# Appendix : On Coresets for Regularized Regression

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## Proof of Theorem 3.2

*Proof.* First we bound the sample size for a fixed query  $\mathbf{q} \in Q$ . Let  $s_i$  be the sensitivity of the  $i^{th}$  point  $\mathbf{x_i}$  and S be the sum of the sensitivities. Let the sampling probability be  $p_i = \frac{s_i}{S}$ .

For all  $\mathbf{q} \in Q$  and  $\mathbf{x}_{\mathbf{i}} \in \mathbf{X}$  define a function  $g_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}}) = \frac{f_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}})}{Sp_i \sum_{j=1}^n f_{\mathbf{q}}(\mathbf{x}_{\mathbf{j}})}$ . So,

$$\mathbb{E}[g_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}})] = \frac{1}{S}$$

and

$$\frac{1}{r} \sum_{\substack{i \in [n] s.t.\\ \tilde{\mathbf{x}_i} \in \mathbf{C}}} g_{\mathbf{q}}(\mathbf{x_i}) = \frac{\sum_{\tilde{\mathbf{x}_i} \in \mathbf{C}} f_{\mathbf{q}}(\tilde{\mathbf{x}_i})}{S \sum_{\mathbf{x_i} \in \mathbf{X}} f_{\mathbf{q}}(\mathbf{x_i})}$$

Let

$$T = \sum_{\substack{i \in [n]s.t.\\ \tilde{\mathbf{x}_i} \in \mathbf{C}}} g_{\mathbf{q}}(\mathbf{x_i})$$

then

$$\mathbb{E}[T] = \sum_{\substack{i \in [n] \ s.t.\\ \bar{\mathbf{x}}_i \in \mathbf{C}}} \mathbb{E}[g_{\mathbf{q}}(\mathbf{x}_i)] = r/S$$

$$\operatorname{var}(g_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}})) \leq \mathbb{E}[(g_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}}))^{2}]$$

$$= \sum_{\mathbf{x}_{\mathbf{i}}\in\mathbf{X}} \frac{(f_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}}))^{2}}{(\sum_{j=1}^{n} f_{\mathbf{q}}(\mathbf{x}_{\mathbf{j}}))^{2}S^{2}p_{i}}$$

$$\leq \sum_{\mathbf{x}_{\mathbf{i}}\in\mathbf{X}} \frac{(f_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}}))^{2}\sum_{j=1}^{n} f_{\mathbf{q}}(\mathbf{x}_{\mathbf{j}})}{(\sum_{j=1}^{n} f_{\mathbf{q}}(\mathbf{x}_{\mathbf{j}}))^{2}f_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}})S}$$

$$= 1/S$$

We get the third equation by replacing values of  $p_i$  and  $s_i$ .

Now  $\operatorname{var}(g_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}})) \leq \mathbb{E}[(g_{\mathbf{q}}(\mathbf{x}_{\mathbf{i}}))^2] \leq 1/S$ . So  $\operatorname{var}(T) \leq r/S$ . Now applying Bernstein Inequality as given in [2] we get,

$$\Pr(|T - \mathbb{E}[T]| \ge r\epsilon') \le \exp(-\frac{r^2 \epsilon'^2}{r/S + r\epsilon'/3})$$
$$\Pr\left(\left|\frac{\sum_{\tilde{\mathbf{x}}_i \in \mathbf{C}} f_{\mathbf{q}}(\tilde{\mathbf{x}}_i)}{S\sum_{\mathbf{x}_i \in \mathbf{X}} f_{\mathbf{q}}(\mathbf{x}_i)} - \frac{1}{S}\right| \ge \epsilon'\right) \le \exp\left(-\frac{r\epsilon'^2}{(1/S) + (\epsilon'/3)}\right)$$

Replacing  $\epsilon'$  with  $\epsilon/S$  we get,

$$\Pr\left(\left|\sum_{\tilde{\mathbf{x}}_{i}\in\mathbf{C}}f_{\mathbf{q}}(\tilde{\mathbf{x}}_{i})-\sum_{\mathbf{x}_{i}\in\mathbf{X}}f_{\mathbf{q}}(\mathbf{x}_{i})\right|\geq\epsilon\sum_{\mathbf{x}_{i}\in\mathbf{X}}f_{\mathbf{q}}(\mathbf{x}_{i})\right)\leq2\exp\left(\frac{-2r\epsilon^{2}}{S(1+\frac{\epsilon}{3})}\right)$$

