

# Appendix

## A PRELIMINARIES

### A.1 Transportation Inequalities

For any function  $f : \mathcal{X} \rightarrow \mathbb{R}$ , we define its span as  $\mathbb{S}(f) := \max_{x \in \mathcal{X}} f(x) - \min_{x \in \mathcal{X}} f(x)$ . For a probability distribution  $P$  supported on the set  $\mathcal{X}$ , let  $\mathbb{E}_P[f] := \mathbb{E}_P[f(X)]$  and  $\mathbb{V}_P[f] := \mathbb{V}_P[f(X)] = \mathbb{E}_P[f(X)^2] - \mathbb{E}_P[f(X)]^2$  denote the mean and variance of the random variable  $f(X)$ , respectively. We now state the following transportation inequalities, which can be adapted from Boucheron et al. (2013, Lemma 4.18).

**Lemma 1** (Transportation inequalities). *Assume  $f$  is such that  $\mathbb{S}(f)$  and  $\mathbb{V}_P[f]$  are finite. Then it holds*

$$\begin{aligned} \forall Q \ll P, \quad \mathbb{E}_Q[f] - \mathbb{E}_P[f] &\leq \sqrt{2 \mathbb{V}_P[f] \text{KL}(Q, P)} + \frac{2 \mathbb{S}(f)}{3} \text{KL}(Q, P), \\ \forall Q \ll P, \quad \mathbb{E}_P[f] - \mathbb{E}_Q[f] &\leq \sqrt{2 \mathbb{V}_P[f] \text{KL}(Q, P)}. \end{aligned}$$

### A.2 Bregman Divergence

For a Legendre function  $F : \mathbb{R}^d \rightarrow \mathbb{R}$ , the Bregman divergence between  $\theta', \theta \in \mathbb{R}^d$  associated with  $F$  is defined as

$$B_F(\theta', \theta) := F(\theta') - F(\theta) - (\theta' - \theta)^\top \nabla F(\theta).$$

Now, for any fixed  $\theta \in \mathbb{R}^d$ , we introduce the function

$$B_{F, \theta}(\lambda) := B_F(\theta + \lambda, \theta) = F(\theta + \lambda) - F(\theta) - \lambda^\top \nabla F(\theta).$$

It then follows that  $B_{F, \theta}$  is a convex function, and we define its dual as

$$B_{F, \theta}^*(x) = \sup_{\lambda \in \mathbb{R}^d} (\lambda^\top x - B_{F, \theta}(\lambda)).$$

We have for any  $\theta, \theta' \in \mathbb{R}^d$ :

$$B_F(\theta', \theta) = B_{F, \theta'}^*(\nabla F(\theta) - \nabla F(\theta')). \quad (4)$$

To see this, we observe that

$$\begin{aligned} &B_{F, \theta'}^*(\nabla F(\theta) - \nabla F(\theta')) \\ &= \sup_{\lambda \in \mathbb{R}^d} \lambda^\top (\nabla F(\theta) - \nabla F(\theta')) - [F(\theta' + \lambda) - F(\theta') - \lambda^\top \nabla F(\theta')] \\ &= \sup_{\lambda \in \mathbb{R}^d} \lambda^\top \nabla F(\theta) - F(\theta' + \lambda) + F(\theta'). \end{aligned}$$

Now an optimal  $\lambda$  must satisfy  $\nabla F(\theta) = \nabla F(\theta' + \lambda)$ . One possible choice is  $\lambda = \theta - \theta'$ . Since, by definition,  $F$  is strictly convex, the supremum will indeed be attained at  $\lambda = \theta - \theta'$ . Plug-in this value, we obtain

$$B_{F, \theta'}^*(\nabla F(\theta) - \nabla F(\theta')) = (\theta - \theta')^\top \nabla F(\theta) - F(\theta) + F(\theta') = B_F(\theta', \theta).$$

(Note that (4) holds for any convex function  $F$ . Only difference is that, in this case,  $B_F(\cdot, \cdot)$  won't correspond to the Bregman divergence.)

### A.3 Exponential Family

In this section, we detail some useful results related to exponential families in our model.

**Derivatives** Let us first take a closer look at the derivative of the log-partition function  $Z_{s,a}$ . As usual with exponential families, these are intimately linked to moments of the random variable. We have on the one hand,

$$\begin{aligned} (\nabla_i Z_{s,a})(\theta) &= \int_{\mathcal{S}} \psi(s')^\top A_i \varphi(s, a) \frac{h(s', s, a) \exp\left(\sum_{i=1}^d \theta_i \psi(s')^\top A_i \varphi(s, a)\right)}{\int_{\mathcal{S}} h(s', s, a) \exp\left(\sum_{i=1}^d \theta_i \psi(s')^\top A_i \varphi(s, a)\right) ds'} ds' \\ &= \mathbb{E}_{s,a}^\theta \left[ \psi(s')^\top A_i \varphi(s, a) \right]. \end{aligned}$$

On the other hand, the entries of the Hessian of  $Z$  are given by

$$\begin{aligned}
 (\nabla_{i,j}^2 Z_{s,a})(\theta) &= \int_{\mathcal{S}} \psi(s')^\top A_i \varphi(s, a) \psi(s')^\top A_j \varphi(s, a) \frac{h(s', s, a) \exp\left(\sum_{i=1}^d \theta_i \psi(s')^\top A_i \varphi(s, a)\right)}{\int_{\mathcal{S}} h(s', s, a) \exp\left(\sum_{i=1}^d \theta_i \psi(s')^\top A_i \varphi(s, a)\right) ds'} ds' \\
 &\quad - \int_{\mathcal{S}} \psi(s')^\top A_i \varphi(s, a) \frac{h(s', s, a) \exp\left(\sum_{i=1}^d \theta_i \psi(s')^\top A_i \varphi(s, a)\right)}{\int_{\mathcal{S}} h(s', s, a) \exp\left(\sum_{i=1}^d \theta_i \psi(s')^\top A_i \varphi(s, a)\right) ds'} ds' (\nabla_j Z_{s,a})(\theta) \\
 &= \mathbb{E}_{s,a}^\theta \left[ \psi(s')^\top A_i \varphi(s, a) \psi(s')^\top A_j \varphi(s, a) \right] \\
 &\quad - \mathbb{E}_{s,a}^\theta \left[ \psi(s')^\top A_i \varphi(s, a) \right] \mathbb{E}_{s,a}^\theta \left[ \psi(s')^\top A_j \varphi(s, a) \right] \\
 &= \varphi(s, a)^\top A_i^\top \left( \mathbb{E}_{s,a}^\theta [\psi(s') \psi(s')^\top] - \mathbb{E}_{s,a}^\theta [\psi(s')] \mathbb{E}_{s,a}^\theta [\psi(s')^\top] \right) A_j \varphi(s, a) \\
 &= \varphi(s, a)^\top A_i^\top \mathbb{C}_{s,a}^\theta [\psi(s')] A_j \varphi(s, a),
 \end{aligned}$$

where we introduce in the last line the  $p \times p$  covariance matrix given by

$$\mathbb{C}_{s,a}^\theta [\psi(s')] = \mathbb{E}_{s,a}^\theta [\psi(s') \psi(s')^\top] - \mathbb{E}_{s,a}^\theta [\psi(s')] \mathbb{E}_{s,a}^\theta [\psi(s')^\top],$$

