\sum_{d' \in [p]} (x_{jd'}^4)^2} \right) \right. \\
&\quad \left. : \forall l, \|x^l\|_2 \leq 1 \right\} \|A^4\|_\infty. \\
&\leq \|A^4\|_\infty = 24.
\end{aligned}$$

Note that the first inequality follows from a simultaneous application of the Holder's inequality with  $p = 1$  and  $q = \infty$  and the Cauchy schwarz inequality. The last step follows from an application of the Cauchy-Schwarz inequality.

**Computation of  $A^4_{\{\{1,2\},\{3,4\}\}}$  :**

$$\begin{aligned}
&= \sup \left\{ \sum_{i_1, i_2, i_3, i_4} a_{i_1, i_2, i_3, i_4} x_{i_1, i_2}^1 x_{i_3, i_4}^2 : \forall l, \|x^l\|_2 \leq 1 \right\} \\
&= \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sum_{d, d' \in [p]} A^4_{id, id', jd, jd'} (x_{id, id'}^1 x_{jd, jd'}^2 + x_{id, jd}^1 x_{id', jd'}^2 + x_{id, jd'}^1 x_{id', jd}^2 \right. \\
&\quad \left. + x_{id, id'}^1 x_{jd', jd}^2 + x_{id, jd}^1 x_{jd', id'}^2 + x_{id, jd'}^1 x_{jd, id'}^2) \right\} \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sqrt{\sum_{d, d' \in [p]} (x_{id, id'}^1)^2 \sum_{d, d' \in [p]} (x_{jd, jd'}^2)^2} + \right. \\
&\quad \left. \sum_{d, d' \in [p]} \sqrt{\sum_{i, j \in [m]} (x_{id, jd}^1)^2 \sum_{i, j \in [m]} (x_{id', jd'}^2)^2} + \right. \\
&\quad \left. \sqrt{\sum_{i, j \in [m]} \sum_{d, d' \in [p]} (x_{id, jd'}^1)^2 \sum_{i, j \in [m]} \sum_{d, d' \in [p]} (x_{id', jd}^2)^2} \right\} 2\|A^4\|_\infty. \\
&\leq 2\|A^4\|_\infty(m + p + 1) = 48(m + p + 1).
\end{aligned}$$

**Computation of  $A^4_{\{\{1\},\{2,3,4\}\}}$  :**

$$\begin{aligned}
&= \sup \left\{ \sum_{i_1, i_2, i_3, i_4} a_{i_1, i_2, i_3, i_4} x_{i_1}^1 x_{i_2, i_3, i_4}^2 : \forall l, \|x^l\|_2 \leq 1 \right\} \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sum_{d, d' \in [p]} A^4_{id, id', jd, jd'} (x_{id}^1 x_{id', jd, jd'}^2 + x_{id}^1 x_{id', jd', jd}^2 + x_{id}^1 x_{id', jd, jd'}^2 \right. \\
&\quad \left. + x_{id}^1 x_{jd', jd, id'}^2 + x_{id}^1 x_{jd, jd', id'}^2 + x_{id}^1 x_{jd', jd, id'}^2) \right\} \cdot \|A^4\|_\infty, \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i \in [m]} \sum_{d \in [p]} |x_{id}^1| \sum_{d' \in [p]} x_{id', jd, jd'}^2 \right\} \cdot 6\|A^4\|_\infty, \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sqrt{p} \sum_{i \in [m]} \sqrt{\sum_{d \in [p]} (x_{id}^1)^2} \left( \sum_{j \in [m]} \sqrt{\sum_{d, d' \in [p]} (x_{id', jd, jd'}^2)^2} \right) \right\} \cdot 6\|A^4\|_\infty, \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sqrt{mp} \sum_{i \in [m]} \sqrt{\sum_{d \in [p]} (x_{id}^1)^2} \left( \sqrt{\sum_{j \in [m]} \sum_{d, d' \in [p]} (x_{id', jd, jd'}^2)^2} \right) \right\} \cdot 6\|A^4\|_\infty, \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sqrt{mp} \sqrt{\sum_{i \in [m]} \sum_{d \in [p]} (x_{id}^1)^2} \left( \sqrt{\sum_{i, j \in [m]} \sum_{d, d' \in [p]} (x_{id', jd, jd'}^2)^2} \right) \right\} \cdot 6\|A^4\|_\infty \\
&\leq 6\|A^4\|_\infty(\sqrt{mp}) = 144\sqrt{mp}.
\end{aligned}$$