To make the above probability less than  $\delta$ , we choose  $r \geq \frac{S}{2\epsilon^2}(1+\frac{\epsilon}{3})\log\frac{2}{\delta}$  which depends on S for a fixed query  $\mathbf{q} \in Q$ . Now to bound the number of samples required to give a uniform bound for all queries simultaneously  $\forall \mathbf{q} \in Q$ , we use the same  $\epsilon$ -net argument as described in [1]. This part is essentially a repeat of their argument. However we present it here for completeness. Observe that function  $g_{\mathbf{q}}(\mathbf{x}_i)$  lies in the interval [0, 1]. Due to the bounded dimension d of Q, the queries in Q span a subspace of  $[0, 1]^d$ . There may be an infinite number of queries in Q. However these may be covered up to  $L_1$  distance  $\epsilon/2$  by some set  $Q^* \subset Q$  of  $O(\epsilon^{-d})$  points [3] as given in [1]. For the  $\epsilon$ -net argument let  $\mathcal{E}$  be the bad event that the coreset property is not satisfied by some  $\mathbf{C}$ . Therefore

$$\begin{aligned} \Pr(\mathcal{E}) &= \Pr\left[\exists \mathbf{q} \in Q : \left|\sum_{\tilde{\mathbf{x}_{i}} \in \mathbf{C}} f_{\mathbf{q}}(\tilde{\mathbf{x}_{i}}) - \sum_{\mathbf{x_{i}} \in \mathbf{X}} f_{\mathbf{q}}(\mathbf{x_{i}})\right| > \epsilon \sum_{\mathbf{x_{i}} \in \mathbf{X}} f_{\mathbf{q}}(\mathbf{x_{i}})\right] \\ &\leq \Pr\left[\exists \mathbf{q} \in Q^{*} : \left|\sum_{\tilde{\mathbf{x}_{i}} \in \mathbf{C}} f_{\mathbf{q}}(\tilde{\mathbf{x}_{i}}) - \sum_{\mathbf{x_{i}} \in \mathbf{X}} f_{\mathbf{q}}(\mathbf{x_{i}})\right| > \frac{\epsilon}{2} \sum_{\mathbf{x_{i}} \in \mathbf{X}} f_{\mathbf{q}}(\mathbf{x_{i}})\right] \\ &\leq 2|Q^{*}|\exp\left(\frac{-2r\epsilon^{2}}{S(1+\frac{\epsilon}{3})}\right)\end{aligned}$$

To make **C** an  $\epsilon$ -coreset with probability at least 1- $\delta$ , we choose  $r = O(\frac{S}{\epsilon^2}(\log |Q^*| + \log \frac{2}{\delta}))$ . Now as  $|Q^*| \in O(\epsilon^{-d})$  we have  $r = O(\frac{S}{\epsilon^2}(d\log \frac{1}{\epsilon} + \log \frac{1}{\delta}))$ .

# Generalizing the Proof of Corollary 4.1.1

The proof can be generalized to the setting when **b** is in the column-space of **A** in the following manner. Suppose  $\mathbf{b} = \mathbf{A}\mathbf{u}$ . Also suppose  $\mathbf{A}_{\mathbf{c}}$  and  $\mathbf{b}_{\mathbf{c}}$  can be obtained as  $\mathbf{A}_{\mathbf{c}} = \mathbf{S}\mathbf{A}$  and  $\mathbf{b}_{\mathbf{c}} = \mathbf{S}\mathbf{b}$  where **S** can be either a sampling and reweighing or a sketching matrix. Now we want to prove the following : If **S** is a coreset creation matrix for  $(\mathbf{A}, \mathbf{b})$  for regression i.e.  $\forall \mathbf{x}, \|\mathbf{A}_{\mathbf{c}}\mathbf{x} - \mathbf{b}_{\mathbf{c}}\|_{p}^{r} \in (1 \pm \epsilon) \|\mathbf{A}\mathbf{x} - \mathbf{b}\|_{p}^{r}$ , then it must be that  $\forall \mathbf{x}, \|\mathbf{A}_{\mathbf{c}}\mathbf{x}\|_{p}^{r} \in (1 \pm \epsilon) \|\mathbf{A}\mathbf{x}\|_{p}^{r}$ . Proving this statement and using