**KL Divergence** For any two  $\theta, \theta'$  and for some pair  $(s, a)$ , we are interested in the following useful relations

$$\begin{aligned}
 \log \left( \frac{P_\theta(s'|s, a)}{P_{\theta'}(s'|s, a)} \right) &= \sum_{i=1}^d (\theta_i - \theta'_i) \psi(s')^\top A_i \varphi(s, a) - Z_{s,a}(\theta) + Z_{s,a}(\theta'), \\
 \text{or } \text{KL} \left( P_\theta(\cdot|s, a), P_{\theta'}(\cdot|s, a) \right) &= \sum_{i=1}^d (\theta_i - \theta'_i) \mathbb{E}_{s,a}^\theta [\psi(s')]^\top A_i \varphi(s, a) - Z_{s,a}(\theta) + Z_{s,a}(\theta') \\
 &= \frac{1}{2} (\theta - \theta')^\top (\nabla^2 Z_{s,a})(\tilde{\theta}) (\theta - \theta'),
 \end{aligned}$$

where in the last line, we used, by a Taylor expansion, that  $Z_{s,a}(\theta') = Z_{s,a}(\theta) + (\nabla Z_{s,a}(\theta))^\top (\theta' - \theta) + \frac{1}{2} (\theta - \theta')^\top (\nabla^2 Z_{s,a}(\tilde{\theta})) (\theta - \theta')$  for some  $\tilde{\theta} \in [\theta, \theta']_\infty$ . Here  $[\theta, \theta']_\infty$  denotes the  $d$ -dimensional hypercube joining  $\theta$  to  $\theta'$ .

## B METHOD OF MIXTURES FOR CONDITIONAL EXPONENTIAL FAMILIES: PROOF OF THEOREM 1

**Step 1: Martingale Construction** First note that by our hypothesis of strict convexity, the log-partition function  $Z_{s,a}$  is a Legendre function.<sup>7</sup> Now for the conditional exponential family model, the KL divergence b/w  $P_\theta(\cdot|s, a)$  and  $P_{\theta'}(\cdot|s, a)$  can be expressed as a Bregman divergence associated to  $Z_{s,a}$  with the parameters reversed, i.e.,

$$\text{KL}_{s,a}(\theta, \theta') := \text{KL}(P_\theta(\cdot|s, a), P_{\theta'}(\cdot|s, a)) = B_{Z_{s,a}}(\theta', \theta). \quad (5)$$

Now, for any  $\lambda \in \mathbb{R}^d$ , we introduce the function  $B_{Z_{s,a}, \theta^*}(\lambda) = B_{Z_{s,a}}(\theta^* + \lambda, \theta)$  and define

$$M_n^\lambda = \exp \left( \lambda^\top S_n - \sum_{t=1}^n B_{Z_{s_t, a_t}, \theta^*}(\lambda) \right),$$

<sup>7</sup>Since we will only use (4) in the proof, the final result would hold even if  $Z_{s,a}$  is only convex.

where  $\forall i \leq d$ , we denote  $(S_n)_i = \sum_{t=1}^n (\psi(s'_t) - \mathbb{E}_{s_t, a_t}^{\theta^*} [\psi(s'_t)])^\top A_i \varphi(s_t, a_t)$ . Note that  $M_n^\lambda > 0$  and it is  $\mathcal{F}_n$ -measurable. Furthermore, we have for all  $(s, a)$ ,

$$\begin{aligned} & \mathbb{E}_{s, a}^{\theta^*} \left[ \exp \left( \sum_{i=1}^d \lambda_i (\psi(s') - \mathbb{E}_{s, a}^{\theta^*} [\psi(s')])^\top A_i \varphi(s, a) \right) \right] \\ &= \exp(-\lambda^\top \nabla Z_{s, a}(\theta^*)) \int_{\mathcal{S}} h(s', s, a) \exp \left( \sum_{i=1}^d (\theta_i^* + \lambda_i) \psi(s')^\top A_i \varphi(s, a) - Z_{s, a}(\theta^*) \right) ds' \\ &= \exp(Z_{s, a}(\theta^* + \lambda) - Z_{s, a}(\theta^*) - \lambda^\top \nabla Z_{s, a}(\theta^*)) = \exp(B_{Z_{s, a}}(\theta^*)). \end{aligned}$$

This implies  $\mathbb{E}[\exp(\lambda^\top S_n) | \mathcal{F}_{n-1}] = \exp(\lambda^\top S_{n-1} + B_{Z_{s_n, a_n}, \theta^*}(\lambda))$  and thus, in turn,  $\mathbb{E}[M_n^\lambda | \mathcal{F}_{n-1}] = M_{n-1}^\lambda$ . Therefore  $\{M_n^\lambda\}_{n=0}^\infty$  is a non-negative martingale adapted to the filtration  $\{\mathcal{F}_n\}_{n=0}^\infty$  and actually satisfies  $\mathbb{E}[M_n^\lambda] = 1$ . For any prior density  $q(\theta)$  for  $\theta$ , we now define a mixture of martingales

$$M_n = \int_{\mathbb{R}^d} M_n^\lambda q(\theta^* + \lambda) d\lambda. \quad (6)$$

Then  $\{M_n\}_{n=0}^\infty$  is also a non-negative martingale adapted to  $\{\mathcal{F}_n\}_{n=0}^\infty$  and in fact,  $\mathbb{E}[M_n] = 1$ .

**Step 2: Method Of Mixtures and Martingale Control** Considering the prior density  $\mathcal{N}(0, (\eta\mathbb{A})^{-1})$ , we obtain from (6) that

$$M_n = c_0 \int_{\mathbb{R}^d} \exp \left( \lambda^\top S_n - \sum_{t=1}^n B_{Z_{s_t, a_t}, \theta^*}(\lambda) - \frac{\eta}{2} \|\theta^* + \lambda\|_{\mathbb{A}}^2 \right) d\lambda, \quad (7)$$

where  $c_0 = \frac{1}{\int_{\mathbb{R}^d} \exp(-\frac{\eta}{2} \|\theta'\|_{\mathbb{A}}^2) d\theta'}$ . We now introduce the function  $Z_n(\theta) = \sum_{t=1}^n Z_{s_t, a_t}(\theta)$ . Note that  $Z_n$  is a also Legendre function and its associated Bregman divergence satisfies

$$B_{Z_n}(\theta', \theta) = \sum_{t=1}^n (Z_{s_t, a_t}(\theta') - Z_{s_t, a_t}(\theta) - (\theta' - \theta)^\top \nabla Z_{s_t, a_t}(\theta)) = \sum_{t=1}^n B_{Z_{s_t, a_t}}(\theta', \theta)$$

Furthermore, we have  $\sum_{t=1}^n B_{Z_{s_t, a_t}, \theta^*}(\lambda) = B_{Z_n, \theta^*}(\lambda)$ .