**Computation of  $A_{\{\{1\},\{2\},\{3,4\}\}}^4$ :**

$$\begin{aligned}
&= \sup \left\{ \sum_{i_1, i_2, i_3, i_4} a_{i_1, i_2, i_3, i_4} \cdot x_{i_1}^1, x_{i_2}^2, x_{i_3, i_4}^3 : \forall l, \|x^l\|_2 \leq 1 \right\} \\
&= \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sum_{d, d' \in [p]} A_{i_d, i_{d'}, j_d, j_{d'}}^4 (x_{i_d}^1, x_{i_{d'}}^2, x_{j_d, j_{d'}}^3 + x_{i_d}^1, x_{j_d}^2, x_{i_{d'}, j_{d'}}^3 + x_{i_d}^1, x_{j_{d'}}^2, x_{i_{d'}, j_d}^3 \right. \\
&\quad \left. + x_{i_d}^1, x_{i_{d'}}^2, x_{j_d, j_d}^3 + x_{i_d}^1, x_{j_d}^2, x_{j_{d'}, i_{d'}}^3 + x_{i_d}^1, x_{j_{d'}}^2, x_{j_d, i_{d'}}^3) \right\} \\
&\leq \sup_{\forall l, \|x^l\|_2 \leq 1} \left\{ \sum_{i, j \in [m]} \sqrt{\sum_{d, d' \in [p]} (x_{i_d}^1 x_{i_{d'}}^2)^2 \sum_{d, d' \in [p]} (x_{j_d, j_{d'}}^3)^2} + \right. \\
&\quad \left. \sum_{d, d' \in [p]} \sqrt{\sum_{i, j \in [m]} (x_{i_d}^1 x_{j_d}^2)^2 \sum_{i, j \in [m]} (x_{i_{d'}, j_{d'}}^3)^2} \right. \\
&\quad \left. + \sqrt{\sum_{i, j \in [m]} \sum_{d, d' \in [p]} (x_{i_d}^1 x_{j_{d'}}^2)^2 \sum_{i, j \in [m]} \sum_{d, d' \in [p]} (x_{i_{d'}, j_d}^3)^2} \right\} \cdot 2 \|A^4\|_\infty. \\
&\leq 2 \|A^4\|_\infty (m + p + 1) = 48(m + p + 1).
\end{aligned}$$

□

Gathering all the norms, we have that for any fixed  $y, z \in \{\pm 1\}^m$ :

$$\begin{aligned}
\mathbb{P}(f(X) \geq C_f(m^2 \rho + m) + t) &\leq 2 \exp \left( -\frac{1}{C} \min \left( \left( \frac{t}{24mp} \right)^2, \left( \frac{t}{24} \right)^{\frac{1}{2}}, \left( \frac{t}{4(m+p+1)} \right), \right. \right. \\
&\quad \left. \left. \left( \frac{t}{\sqrt{mp}} \right), \left( \frac{t}{4(m+p+1)} \right)^{\frac{2}{3}}, \left( \frac{t}{p\sqrt{p \log p}} \right)^2, \left( \frac{t}{p^2} \right), \left( \frac{t}{m\sqrt{p \log p}} \right)^2, \left( \frac{t}{\sqrt{p \log p}} \right)^{\frac{2}{3}}, \left( \frac{t}{\sqrt{mp}} \right) \right) \right). \tag{12}
\end{aligned}$$

Applying a union bound over all possible  $y, z \in \{\pm 1\}^m$  and the setting the R.H.S of Equation 12 to  $\exp(-(1+\epsilon)m \log 2)$ , for some arbitrarily small constant  $\epsilon > 0$  we have that w.h.p,

$$\begin{aligned}
\sup_{\{z, y \in \pm 1\}^m} \kappa \sum_{i, j=1}^m y_i z_j R_{i, j}^{(2)} &\leq \\
&\frac{C_2 \kappa}{p^2} (\rho m(m-1) + m + (mp\sqrt{m} \vee m^2\sqrt{m} \vee p^2\sqrt{m})). \tag{13}
\end{aligned}$$

for some constant  $C_2 > 0$ .