Theorem 4.1 essentially proves the corollary for the more general setting of **b** in column space of **A**. To prove the statement we use contradiction. Let us suppose that the statement is false. Then  $\exists \mathbf{v} \in \mathbb{R}^d$  s.t.  $\|\mathbf{A_c v}\|_p^r > (1+\epsilon)\|\mathbf{Av}\|_p^r$ . We will create a **y** s.t that  $\|\mathbf{A_c y} - \mathbf{b_c}\|_p^r > (1+\epsilon)\|\mathbf{Ay} - \mathbf{b}\|_p^r$ . Consider the ratio

$$\frac{\|\mathbf{S}(\mathbf{A}\mathbf{y} - \mathbf{b})\|_p^r}{\|\mathbf{A}\mathbf{y} - \mathbf{b}\|_p^r} = \frac{\|\mathbf{S}\mathbf{A}(\mathbf{y} - \mathbf{u})\|_p^r}{\|\mathbf{A}(\mathbf{y} - \mathbf{u})\|_p^r}$$

Now if we choose  $\mathbf{y} = \mathbf{u} + \mathbf{v}$  then we have  $\frac{\|\mathbf{SAv}\|_p^r}{\|\mathbf{Av}\|_p^r} > (1+\epsilon)$ . This a contradiction to the fact that  $(\mathbf{SA}, \mathbf{Sb})$  is coreset for  $\|\mathbf{Ay} - \mathbf{b}\|_p^r$ . Hence our assumption is false. So  $\forall \mathbf{x}, \|\mathbf{A_cx}\|_p^r \leq (1+\epsilon)\|\mathbf{Ax}\|_p^r$ . The other direction for coreset definition is proved in similar manner. This combined with the Theorem 4.1 gives our corollary

#### Proof of Corollary 6.1.1

*Proof.* For  $\hat{\mathbf{A}} = [\mathbf{A} - \mathbf{B}]$  and  $\hat{\mathbf{X}} = \begin{bmatrix} \mathbf{X} \\ \mathbf{I}_{\mathbf{k}} \end{bmatrix}$  where  $\mathbf{I}_{\mathbf{k}}$  is k-dimensional identity matrix, the sensitivity of Multiresponse RLAD problem is given as

$$s_i = \sup_{\hat{\mathbf{X}}} \frac{\|\hat{\mathbf{a}}_i^T \hat{\mathbf{X}}\|_1 + \frac{\lambda \|\hat{\mathbf{X}}\|_1}{n}}{\sum_j \|\hat{\mathbf{a}}_j^T \hat{\mathbf{X}}\|_1 + \lambda \|\hat{\mathbf{X}}\|_1}$$

Let  $\hat{\mathbf{A}} = \mathbf{U}\mathbf{Y}$  where  $\mathbf{U}$  is an  $(\alpha, \beta, 1)$  well conditioned basis for  $\hat{\mathbf{A}}$ . So  $\hat{\mathbf{a}}_j^T \hat{\mathbf{X}} = \mathbf{u}_j^T \mathbf{Y} \hat{\mathbf{X}}$ . Let  $\mathbf{Y} \hat{\mathbf{X}} = \mathbf{Z}$ . So the sensitivity equation becomes

$$s_i = \sup_{\mathbf{Z}} \frac{\|\mathbf{u}_i^T \mathbf{Z}\|_1 + \lambda \|\mathbf{Y}^{-1} \mathbf{Z}\|_1}{\sum_j \|\mathbf{u}_j^T \mathbf{Z}\|_1 + \lambda \|\mathbf{Y}^{-1} \mathbf{Z}\|_1} \\ \leq \sup_{\mathbf{Z}} \frac{\|\mathbf{u}_i^T \mathbf{Z}\|_1}{\sum_j \|\mathbf{u}_j^T \mathbf{Z}\|_1 + \lambda \|\mathbf{Y}^{-1} \mathbf{Z}\|_1} + \frac{1}{n}$$

Instead of supremum of the first quantity on the right hand side, we take the infimum of its reciprocal. Lets call it m.

$$m = \inf_{\mathbf{Z}} \frac{\sum_{j} \|\mathbf{u}_{j}^{T}\mathbf{Z}\|_{1} + \lambda \|\mathbf{Y}^{-1}\mathbf{Z}\|_{1}}{\|\mathbf{u}_{i}^{T}\mathbf{Z}\|_{1}}$$
  
$$\geq \inf_{\mathbf{Z}} \frac{\sum_{j} \|\mathbf{u}_{j}^{T}\mathbf{Z}\|_{1}}{\|\mathbf{u}_{i}^{T}\mathbf{Z}\|_{1}} + \inf_{\mathbf{Z}} \frac{\lambda \|\mathbf{Y}^{-1}\mathbf{Z}\|_{1}}{\|\mathbf{u}_{i}^{T}\mathbf{Z}\|_{1}}$$