From the penalized likelihood formula (2), recall that

$$\forall i \leq d, \quad \sum_{t=1}^n \nabla_i Z_{s_t, a_t}(\theta_n) + \frac{\eta}{2} \nabla_i \|\theta_n\|_{\mathbb{A}}^2 = \sum_{t=1}^n \psi(s'_t)^\top A_i \varphi(s_t, a_t).$$

This yields

$$S_n = \sum_{t=1}^n (\nabla Z_{s_t, a_t}(\theta_n) - \nabla Z_{s_t, a_t}(\theta^*)) + \eta \mathbb{A} \theta_n = \nabla Z_n(\theta_n) - \nabla Z_n(\theta^*) + \eta \mathbb{A} \theta_n. \quad (8)$$

We now obtain from (7) and (8) that

$$M_n = c_0 \cdot \exp \left( -\frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \right) \int_{\mathbb{R}^d} \exp(\lambda^\top x_n - B_{Z_n, \theta^*}(\lambda) + g_n(\lambda)) d\lambda, \quad (9)$$

where we have introduced  $g_n(\lambda) = \frac{\eta}{2} (2\lambda^\top \mathbb{A} \theta_n + \|\theta^*\|_{\mathbb{A}}^2 - \|\theta^* + \lambda\|_{\mathbb{A}}^2)$  and  $x_n = \nabla Z_n(\theta_n) - \nabla Z_n(\theta^*)$ .

Now, note that  $\sup_{\lambda \in \mathbb{R}^d} g_n(\lambda) = \frac{\eta}{2} \|\theta^* - \theta_n\|_{\mathbb{A}}^2$ , where the supremum is attained at  $\lambda^* = \theta_n - \theta^*$ . We then have

$$\begin{aligned} g_n(\lambda) &= g_n(\lambda) + \sup_{\lambda \in \mathbb{R}^d} g_n(\lambda) - g_n(\lambda^*) \\ &= \frac{\eta}{2} \|\theta_n - \theta^*\|_{\mathbb{A}}^2 + \eta(\lambda - \lambda^*)^\top \mathbb{A}(\theta^* + \lambda^*) + \frac{\eta}{2} \|\theta^* + \lambda^*\|_{\mathbb{A}}^2 - \frac{\eta}{2} \|\theta^* + \lambda\|_{\mathbb{A}}^2 \\ &= B_{Z_0}(\theta^*, \theta_n) + (\lambda - \lambda^*)^\top \nabla Z_0(\theta^* + \lambda^*) + Z_0(\theta^* + \lambda^*) - Z_0(\theta^* + \lambda), \end{aligned} \quad (10)$$

where we have introduced the Legendre function  $Z_0(\theta) = \frac{\eta}{2} \|\theta\|_{\mathbb{A}}^2$ . We now have from (4) that

$$\begin{aligned} & \sup_{\lambda \in \mathbb{R}^d} (\lambda^\top x_n - B_{Z_n, \theta^*}(\lambda)) \\ &= B_{Z_n, \theta^*}^*(x_n) = B_{Z_n, \theta^*}^*(\nabla Z_n(\theta_n) - \nabla Z_n(\theta^*)) = B_{Z_n}(\theta^*, \theta_n). \end{aligned}$$

Further, any optimal  $\lambda$  must satisfy

$$\nabla Z_n(\theta^* + \lambda) - \nabla Z_n(\theta^*) = x_n \implies \nabla Z_n(\theta^* + \lambda) = \nabla Z_n(\theta_n).$$

One possible solution is  $\lambda = \lambda^*$ . Now, since  $Z_n$  is strictly convex, the supremum is indeed attained at  $\lambda = \lambda^*$ . We then have

$$\begin{aligned}
 & \lambda^\top x_n - B_{Z_n, \theta^*}(\lambda) \\
 &= \lambda^\top x_n - B_{Z_n, \theta^*}(\lambda) + B_{Z_n}(\theta^*, \theta_n) - \left( \lambda^{*\top} x_n - B_{Z_n, \theta^*}(\lambda^*) \right) \\
 &= B_{Z_n}(\theta^*, \theta_n) + (\lambda - \lambda^*)^\top \nabla Z_n(\theta^* + \lambda^*) + B_{Z_n, \theta^*}(\lambda^*) - B_{Z_n, \theta^*}(\lambda) - (\lambda - \lambda^*)^\top \nabla Z_n(\theta^*) \\
 &= B_{Z_n}(\theta^*, \theta_n) + (\lambda - \lambda^*)^\top \nabla Z_n(\theta^* + \lambda^*) + Z_n(\theta^* + \lambda^*) - Z_n(\theta^* + \lambda) .
 \end{aligned} \tag{11}$$

Plugging (10) and (11) in (9), we now obtain

$$\begin{aligned}
 M_n &= c_0 \cdot \exp \left( \sum_{j \in \{0, n\}} B_{Z_j}(\theta^*, \theta_j) - \frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \right) \\
 &\quad \times \int_{\mathbb{R}^d} \exp \left( \sum_{j \in \{0, n\}} ((\lambda - \lambda^*)^\top \nabla Z_j(\theta^* + \lambda^*) + Z_j(\theta^* + \lambda^*) - Z_j(\theta^* + \lambda)) \right) d\lambda \\
 &= c_0 \cdot \exp \left( \sum_{j \in \{0, n\}} B_{Z_j}(\theta^*, \theta_n) - \frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \right) \cdot \exp \left( - \sum_{j \in \{0, n\}} ((\theta^* + \lambda^*)^\top \nabla Z_j(\theta^* + \lambda^*) - Z_j(\theta^* + \lambda^*)) \right) \\
 &\quad \times \int_{\mathbb{R}^d} \exp \left( \sum_{j \in \{0, n\}} ((\theta^* + \lambda)^\top \nabla Z_j(\theta^* + \lambda^*) - Z_j(\theta^* + \lambda)) \right) d\lambda \\
 &= \frac{c_0}{c_n} \cdot \exp \left( \sum_{j \in \{0, n\}} B_{Z_j}(\theta^*, \theta_n) - \frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \right) \cdot \frac{\int_{\mathbb{R}^d} \exp \left( \sum_{j \in \{0, n\}} ((\theta^* + \lambda)^\top \nabla Z_j(\theta^* + \lambda^*) - Z_j(\theta^* + \lambda)) \right) d\lambda}{\int_{\mathbb{R}^d} \exp \left( \sum_{j \in \{0, n\}} ((\theta')^\top \nabla Z_j(\theta^* + \lambda^*) - Z_j(\theta')) \right) d\theta'} \\
 &= \frac{c_0}{c_n} \cdot \exp \left( B_{Z_n}(\theta^*, \theta_n) + B_{Z_0}(\theta^*, \theta_n) - \frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \right) \cdot 1 \\
 &= \frac{c_0}{c_n} \cdot \exp \left( \sum_{t=1}^n B_{Z_{s_t, a_t}}(\theta^*, \theta_n) + \frac{\eta}{2} \|\theta^* - \theta_n\|_{\mathbb{A}}^2 - \frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \right) ,
 \end{aligned}$$