**Proof of Lemma 8.** For any fixed  $y, z \in \{\pm 1\}^m$ ,

$$\begin{aligned} & \sum_{i,j \in [m]} y_i z_j (\langle x_i, x_j \rangle - \mathbb{E}\langle x_i, x_j \rangle) = \\ & \sum_{d \in [p]} \left[ \left( \sum_{i \in [m]} y_i x_{id} \right) \left( \sum_{j \in [m]} z_j x_{jd} \right) - \mathbb{E} \left( \sum_{i \in [m]} y_i x_{id} \right) \left( \sum_{j \in [m]} z_j x_{jd} \right) \right]. \end{aligned} \quad (14)$$

Since each  $x_{id}$  is a normally distributed random variable,  $\sum_{i \in [m]} y_i x_{id}$  is a sub-Gaussian random variable with

$$\left\| \sum_{i \in [m]} y_i x_{id} \right\|_{\psi_2} \leq \sqrt{m} \|x_{id}\|_{\psi_2} \leq \sqrt{m} (1 + O(\log p/p)).$$

Therefore, for each  $d \in [p]$ ,  $\left( \sum_{i \in [m]} y_i x_{id} \right) \left( \sum_{j \in [m]} z_j x_{jd} \right)$  is a sub-exponential random variable with sub-exponential norm:

$$\left\| \left( \sum_{i \in [m]} y_i x_{id} \right) \left( \sum_{j \in [m]} z_j x_{jd} \right) \right\|_{\psi_1} \leq \left\| \sum_{i \in [m]} y_i x_{id} \right\|_{\psi_2} \left\| \sum_{j \in [m]} z_j x_{jd} \right\|_{\psi_2} \leq m(1 + O(\log p/p))^2.$$

Applying Bernstein's inequality for sums of independent sub-exponential random variables, we have that  $\forall t > 0$ ,

$$\begin{aligned} & \Pr \left( \sum_{d \in [p]} \left[ \left( \sum_{i \in [m]} y_i x_{id} \right) \left( \sum_{j \in [m]} z_j x_{jd} \right) - \mathbb{E} \left( \sum_{i \in [m]} y_i x_{id} \right) \left( \sum_{j \in [m]} z_j x_{jd} \right) \right] > t \right) \leq \\ & \exp \left( -c \min \left( \frac{t^2}{pm^2(1 + O(\log p/p))^4}, \frac{t}{m(1 + O(\log p/p))^2} \right) \right). \end{aligned} \quad (15)$$

Applying a union bound over all possible partitions  $y, z \in \{\pm 1\}^m$ , we can see that w.h.p

$$\sup_{y, z \in \{\pm 1\}^m} \sum_{i, j \in [m]} y_i z_j (\langle x_i, x_j \rangle - \mathbb{E}\langle x_i, x_j \rangle) \leq C_1 \left( \frac{m^2}{\sqrt{\alpha}} \vee m^2 \right) \quad (16)$$

for some constant  $C_1 > 0$ . □

**Proof of Lemma 1.** Since for each  $i \in [m]$  and each  $d \in [p]$ ,  $x_{id}$  is a normally distributed random variable, for each  $i, j \in [m]$ ,  $x_{id}x_{jd}$  is a sub-exponential random variable with:

$$\|x_{id}x_{jd}\|_{\psi_1} \leq \|x_{id}\|_{\psi_2} \|x_{jd}\|_{\psi_2} \leq (1 + O(\log p/p))^2.$$

From an application of Bernstein's inequality for sub-exponential random variables, we have that:

$$\Pr(|\langle x_i, x_j \rangle - \mathbb{E}\langle x_i, x_j \rangle| > t) \leq 2 \exp \left( -c \left( \frac{t^2}{p(1 + O(\log p/p))^4} \wedge \frac{t}{(1 + O(\log p/p))^2} \right) \right).$$

Taking a union bound over all  $i, j \in [m]$ , we have that:

$$\max_{i, j \in [m]} \langle x_i, x_j \rangle \leq \mathbb{E} \langle x_i, x_j \rangle + O(\sqrt{p \log p}).$$

and

$$\min_{i, j \in [m]} \langle x_i, x_j \rangle \geq \mathbb{E} \langle x_i, x_j \rangle - O(\sqrt{p \log p}).$$