Let us consider the first part. **U** is an  $(\alpha, \beta, 1)$ - well conditioned basis for  $\mathbf{\hat{A}}$ . Hence by definition  $\|\mathbf{U}\|_1 \leq \alpha$  and  $\forall \mathbf{z} \in \mathbb{R}^{d+k}, \|\mathbf{z}\|_{\infty} \leq \beta \|\mathbf{U}\mathbf{z}\|_1$ . So the first term in the infimum

$$\begin{split} \inf_{\mathbf{Z}} \frac{\sum_{j} \|\mathbf{u}_{j}^{T}\mathbf{Z}\|_{1}}{\|\mathbf{u}_{i}^{T}\mathbf{Z}\|_{1}} \\ &= \inf_{\mathbf{Z}} \frac{\sum_{l=1}^{k} \|\mathbf{U}\mathbf{z}^{l}\|_{1}}{\sum_{l=1}^{k} |\mathbf{u}_{i}^{T}\mathbf{z}^{l}|} \\ &\geq \frac{\frac{1}{\beta} \sum_{l=1}^{k} \|\mathbf{z}^{l}\|_{\infty}}{\|\mathbf{u}_{i}\|_{1} \sum_{l=1}^{k} \|\mathbf{z}^{l}\|_{\infty}} \\ &= \frac{1}{\beta \|\mathbf{u}_{i}\|_{1}} \end{split}$$

Now for the second term in the infimum let us consider instead

$$\inf_{\mathbf{Z}} \frac{\|\mathbf{A}\mathbf{Y}^{-1}\mathbf{Z}\|_{1}}{\|\mathbf{u}_{i}^{T}\mathbf{Z}\|_{1}} \\ = \inf_{\mathbf{Z}} \frac{\|\mathbf{U}\mathbf{Z}\|_{1}}{\|\mathbf{u}_{i}^{T}\mathbf{Z}\|_{1}} \\ \ge \frac{1}{\beta\|\mathbf{u}_{i}\|_{1}}$$

Now  $\|\mathbf{A}\mathbf{Y}^{-1}\mathbf{Z}\|_1 \le \|\mathbf{A}\|_{(1)} \|\mathbf{Y}^{-1}\mathbf{Z}\|_1$ . Therefore

$$\begin{split} \inf_{\mathbf{Z}} \frac{\|\mathbf{Y}^{-1}\mathbf{Z}\|_{1}}{\|\mathbf{u}_{i}^{T}\mathbf{Z}\|_{1}} \\ \geq \inf_{\mathbf{Z}} \frac{\|\mathbf{A}\mathbf{Y}^{-1}\mathbf{Z}\|_{1}}{\|\mathbf{A}\|_{(1)}\|\mathbf{u}_{i}^{T}\mathbf{Z}\|_{1}} \\ \geq \frac{1}{\beta\|\mathbf{A}\|_{(1)}\|\mathbf{u}_{i}\|_{1}} \end{split}$$

Combining both these

$$m \ge \frac{1}{\beta \|\mathbf{u}_{\mathbf{i}}\|_1} \left( 1 + \frac{\lambda}{\|\mathbf{A}\|_{(1)}} \right)$$

Now sensitivity of  $i^{th}$  point is bounded as  $s_i \leq \frac{1}{m} + \frac{1}{n}$ . Therefore  $s_i \leq \frac{\beta \|\mathbf{u}_i\|_1}{1 + \frac{\lambda}{\|\mathbf{A}\|_{(1)}}} + \frac{1}{n}$ . So the sum of sensitivities is bounded by  $S \leq \frac{\alpha\beta}{1 + \frac{\lambda}{\|\mathbf{A}\|_{(1)}}} + 1$ . This fact combined with fact that dimension of  $\mathbf{X}$  is dk and applying theorem 3.2 proves the corollary

## References

[1] Olivier Bachem, Mario Lucic, and Andreas Krause. Practical coreset constructions for machine learning. arXiv preprint arXiv:1703.06476, 2017.

- [2] Devdatt P Dubhashi and Alessandro Panconesi. Concentration of measure for the analysis of randomized algorithms. Cambridge University Press, 2009.
- [3] David Haussler. Sphere packing numbers for subsets of the boolean n-cube with bounded vapnik-chervonenkis dimension. J. Comb. Theory, Ser. A, 69(2):217–232, 1995.