where we have introduced  $c_n = \frac{\exp(\sum_{j \in \{0, n\}} ((\theta^* + \lambda^*)^\top \nabla Z_j(\theta^* + \lambda^*) - Z_j(\theta^* + \lambda^*)))}{\int_{\mathbb{R}^d} \exp(\sum_{j \in \{0, n\}} ((\theta')^\top \nabla Z_j(\theta^* + \lambda^*) - Z_j(\theta'))) d\theta'}$ . Since  $\lambda^* = \theta_n - \theta^*$ , we have

$$c_n = \frac{1}{\int_{\mathbb{R}^d} \exp \left( - \sum_{j \in \{0, n\}} B_{Z_j}(\theta', \theta^* + \lambda^*) \right) d\theta'} = \frac{1}{\int_{\mathbb{R}^d} \exp \left( - \sum_{t=1}^n B_{Z_{s_t, a_t}}(\theta', \theta_n) - \frac{\eta}{2} \|\theta' - \theta_n\|_{\mathbb{A}}^2 \right) d\theta'} .$$

Therefore, we have from (5) that

$$C_{\mathbb{A}, n} := \frac{c_n}{c_0} = \frac{\int_{\mathbb{R}^d} \exp \left( - \frac{\eta}{2} \|\theta'\|_{\mathbb{A}}^2 \right) d\theta'}{\int_{\mathbb{R}^d} \exp \left( - \sum_{t=1}^n \text{KL}_{s_t, a_t}(\theta_n, \theta') - \frac{\eta}{2} \|\theta' - \theta_n\|_{\mathbb{A}}^2 \right) d\theta'}$$

An application of Markov's inequality now yields

$$\mathbb{P} \left[ \sum_{t=1}^n \text{KL}_{s_t, a_t}(\theta_n, \theta^*) + \frac{\eta}{2} \|\theta^* - \theta_n\|_{\mathbb{A}}^2 - \frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \geq \log \left( \frac{C_{\mathbb{A}, n}}{\delta} \right) \right] = \mathbb{P} \left[ M_n \geq \frac{1}{\delta} \right] \leq \delta \cdot \mathbb{E} [M_n] = \delta . \tag{12}$$

**Step 3: A Stopped Martingale and Its Control** Let  $N$  be a stopping time with respect to the filtration  $\{\mathcal{F}_n\}_{n=0}^\infty$ . Now, by the martingale convergence theorem,  $M_\infty = \lim_{n \rightarrow \infty} M_n$  is almost surely well-defined, and thus  $M_N$  is well-defined as well irrespective of whether  $N < \infty$  or not. Let  $Q_n = M_{\min\{N, n\}}$  be a stopped version of  $\{M_n\}_n$ . Then an application of Fatou's lemma yields

$$\mathbb{E} [M_N] = \mathbb{E} \left[ \liminf_{n \rightarrow \infty} Q_n \right] \leq \liminf_{n \rightarrow \infty} \mathbb{E} [Q_n] = \liminf_{n \rightarrow \infty} \mathbb{E} [M_{\min\{N, n\}}] \leq 1 ,$$

since the stopped martingale  $\{M_{\min\{N, n\}}\}_{n \geq 1}$  is also a martingale. Therefore, by the properties of  $M_n$ , (12) also holds for any random stopping time  $N < \infty$ .

To complete the proof, we now employ a random stopping time construction as in Abbasi-Yadkori et al. (2011).

We define a random stopping time  $N$  by

$$N = \min \left\{ n \geq 1 : \sum_{t=1}^n \text{KL}_{s_t, a_t}(\theta_n, \theta^*) + \frac{\eta}{2} \|\theta^* - \theta_n\|_{\mathbb{A}}^2 - \frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \geq \log \left( \frac{C_{\mathbb{A}, n}}{\delta} \right) \right\},$$

with  $\min\{\emptyset\} := \infty$  by convention. We then have

$$\mathbb{P} \left[ \exists n \geq 1, \sum_{t=1}^n \text{KL}_{s_t, a_t}(\theta_n, \theta^*) + \frac{\eta}{2} \|\theta^* - \theta_n\|_{\mathbb{A}}^2 - \frac{\eta}{2} \|\theta^*\|_{\mathbb{A}}^2 \geq \log \left( \frac{C_{\mathbb{A}, n}}{\delta} \right) \right] = \mathbb{P}[N < \infty] \leq \delta,$$

which concludes the proof of the first part.

**Proof of Second Part: Upper Bound on  $C_{\mathbb{A}, n}$**  First, we have for some  $\tilde{\theta} \in [\theta_n, \theta']_{\infty}$  that

$$\text{KL}_{s, a}(\theta_n, \theta') = \frac{1}{2} \sum_{i, j=1}^d (\theta' - \theta_n)_i \varphi(s, a)^\top A_i^\top \mathbb{C}_{s, a}^{\tilde{\theta}} [\psi(s')] A_j \varphi(s, a) (\theta' - \theta_n)_j. \quad (13)$$

Now (13) implies that

$$\sum_{t=1}^n \text{KL}_{s_t, a_t}(\theta_n, \theta') \leq \frac{\beta}{2} \sum_{t=1}^n \sum_{i, j=1}^d (\theta' - \theta_n)_i \varphi(s_t, a_t)^\top A_i^\top A_j \varphi(s_t, a_t) (\theta' - \theta_n)_j = \frac{\beta}{2} \|\theta' - \theta_n\|_{\sum_{t=1}^n G_{s_t, a_t}}^2,$$

where  $\beta := \sup_{\theta, s, a} \lambda_{\max}(\mathbb{C}_{s, a}^{\theta}[\psi(s')])$  and  $\forall i, j \leq d$ ,  $(G_{s, a})_{i, j} := \varphi(s, a)^\top A_i^\top A_j \varphi(s, a)$ . Therefore, we obtain

$$\begin{aligned} C_{\mathbb{A}, n} &\leq \frac{\int_{\mathbb{R}^d} \exp\left(-\frac{\eta}{2} \|\theta'\|_{\mathbb{A}}^2\right) d\theta'}{\int_{\mathbb{R}^d} \exp\left(-\frac{1}{2} \|\theta' - \theta_n\|_{(\beta \sum_{t=1}^n G_{s_t, a_t} + \eta \mathbb{A})}^2\right) d\theta'} \\ &= \frac{(2\pi)^{d/2}}{\det(\eta \mathbb{A})^{1/2}} \cdot \frac{\det(\beta \sum_{t=1}^n G_{s_t, a_t} + \eta \mathbb{A})^{1/2}}{(2\pi)^{d/2}} = \det\left(I + \beta \eta^{-1} \mathbb{A}^{-1} \sum_{t=1}^n G_{s_t, a_t}\right), \end{aligned}$$

which completes the proof of the second part.