Since  $\forall i \neq j, \mathbb{E} \langle x_i, x_j \rangle = O(1)$ , we have that w.h.p:

$$\max_{i \neq j \in [m]} \frac{\langle x_i, x_j \rangle}{p} \leq O(\log p / \sqrt{p}), \quad \min_{i \neq j \in [m]} \frac{\langle x_i, x_j \rangle}{p} \geq -O(\sqrt{\log p / p})$$

and  $\forall i \in [m], \mathbb{E} \|x_i\|^2 = p + O(1)$ . So,

$$\max_{i \in [m]} \frac{\|x_i\|}{p} \leq 1 + O(\log p / \sqrt{p}), \quad \min_{i \in [m]} \frac{\|x_i\|}{p} \geq 1 - O(\sqrt{\log p / p})$$

□

**Proof of Lemma 2.** For any partition  $\sigma$  such that  $\|\beta(\sigma, \sigma_*)\|_F^2 \leq 1 + (k-1)\epsilon$ ,

$$\frac{k}{m} \sum_{s=1}^k \sum_{\substack{\sigma(i)=s \\ \sigma(j)=s}} \langle x_i, x_j \rangle = \frac{k}{m} \sum_{s=1}^k \left\| \sum_{\sigma(i)=s} x_i \right\|^2.$$

$\sum_{\sigma(i)=s} x_i$  is the sum of independent normally distributed random variable and is also normally distributed. Therefore,  $\left\| \sum_{\sigma(i)=s} x_i \right\|^2$  follows a non central chi-square distribution with non-centrality:

$$\frac{\alpha\rho}{k} \sum_{s' \in [k]} \sum_{s, t \in [k]} \beta_{s, s'} \beta_{t, s'} \langle \mu_s, \mu_t \rangle = \frac{p\alpha\rho}{k-1} (\|\beta\|_F^2 - 1) + O(\sqrt{p \log p})$$

and  $pk$  degrees of freedom. Applying upper tail bounds from proposition 5, followed a union bound over all such partitions and setting  $t = (1 + \epsilon)m \log k$ , we obtain the following inequality which holds with high probability:

$$\begin{aligned} \max_{\substack{\sigma: \|\beta(\sigma, \sigma_*)\|_F^2 \\ \leq 1 + (k-1)\epsilon}} \frac{k}{m} \sum_{s=1}^k \sum_{\substack{\sigma(i)=s \\ \sigma(j)=s}} Q_{i, j}^{1\sigma} &\leq k + \alpha\rho\epsilon + 2(1 + \epsilon)\alpha \log k \\ &+ 2\sqrt{(1 + \epsilon)(k + 2\alpha\rho\epsilon)\alpha \log k} + O(\sqrt{\log p / p}). \end{aligned} \quad (17)$$

Similarly, the random variable  $\sum_{i=1}^m \sum_{d=1}^p x_{i, d}^2$  is distributed according to a non-central chi-squared distribution with non-centrality  $(\mu^2) p\alpha\rho$  and  $mp$  degrees of freedom ( $d$ ). Note that it is independent of the partition. Using the lower tail bounds from proposition 5 and setting  $t = \log(p)$ , w.p.a.l  $(1 - \frac{1}{p})$ .

$$\max_{\substack{\sigma: \|\beta(\sigma, \sigma_*)\|_F^2 \\ \leq 1 + (k-1)\epsilon}} -\gamma_{\max} Q_i^5 \leq -\frac{k\gamma_{\max} \tau}{mp} \left( mp + p\alpha\rho - 2\sqrt{(mp + 2p\alpha\rho) \log p} \right). \quad (18)$$

□

**Proof of Lemma 4.** Using the inequality,  $\sum_{i=1}^n a_i \cdot b_i \leq \sup_{i \in [n]} |b_i| \cdot \sum_{i=1}^n |a_i|$ , we have:

$$\frac{k}{m} \sum_{i \in [m]} \left( \frac{\|x_i\|^2}{p} - \tau \right)^2 \leq k \max_{i \in [m]} \left( \frac{\|x_i\|^2}{p} - \tau \right)^2 \leq kO\left(\frac{\log p}{p}\right).$$