## C REGRET BOUND OF Exp-UCRL: PROOF OF THEOREM 2

**Step 1: Optimism** Let us consider the start of episode  $t$ , i.e., when the total number of steps completed is  $n = (t-1)H$ . Recall that  $\theta_n \equiv \theta_{(t-1)H}$  denotes the penalized MLE and  $\Theta_n \equiv \Theta_{(t-1)H}$  the confidence set around the MLE after  $n$  steps. Now, let  $\hat{\theta}_n \equiv \hat{\theta}_{(t-1)H}$  denotes the most optimistic realization from the confidence set  $\Theta_n$ , i.e.,

$$V_{\hat{\theta}_n, 1}^{\pi_t}(s_1^t) = \max_{\pi \in \Pi} \max_{\theta \in \Theta_n} V_{\theta, 1}^{\pi}(s_1^t),$$

where  $s_1^t$  denotes the starting state at episode  $t$ . Therefore, as long as the true parameter  $\theta^*$  belongs to  $\Theta_n$ ,  $V_{\hat{\theta}_n, 1}^{\pi_t}(s_1^t)$  gives an optimistic estimate of the value  $V_{\theta^*, 1}^{\pi_t}(s_1^t)$  of the episode, i.e.,

$$V_{\hat{\theta}_n, 1}^{\pi_t}(s_1^t) \geq V_{\theta^*, 1}^{\pi_t}(s_1^t). \quad (14)$$

An application of 1 implies that with probability at least  $1 - \delta/2$ ,  $\theta^* \in \Theta_n$  across all episodes. We then have from (14) that with probability at least  $1 - \delta/2$ , the cumulative regret is controlled by

$$\mathcal{R}(N) \leq \sum_{t=1}^T \left( V_{\hat{\theta}_n, 1}^{\pi_t}(s_1^t) - V_{\theta^*, 1}^{\pi_t}(s_1^t) \right), \quad (15)$$

where  $N = TH$  denotes the total number of steps completed after  $T$  episodes.

**Step 2: Bellman Recursion, Transportation Inequalities and Martingale Control** For any parameter  $\theta \in \mathbb{R}^d$  and policy  $\pi \in \Pi$ , the Bellman operator  $\mathcal{T}_{\theta, h}^{\pi} : (\mathcal{S} \rightarrow \mathbb{R}) \rightarrow (\mathcal{S} \rightarrow \mathbb{R})$  is defined for all  $s \in \mathcal{S}$  and  $h \in [H]$  as

$$\mathcal{T}_{\theta, h}^{\pi}(V)(s) = R(s, \pi(s, h)) + \mathbb{E}_{s, \pi(s, h)}^{\theta}[V],$$

where  $V : \mathcal{S} \rightarrow \mathbb{R}$ . By the Bellman equation, we have

$$V_{\theta, h}^{\pi}(s) = \mathcal{T}_{\theta, h}^{\pi}(V_{\theta, h+1}^{\pi})(s), \quad \forall h \in [H] \quad (\text{with } V_{\theta, H+1}^{\pi}(s) := 0).$$

Following, e.g., Chowdhury and Gopalan (2019), a recursive application of Bellman equation now yields

$$V_{\hat{\theta}_{n,1}}^{\pi_t}(s_1^t) - V_{\theta^*,1}^{\pi_t}(s_1^t) = \sum_{h=1}^H \left( \mathcal{T}_{\hat{\theta}_{n,h}}^{\pi_t} \left( V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right) (s_h^t) - \mathcal{T}_{\theta^*,h}^{\pi_t} \left( V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right) (s_h^t) + m_h^t \right),$$

where  $m_h^t = \mathbb{E}_{s_h^t, a_h^t}^{\theta^*} \left[ V_{\hat{\theta}_{n,h+1}}^{\pi_t}(s_{h+1}^t) - V_{\theta^*,h+1}^{\pi_t}(s_{h+1}^t) \right] - \left( V_{\hat{\theta}_{n,h+1}}^{\pi_t}(s_{h+1}^t) - V_{\theta^*,h+1}^{\pi_t}(s_{h+1}^t) \right)$ . Note that  $\{m_h^t\}_{t,h}$  is a martingale sequence satisfying  $|m_h^t| \leq 2H$ . Therefore, by the Azuma-Hoeffding inequality (Boucheron et al., 2013), with probability at least  $1 - \delta/2$ , we obtain

$$\sum_{t=1}^T \sum_{h=1}^H m_h^t \leq 2H \sqrt{2TH \ln(2/\delta)} = 2H \sqrt{2N \ln(2/\delta)}.$$

Then, using a union bound argument along with (15), the cumulative regret can be upper bounded with probability at least  $1 - \delta$  as

$$\mathcal{R}(N) \leq \sum_{t=1}^T \sum_{h=1}^H \left( \mathcal{T}_{\hat{\theta}_{n,h}}^{\pi_t} \left( V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right) (s_h^t) - \mathcal{T}_{\theta^*,h}^{\pi_t} \left( V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right) (s_h^t) \right) + 2H \sqrt{2N \ln(2/\delta)}. \quad (16)$$