Therefore,

$$\max_{\substack{\sigma: \|\beta(\sigma, \sigma^*)\|_F^2 \\ \leq 1 + (k-1)\epsilon}} \sum_{i \in [m]} \frac{k\gamma_{\max}(e^\tau - 1)}{2m} Q_i^{4\sigma} \leq C_0 k \gamma_{\max}(e^\tau - 1) (\log p)^2 / 2p. \quad (19)$$

□

**Proof of Lemma 5.** For the true partition  $\sigma^*$ ,  $\|\sum_{\sigma^*(i)=s} x_i\|^2$  follows a non central chi-square distribution with non-centrality:

$$p\alpha\rho + O(\sqrt{p \log p})$$

and  $pk$  degrees of freedom. Applying lower tail bounds from proposition 5 and setting  $t = \log p$ , we obtain the following inequality which holds with high probability:

$$\frac{k}{m} \sum_{s=1}^k \sum_{\substack{\sigma^*(i)=s \\ \sigma^*(j)=s}} Q_{i,j}^{1\sigma} \geq k + \alpha\rho - O(\sqrt{\log p/p}). \quad (20)$$

Similarly, as noted earlier, the random variable  $\sum_{i=1}^m \sum_{d=1}^p x_{id}^2$  is distributed according to a non-central chi-squared distribution with non-centrality  $(\mu^2) p\alpha\rho$  and  $mp$  degrees of freedom  $(d)$ . Using the upper tail bounds from proposition 5 and setting  $t = \log(p)$ , w.p.a.l  $(1 - \frac{1}{p})$ :

$$-\gamma_{\min} Q^5 > -\frac{k\gamma_{\min}\tau}{mp} \left( mp + p\alpha\rho + 2 \log p - 2\sqrt{(mp + 2p\alpha\rho) \log p} \right). \quad (21)$$

□

**Proof of Lemma 7.**

$$\tilde{K}(i, j) = f(0) + \begin{cases} \frac{f'(0)\rho\langle\mu_i, \mu_j\rangle}{p^2} + \frac{\kappa\rho^2\langle\mu_i, \mu_j\rangle^2}{p^4} + \frac{\kappa}{p} & \text{if } i \neq j \\ \frac{f'(0)(p^2 + \rho\|\mu_i\|^2)}{p^2} + \frac{\kappa(p^2 + \rho\|\mu_i\|^2)^2}{p^4} + \frac{\kappa}{p} & \text{otherwise.} \end{cases}$$

$$\begin{aligned} \langle \tilde{K}, X^* - \hat{X} \rangle &= \sum_S \sum_{i \neq j \in S} \tilde{K}_{i,j} (1 - \hat{X}_{i,j}) + \sum_{S, S'} \sum_{\substack{i \in S \\ j \in S'}} \tilde{K}_{i,j} (-\hat{X}_{i,j}) \\ &\geq \min_S \min_{i \neq j \in S} \tilde{K}_{i,j} \sum_S \sum_{i \neq j \in S} (1 - \hat{X}_{i,j}) - \max_{S, S'} \max_{\substack{i \in S \\ j \in S'}} \tilde{K}_{i,j} \sum_{\substack{i \in S \\ j \in S'}} \sum_{\substack{i \in S \\ j \in S'}} (\hat{X}_{i,j}) \\ &= (\min_S \min_{i \neq j \in S} \tilde{K}_{i,j} - \max_{S, S'} \max_{\substack{i \in S \\ j \in S'}} \tilde{K}_{i,j}) \sum_S \sum_{i \neq j \in S} (1 - \hat{X}_{i,j}). \end{aligned}$$

The last inequality is obtained using the property that the sum of entries of each row of  $\hat{X}$  is equal to  $\frac{m}{k}$ . Similarly,

$$\begin{aligned}\|X^* - \hat{X}\|_1 &= \sum_S \sum_{i \neq j \in S} (1 - \hat{X}_{i,j}) + \sum_{S,S'} \sum_{\substack{i \in S \\ j \in S'}} (\hat{X}_{i,j}) \\ &\leq 2 \sum_S \sum_{i \neq j \in S} (1 - \hat{X}_{i,j}).\end{aligned}$$

Therefore,

$$\|X^* - \hat{X}\|_1 \leq \frac{2}{\left(\min_S \min_{i \neq j \in S} \tilde{K}_{i,j} - \max_{S,S'} \max_{\substack{i \in S \\ j \in S'}} \tilde{K}_{i,j}\right)} \langle \tilde{K}, X^* - \hat{X} \rangle. \quad (22)$$

Substituting the values of  $\tilde{K}_{i,j}$ , we have that

$$\begin{aligned}\min_S \min_{i \neq j \in S} \tilde{K}_{i,j} &= f(0) + f'(0) \min_S \frac{\rho \|\mu_s\|^2}{p^2} + e^\tau \gamma_{\max} \min_S \frac{\rho^2 \|\mu_s\|^4}{p^4} \\ &= f(0) + f'(0) \frac{\rho(1 + O(\sqrt{\frac{\log p}{p}}))}{p} + e^\tau \gamma_{\max} \frac{\rho^2(1 + O(\sqrt{\frac{\log p}{p}}))^2}{p^2}.\end{aligned}$$

$$\begin{aligned}\max_{S \neq S'} \max_{i \in S, j \in S'} \tilde{K}_{i,j} &= f(0) + f'(0) \max_{S \neq S'} \frac{\rho \langle \mu_s, \mu_{s'} \rangle}{p^2} + e^\tau \gamma_{\max} \max_{S \neq S'} \frac{\rho^2 \langle \mu_s, \mu_{s'} \rangle^2}{p^4} \\ &= f(0) + f'(0) \frac{\rho(\frac{-1}{k-1} + O(\sqrt{\frac{\log p}{p}}))}{p} + e^\tau \gamma_{\max} \frac{\rho^2(\frac{-1}{k-1} + O(\sqrt{\frac{\log p}{p}}))^2}{p^2}.\end{aligned}$$

where, the second equalities for both quantities arise from substituting the values of  $\min_S \|\mu_s\|^2$  and  $\min_{s,s'} \langle \mu_s, \mu_{s'} \rangle$ . Therefore,

$$\begin{aligned}\min_S \min_{i \neq j \in S} \tilde{K}_{i,j} - \max_{S \neq S'} \max_{i \in S, j \in S'} \tilde{K}_{i,j} &= \frac{\rho}{p} \left( \frac{k}{k-1} + O(\sqrt{\frac{\log p}{p}}) + O(1/p) \right) \text{ and} \\ \|X^* - \hat{X}\|_1 &\leq \frac{2 \langle \tilde{K}, X^* - \hat{X} \rangle}{\frac{\rho}{p} \left( \frac{k}{k-1} + O(\sqrt{\frac{\log p}{p}}) + O(1/p) \right)}.\end{aligned}$$

□

**Proofs of Lemma 3,6.**  $f_{\sigma_*}(X) = \sum_{s \in [k]} \sum_{i,j \in \sigma_*^{-1}(s)} \langle x_i, x_j \rangle^2$ .

$$\mathbb{E} \sum_{s \in [k]} \sum_{i,j \in \sigma_*^{-1}(s)} \langle x_i, x_j \rangle^2 = \frac{mp^2}{k} \left( k + \frac{\alpha}{k} + O\left(\frac{1}{p}\right) \right).$$

We need the following tensors and their corresponding tensor norms to establish the results of lemma 3 and 6 via an application of Proposition 6.

- For each  $s \in [4]$ ,  $s^{\text{th}}$  order tensors  $A_s$  of expectations of  $s^{\text{th}}$  order derivatives of  $f$  with respect to each  $\{x_{i_1, d_1}, \dots, x_{i_s, d_s}\}_{i_j \in [m], d_j \in [p]}$ .

- Tensor norms for  $A_s$  with respect to each  $\mathcal{J}$  in  $P_s$ .