We now proceed to bound the first term in (16). Since  $V_{\hat{\theta}_{n,h+1}}^{\pi_t}(s) \leq H$ ,  $\forall s$ , we have its span  $\mathbb{S} \left( V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right) \leq H$  and variance  $\mathbb{V}_{s_h^t, a_h^t}^{\theta} \left[ V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right] \leq H^2$ ,  $\forall \theta, \forall (s, a)$ . Therefore, we obtain

$$\begin{aligned} & \mathcal{T}_{\hat{\theta}_{n,h}}^{\pi_t} \left( V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right) (s_h^t) - \mathcal{T}_{\theta^*,h}^{\pi_t} \left( V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right) (s_h^t) \\ &= \mathbb{E}_{s_h^t, a_h^t}^{\hat{\theta}_{n,h}} \left[ V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right] - \mathbb{E}_{s_h^t, a_h^t}^{\theta^*} \left[ V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right] \\ &= \mathbb{E}_{s_h^t, a_h^t}^{\hat{\theta}_{n,h}} \left[ V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right] - \mathbb{E}_{s_h^t, a_h^t}^{\theta_{n,h}} \left[ V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right] + \mathbb{E}_{s_h^t, a_h^t}^{\theta_{n,h}} \left[ V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right] - \mathbb{E}_{s_h^t, a_h^t}^{\theta^*} \left[ V_{\hat{\theta}_{n,h+1}}^{\pi_t} \right] \\ &\leq H \sqrt{2 \text{KL}_{s_h^t, a_h^t}(\theta_{n,h}, \hat{\theta}_{n,h})} + H \sqrt{2 \text{KL}_{s_h^t, a_h^t}(\theta_{n,h}, \theta^*)} + \frac{2H}{3} \text{KL}_{s_h^t, a_h^t}(\theta_{n,h}, \theta^*), \end{aligned}$$

where the last step follows from the transportation inequalities (Lemma 1). We then obtain from 16 that

$$\mathcal{R}(N) \leq H \sum_{t=1}^T \sum_{h=1}^H \left( \sqrt{2 \text{KL}_{s_h^t, a_h^t}(\theta_{n,h}, \hat{\theta}_{n,h})} + \sqrt{2 \text{KL}_{s_h^t, a_h^t}(\theta_{n,h}, \theta^*)} + \frac{2}{3} \text{KL}_{s_h^t, a_h^t}(\theta_{n,h}, \theta^*) \right) + 2H \sqrt{2N \ln(2/\delta)}. \quad (17)$$

**Step 3: Sum of KL Divergences Along the Transition Trajectory** First, we obtain from (13) that

$$\forall (s, a) \in \mathcal{S} \times \mathcal{A}, \quad \forall \theta, \theta' \in \mathbb{R}^d, \quad \frac{\alpha}{2} \|\theta' - \theta\|_{G_{s,a}}^2 \leq \text{KL}_{s,a}(\theta, \theta') \leq \frac{\beta}{2} \|\theta' - \theta\|_{G_{s,a}}^2,$$

where  $\alpha := \inf_{\theta, s, a} \lambda_{\min}(\mathbb{C}_{s,a}^{\theta}[\psi(s')])$ ,  $\beta := \sup_{\theta, s, a} \lambda_{\max}(\mathbb{C}_{s,a}^{\theta}[\psi(s')])$ , and  $\forall i, j \leq d$ ,  $(G_{s,a})_{i,j} := \varphi(s, a)^\top A_i^\top A_j \varphi(s, a)$ .

We then have

$$\forall (s, a), \quad \forall \theta, \quad \text{KL}_{s,a}(\theta_n, \theta) \leq \frac{\beta}{2} \|\theta - \theta_n\|_{G_{s,a}}^2 \leq \beta \left\| \bar{G}_n^{-1/2} G_{s,a} \bar{G}_n^{-1/2} \right\| \frac{1}{2} \|\theta - \theta_n\|_{\bar{G}_n}^2,$$

where  $\bar{G}_n \equiv \bar{G}_{(t-1)H} := G_n + \alpha^{-1} \eta \mathbb{A}$  and  $G_n \equiv G_{(t-1)H} := \sum_{\tau=1}^{t-1} \sum_{h=1}^H G_{s_h^\tau, a_h^\tau}$ . Furthermore, note that

$$\frac{1}{2} \|\theta - \theta_n\|_{\bar{G}_n}^2 = \frac{\alpha^{-1} \eta}{2} \|\theta - \theta_n\|_{\mathbb{A}}^2 + \sum_{\tau=1}^{t-1} \sum_{h=1}^H \frac{1}{2} \|\theta - \theta_n\|_{G_{s_h^\tau, a_h^\tau}}^2 \leq \alpha^{-1} \left( \frac{\eta}{2} \|\theta - \theta_n\|_{\mathbb{A}}^2 + \sum_{\tau=1}^{t-1} \sum_{h=1}^H \text{KL}_{s_h^\tau, a_h^\tau}(\theta_n, \theta) \right).$$

Therefore, for any  $\theta \in \Theta_n$ , we obtain

$$\forall (s, a), \quad \text{KL}_{s,a}(\theta_n, \theta) \leq \frac{\beta}{\alpha} \cdot \beta_n(\delta) \left\| \bar{G}_n^{-1/2} G_{s,a} \bar{G}_n^{-1/2} \right\| = \frac{\beta}{\alpha} \cdot \beta_n(\delta) \left\| \bar{G}_n^{-1} G_{s,a} \right\|, \quad (18)$$

where  $\beta_n(\delta) \equiv \beta_{(t-1)H}(\delta) = \frac{\eta}{2} B_{\mathbb{A}}^2 + \log(2C_{\mathbb{A},(t-1)H}/\delta)$ .

Now, since  $G_n$  is positive semi-definite, we have  $\bar{G}_n \succeq \alpha^{-1} \eta \mathbb{A}$ , and thus, in turn

$$\left\| \bar{G}_n^{-1} G_{s,a} \right\| \leq \frac{\alpha}{\eta} \left\| \mathbb{A}^{-1} G_{s,a} \right\| \leq \frac{\alpha B_{\varphi, \mathbb{A}}}{\eta}, \quad \forall (s, a),$$

where  $B_{\varphi, \mathbb{A}} := \sup_{s,a} \|\mathbb{A}^{-1} G_{s,a}\|$ . This further yields

$$\left\| I + \bar{G}_n^{-1} \sum_{h=1}^H G_{s_h^t, a_h^t} \right\| \leq 1 + \sum_{h=1}^H \left\| \bar{G}_n^{-1} G_{s_h^t, a_h^t} \right\| \leq 1 + \frac{\alpha B_{\varphi, \mathbb{A}} H}{\eta}. \quad (19)$$

Now, we define  $\bar{G}_{n+H} := \bar{G}_n + \sum_{h=1}^H G_{s_h^t, a_h^t}$ . Hence,  $\bar{G}_{n+H}^{-1} G_{s,a} = \left( I + \bar{G}_n^{-1} \sum_{h=1}^H G_{s_h^t, a_h^t} \right)^{-1} \bar{G}_n^{-1} G_{s,a}$ . We therefore deduce from (19) that

$$\forall (s, a), \quad \left\| \bar{G}_n^{-1} G_{s,a} \right\| = \left\| \left( I + \bar{G}_n^{-1} \sum_{h=1}^H G_{s_h^t, a_h^t} \right) \bar{G}_{n+H}^{-1} G_{s,a} \right\| \leq \left( 1 + \frac{\alpha B_{\varphi, \mathbb{A}} H}{\eta} \right) \left\| \bar{G}_{n+H}^{-1} G_{s,a} \right\|. \quad (20)$$