**Computing  $A^1$ :** The first order derivative of  $f$  with respect to  $x_{id}$  for any  $i \in [m]$  and  $d \in [p]$ :  $\frac{\partial f_{\sigma_*}(X)}{\partial x_{id}} =$

$$4x_{id}^3 + \sum_{d' \neq d} 2x_{id}x_{id'}^2 + \sum_{s \in [k]} \sum_{j \neq i \in \sigma_*^{-1}(s)} 2x_{id}x_{jd}^2 + \sum_{d' \neq d} \sum_{j \neq i \in \sigma_*^{-1}(s)} x_{id'}x_{jd}x_{jd'}. \quad (23)$$

$$\mathbb{E}\left(\frac{\partial f_{\sigma_*}(X)}{\partial x_{id}}\right) = O(\sqrt{p \log p}). \quad (24)$$

Therefore,

$$A^1 = O(\sqrt{p \log p})\mathbb{J}_{mp}$$

, where  $\mathbb{J}_{mp} \in \mathbb{R}^{mp}$  denotes the vector of ones.

**Computing  $A^2$ :** The second order derivative of  $f$  with respect to  $x_{id}, x_{k\beta}$  for any  $i, k \in [m]$  and  $d, \beta \in [p]$ :  $\frac{\partial^2 f_{\sigma_*}(X)}{\partial x_{id} \partial x_{k\beta}} =$

$$\begin{cases} 12x_{id}^2 + \sum_{d' \neq d} 2x_{id'}^2 + \sum_{j \neq i \in \sigma_*^{-1}(s)} 2x_{jd}^2 & \text{if } k = i; \beta = d, \\ 4x_{id}x_{i\beta} + \sum_{j \neq i \in \sigma_*^{-1}(s)} 2x_{jd}x_{j\beta} & \text{if } k = i; \beta \neq d, \\ 4x_{id}x_{kd} + \sum_{d' \neq d} x_{id'}x_{kd'} & \text{if } k \neq i \in \sigma_*^{-1}(s); s \in [k]; \beta = d, \\ x_{i\beta}x_{kd} & \text{if } k \neq i \in \sigma_*^{-1}(s); s \in [k]; \beta \neq d, \\ 0 & \text{otherwise .} \end{cases} \quad (25)$$

Then  $A^2(i, j) =$

$$\begin{cases} O(p) & \text{if } k = i; \beta = d, \\ O(\log p) & \text{if } k = i; \beta \neq d, \\ O(1) & \text{if } k \neq i \in \sigma_*^{-1}(s); s \in [k]; \beta = d, \\ O(\log p/p) & \text{if } k \neq i \in \sigma_*^{-1}(s); s \in [k]; \beta \neq d, \\ 0 & \text{otherwise .} \end{cases} \quad (26)$$

**Computing  $A^3$ :** The third order derivative of  $f_{\sigma_*}(X)$  with respect to  $x_{id}, x_{k\beta}, x_{\alpha l}$  for any  $i, k \in [m]$  and  $d, \beta \in [p]$ :  $\frac{\partial^3 f_{\sigma_*}(X)}{\partial x_{id} \partial x_{k\beta} \partial x_{\alpha l}} =$

$$\begin{cases} 24x_{id} & \text{if } \alpha = k = i; l = \beta = d, \\ 4x_{il} & \text{if } \alpha = k = i; l \neq \beta = d, \\ 4x_{\alpha d} & \text{if } \alpha \neq k = i; l = \beta = d, \alpha \neq i \in \sigma_*^{-1}(s); s \in [k]; \\ x_{\alpha\beta} & \text{if } \alpha \neq k = i; l = d \neq \beta, \alpha \neq i \in \sigma_*^{-1}(s); s \in [k]; \\ 0 & \text{otherwise .} \end{cases} \quad (27)$$

Then  $A_{id,k\beta,\alpha l}^3 =$

$$\begin{cases} O\left(\frac{\log p}{p}\right) & \text{if } \alpha = k = i; l = \beta = d, \\ O\left(\frac{\log p}{p}\right) & \text{if } \alpha = k = i; l \neq \beta = d, \\ O\left(\frac{\log p}{p}\right) & \text{if } \alpha \neq k = i; l = \beta = d, \\ O\left(\frac{\log p}{p}\right) & \text{if } \alpha \neq k = i; l = d \neq \beta, \\ 0 & \text{otherwise .} \end{cases} \quad (28)$$

**Computing  $A^4$ :** The fourth order derivative of  $f_{\sigma_*}(X)$  with respect to  $x_{id}, x_{k\beta}, x_{\alpha l}, x_{q\gamma}$  for any  $i, k, \alpha, q \in [m]$  and  $d, \beta, l, \gamma \in [p]$ :  $\frac{\partial^4 f_{\sigma_*}(X)}{\partial x_{id} \partial x_{k\beta} \partial x_{\alpha l} \partial x_{q\gamma}} =$

$$\begin{cases} 24 & \text{if } q = \alpha = k = i; \gamma = l = \beta = d, \\ 4 & \text{if } q = \alpha = k = i; \gamma = l \neq \beta = d, \\ 4 & \text{if } q = \alpha \neq k = i; \gamma = l = \beta = d, \alpha \neq i \in \sigma_*^{-1}(s); s \in [k]; \\ 1 & \text{if } q = \alpha \neq k = i; l = d \neq \beta = \gamma, \alpha \neq i \in \sigma_*^{-1}(s); s \in [k]; \\ 0 & \text{otherwise .} \end{cases} \quad (29)$$

Computing all the tensor norms, (see the proof of Lemma 9 for how the norms are computed) we have:

$$\begin{aligned} \|A^1\|_{\{1\}} &= O(p\sqrt{p \log p}); & \|A^2\|_{\{1,2\}} &= O(p^2); & \|A^2\|_{\{\{1\},\{2\}\}} &= O(p \log p); \\ \|A^3\|_{\{1,2,3\}} &= O(m\sqrt{p \log p}); & \|A^3\|_{\{1,2\},\{3\}} &= O(\sqrt{mp}); & \|A^3\|_{\{\{1\},\{2\},\{3\}\}} &= O(\sqrt{p \log p}); \\ \|A^4\|_{\{1,2,3,4\}} &= O(mp); & \|A^4\|_{\{1,2\},\{3,4\}} &= O(p); & \|A^4\|_{\{\{1\},\{2\},\{3\},\{4\}\}} &= O(1); \\ \|A^4\|_{\{1\},\{2,3,4\}} &= O(\sqrt{mp}); & \|A^4\|_{\{\{1\},\{2\},\{3,4\}\}} &= O(p); \end{aligned}$$

Applying the lower tail bounds for  $f_{\sigma_*}(X)$  from Proposition 6, and setting the R.H.S of the inequality to  $\frac{1}{p}$ , we derive the following upper bound that holds with probability at least  $1 - 1/p$ :

$$f_{\sigma_*}(X) > \frac{mp^2}{k} \left( k + \frac{\alpha}{k} + O\left(\frac{1}{p}\right) \right) - (O(p^2 \sqrt{\log p}) \vee O(mp \sqrt{\log p}))$$

Therefore,

$$\gamma_{\min} Q_{2\sigma_*} > \gamma_{\min} \left( 1 + \frac{1}{k} + O\left(\frac{1}{p}\right) \right) - C_2 \gamma_{\min} \left( \sqrt{\frac{\log p}{p^2}} \vee \alpha \sqrt{\frac{\log p}{p^2}} \right).$$

Fix some  $\epsilon > 0$  be an arbitrarily small constant. For any  $\sigma : \|\beta(\sigma, \sigma_*)\|_F^2 < 1 + (k-1)\epsilon$ , let  $f_\sigma(X) = \sum_{s \in [k]} \sum_{i, j \in \sigma^{-1}(s)} \langle x_i, x_j \rangle^2$ .

We can show that  $\mathbb{E} f_\sigma(X) \leq \frac{mp^2}{k} \left( k + \frac{\alpha}{k} + O\left(\frac{1}{p}\right) \right)$ . Computing the tensors and their respective norms similarly as above, applying proposition 6, followed by an union bound over all such partitions and we can show that w.h.p,

$$\max_{\substack{\sigma : \|\beta(\sigma, \sigma_*)\|_F^2 \\ \leq 1 + (k-1)\epsilon}} \gamma_{\max} Q_{2\sigma} \leq \gamma_{\max} \left( 1 + \frac{1}{k} + O\left(\frac{1}{p}\right) \right) + C_2 \gamma_{\max} O \left( \sqrt{\frac{\alpha}{p}} \vee \alpha \sqrt{\frac{\alpha}{p}} \vee \sqrt{\frac{1}{\alpha p}} \right). \quad (30)$$

□