Now see that

$$\sum_{t=1}^T \sum_{h=1}^H \left\| \bar{G}_{n+H}^{-1} G_{s_h^t, a_h^t} \right\| \leq \sum_{t=1}^T \sum_{h=1}^H \text{tr} \left( \bar{G}_{n+H}^{-1} G_{s_h^t, a_h^t} \right) = \sum_{t=1}^T \text{tr} \left( \bar{G}_{n+H}^{-1} (\bar{G}_{n+H} - \bar{G}_n) \right) \leq \sum_{t=1}^T \log \frac{\det(\bar{G}_{n+H})}{\det(\bar{G}_n)},$$

where we have used that for two positive definite matrices  $A$  and  $B$  such that  $A - B$  is positive semi-definite,  $\text{tr}(A^{-1}(A - B)) \leq \log \frac{\det(A)}{\det(B)}$ . We can now control the R.H.S. of the above equation, as

$$\sum_{t=1}^T \log \frac{\det(\bar{G}_{n+H})}{\det(\bar{G}_n)} = \sum_{t=1}^T \log \frac{\det(\bar{G}_{tH})}{\det(\bar{G}_{(t-1)H})} = \log \frac{\det(\bar{G}_{TH})}{\det(\bar{G}_0)} = \log \frac{\det(\bar{G}_N)}{\det(\alpha^{-1} \eta \mathbb{A})} = \log \det (I + \alpha \eta^{-1} \mathbb{A}^{-1} G_N).$$

Therefore, we have from (20) and that

$$\sum_{t=1}^T \sum_{h=1}^H \left\| \bar{G}_n^{-1} G_{s_h^t, a_h^t} \right\| \leq \left( 1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta} \right) \log \det (I + \beta \eta^{-1} \mathbb{A}^{-1} G_N), \quad (21)$$

where we have used that  $\alpha \leq \beta$ .

It now remains to bound the log determinant term in the above equation. By the trace-determinant inequality, we have

$$\det (I + \beta \eta^{-1} \mathbb{A}^{-1} G_n) \leq \left( \frac{\text{tr} (I + \beta \eta^{-1} \mathbb{A}^{-1} G_n)}{d} \right)^d \leq \left( 1 + \frac{\beta \eta^{-1}}{d} \text{tr} (\mathbb{A}^{-1} G_n) \right)^d.$$

Now see that  $\text{tr} (\mathbb{A}^{-1} G_n) \leq n \sup_{s,a} \text{tr} (\mathbb{A}^{-1} G_{s,a}) \leq d B_{\varphi, \mathbb{A}} n$ . Therefore, we have

$$\log \det (I + \beta \eta^{-1} \mathbb{A}^{-1} G_n) \leq d \log (1 + \beta \eta^{-1} B_{\varphi, \mathbb{A}} n). \quad (22)$$

This further implies that the confidence radius

$$\beta_n(\delta) \leq \frac{\eta}{2} B_{\mathbb{A}}^2 + \log (2 \det (I + \beta \eta^{-1} \mathbb{A}^{-1} G_n) / \delta) \leq \frac{\eta}{2} B_{\mathbb{A}}^2 + d \log (1 + \beta \eta^{-1} B_{\varphi, \mathbb{A}} n) + \log(2/\delta),$$

which is an increasing function in the total number of steps  $n$ , hence, in the number of episodes  $t$ . We then have from (18) and (21) that

$$\forall \theta \in \Theta_n, \quad \sum_{t=1}^T \sum_{h=1}^H \text{KL}_{s_h^t, a_h^t}(\theta_n, \theta) \leq \frac{\beta}{\alpha} \left( 1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta} \right) \beta_N(\delta) \gamma_N, \quad (23)$$

where we define  $\gamma_N := d \log (1 + \beta \eta^{-1} B_{\varphi, \mathbb{A}} N)$  and  $\beta_N(\delta) := \frac{\eta}{2} B_{\mathbb{A}}^2 + \gamma_N + \log(2/\delta)$ .

**Final Step:** First, an application of Cauchy-Schwartz's inequality yields

$$\forall \theta \in \Theta_n, \quad \sum_{t=1}^T \sum_{h=1}^H \sqrt{\text{KL}_{s_h^t, a_h^t}(\theta_n, \theta)} \leq \sqrt{N \sum_{t=1}^T \sum_{h=1}^H \text{KL}_{s_h^t, a_h^t}(\theta_n, \theta)} \leq \sqrt{\frac{\beta}{\alpha} \left( 1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta} \right) \beta_N(\delta) N \gamma_N}. \quad (24)$$

At this point, we note that by design,  $\hat{\theta}_n \in \Theta_n$  and by Theorem 1,  $\theta^* \in \Theta_n$  with probability at least  $1 - \delta/2$ . We now obtain from (17), (23) and (24) that the cumulative regret

$$\mathcal{R}(N) \leq 2H \sqrt{\frac{\beta}{\alpha} \left( 1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta} \right) 2\beta_N(\delta) N \gamma_N} + 2H \sqrt{2N \ln(2/\delta)} + \frac{2H}{3} \frac{\beta}{\alpha} \left( 1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta} \right) \beta_N(\delta) \gamma_N,$$

which completes the proof.

## D REGRET BOUND OF Exp-PSRL: PROOF OF THEOREM 3

Let us consider the start of episode  $t$ , i.e., when the total number of steps completed is  $n = (t-1)H$ . Recall that we sample  $\tilde{\theta}_n \equiv \tilde{\theta}_{(t-1)H} \sim \mu_n$ , where  $\mu_n \equiv \mu_{(t-1)H} = \mathbb{P}(\theta^* \in \cdot | \mathcal{H}_n)$  denotes the posterior distribution of  $\theta^*$ , given the history of transitions  $\mathcal{H}_n \equiv \mathcal{H}_{(t-1)H} = \{(s_{\tau}^{\top}, a_{\tau}^{\top}, s_{\tau+1}^{\top})_{\tau < t, h \leq H}\}$ . A key property of posterior sampling is that for any  $\sigma(\mathcal{H}_n)$ -measurable function  $f$ , we have  $\mathbb{E}[f(\tilde{\theta}_n)] = \mathbb{E}[f(\theta^*)]$  (Osband et al., 2013). This implies that the optimal policy  $\pi^*$  and selected policy  $\pi^t$  are identically distributed conditioned on the history  $\mathcal{H}_n$ . Therefore, we have  $\mathbb{E}\left[V_{\tilde{\theta}_n, 1}^{\pi^t}(s_1^t)\right] = \mathbb{E}\left[V_{\theta^*, 1}^{\pi^t}(s_1^t)\right]$ , and thus, in turn, the Bayes regret

$$\mathbb{E}[\mathcal{R}(N)] = \mathbb{E}\left[\sum_{t=1}^T \left(V_{\tilde{\theta}_n, 1}^{\pi^t}(s_1^t) - V_{\theta^*, 1}^{\pi^t}(s_1^t)\right)\right].$$

A recursive application of the Bellman equation now yields a result similar to (16):

$$\mathbb{E}[\mathcal{R}(N)] = \mathbb{E}\left[\sum_{t=1}^T \sum_{h=1}^H \left(\mathcal{T}_{\tilde{\theta}_n, h}^{\pi^t} \left(V_{\tilde{\theta}_n, h+1}^{\pi^t}\right)(s_h^t) - \mathcal{T}_{\theta^*, h}^{\pi^t} \left(V_{\theta^*, h+1}^{\pi^t}\right)(s_h^t)\right) + \sum_{t=1}^T \sum_{h=1}^H m_h^t\right],$$

where  $m_h^t = \mathbb{E}_{s_h^t, a_h^t}^{\theta^*} \left[V_{\tilde{\theta}_n, h+1}^{\pi^t}(s_{h+1}^t) - V_{\theta^*, h+1}^{\pi^t}(s_{h+1}^t)\right] - \left(V_{\tilde{\theta}_n, h+1}^{\pi^t}(s_{h+1}^t) - V_{\theta^*, h+1}^{\pi^t}(s_{h+1}^t)\right)$  is a martingale difference sequence satisfying  $\mathbb{E}[m_h^t] = 0$ . Then an application of the transportation inequalities (Lemma 1) yields a result similar to (17):

$$\mathbb{E}[\mathcal{R}(N)] \leq H \mathbb{E}\left[\sum_{t=1}^T \sum_{h=1}^H \left(\sqrt{2 \text{KL}_{s_h^t, a_h^t}(\theta_n, \tilde{\theta}_n)} + \sqrt{2 \text{KL}_{s_h^t, a_h^t}(\theta_n, \theta^*)} + \frac{2}{3} \text{KL}_{s_h^t, a_h^t}(\theta_n, \theta^*)\right)\right], \quad (25)$$

where  $\theta_n \equiv \theta_{(t-1)H}$  denotes the penalized MLE (as computed by Exp-UCRL) after  $n = (t-1)H$  steps.

We now define for any  $\delta \in (0, 1]$ , the events  $\mathcal{E}^* = \{\forall t \geq 1, \theta^* \in \Theta_n\}$  and  $\tilde{\mathcal{E}} = \{\forall t \geq 1, \tilde{\theta}_n \in \Theta_n\}$ , where  $\Theta_n \equiv \Theta_{(t-1)H}$  is confidence set (as constructed by Exp-UCRL) after  $n = (t-1)H$  steps. Under the event  $\mathcal{E}^* \cap \tilde{\mathcal{E}}$ , we have from (23) and (24) that

$$\begin{aligned} \sum_{t=1}^T \sum_{h=1}^H \text{KL}_{s_h^t, a_h^t}(\theta_n, \theta^*) &\leq \frac{\beta}{\alpha} \left(1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta}\right) \beta_N(\delta) \gamma_N, \\ \sum_{t=1}^T \sum_{h=1}^H \sqrt{\text{KL}_{s_h^t, a_h^t}(\theta_n, \theta^*)} &\leq \sqrt{\frac{\beta}{\alpha} \left(1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta}\right) \beta_N(\delta) N \gamma_N} \quad \text{and} \\ \sum_{t=1}^T \sum_{h=1}^H \sqrt{\text{KL}_{s_h^t, a_h^t}(\theta_n, \tilde{\theta}_n)} &\leq \sqrt{\frac{\beta}{\alpha} \left(1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta}\right) \beta_N(\delta) N \gamma_N}. \end{aligned}$$

Therefore, we obtain from (25), the following:

$$\mathbb{E}[\mathcal{R}(N) \mathbb{I}_{\mathcal{E}^* \cap \tilde{\mathcal{E}}}] \leq 2H \sqrt{\frac{\beta}{\alpha} \left(1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta}\right) \beta_N(\delta) N \gamma_N} + \frac{2H}{3} \frac{\beta}{\alpha} \left(1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta}\right) \beta_N(\delta) \gamma_N.$$

Since we can always bound  $\mathcal{R}(N) \leq N$ , we have

$$\mathbb{E}[\mathcal{R}(N)] = \mathbb{E}\left[\mathcal{R}(N) \mathbb{I}_{\mathcal{E}^* \cap \tilde{\mathcal{E}}} + \mathcal{R}(N) \mathbb{I}_{(\mathcal{E}^* \cap \tilde{\mathcal{E}})^c}\right] \leq \mathbb{E}[\mathcal{R}(N) \mathbb{I}_{\mathcal{E}^* \cap \tilde{\mathcal{E}}}] + N(1 - \mathbb{P}(\mathcal{E}^* \cap \tilde{\mathcal{E}})).$$

Now from the property of Posterior sampling,  $\mathbb{P}(\tilde{\mathcal{E}}) = \mathbb{P}(\mathcal{E}^*)$  and from Theorem 1,  $\mathbb{P}(\mathcal{E}^*) \geq 1 - \delta/2$ . Therefore, by a union bound,  $\mathbb{P}(\mathcal{E}^* \cap \tilde{\mathcal{E}}) \geq 1 - \delta$ . This implies for any  $\delta \in (0, 1]$  that the Bayes regret

$$\mathbb{E}[\mathcal{R}(N)] \leq 2H \sqrt{\frac{\beta}{\alpha} \left(1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta}\right) \beta_N(\delta) N \gamma_N} + \frac{2H}{3} \frac{\beta}{\alpha} \left(1 + \frac{\beta B_{\varphi, \mathbb{A}} H}{\eta}\right) \beta_N(\delta) \gamma_N + N\delta.$$

The proof now can be completed by setting  $\delta = 1/N$ .

## E ON THE CHOICE OF PENALTY FUNCTION

In this paper, we have considered the penalty function  $\text{pen}(\theta) = \frac{1}{2} \|\theta\|_{\mathbb{A}}^2$ , where  $\forall i, j \leq d$ ,  $\mathbb{A}_{i,j} = \text{tr}(A_i A_j^{\top})$ . We however note that all our results (Theorem 1, 2, 3) hold for any choice of the (regularizing) matrix  $\mathbb{A}$ . For any



such choice of  $\mathbb{A}$ , we only need to ensure that there exist a known constant  $B_{\mathbb{A}}$  such that  $\|\theta^*\|_{\mathbb{A}} \leq B_{\mathbb{A}}$ . In fact for our particular choice, as we have seen in Section 4, we obtain  $\mathbb{A} = I$  for factored and tabular MDPs and  $\mathbb{A} = m_1 I$  for the linearly controlled dynamical systems. (The scaling with  $m_1$  arises because of our parameterization and can be suppressed for the special case of  $\Sigma_{s,a} = cI$ ,  $c > 0$ ,  $\forall(s, a)$  by using a reparameterization.) We leave it to future work to study the effect of other possible regularizing matrices and penalty functions.