

A Self-normalized Martingale Tail Inequality

The self-normalized martingale tail inequality that we present here is the scalar-valued version of the more general vector-valued results obtained by Abbasi-Yadkori et al. (2011b,a). We include the proof for completeness.

Theorem 7 (Self-normalized bound for martingales). *Let $\{F_t\}_{t=1}^\infty$ be a filtration. Let τ be a stopping time w.r.t. to the filtration $\{F_{t+1}\}_{t=1}^\infty$ i.e. the event $\{\tau \leq t\}$ belongs to F_{t+1} . Let $\{Z_t\}_{t=1}^\infty$ be a sequence of real-valued variables such that Z_t is F_t -measurable. Let $\{\eta_t\}_{t=1}^\infty$ be a sequence of real-valued random variables such that η_t is F_{t+1} -measurable and is conditionally R -sub-Gaussian. Let $V > 0$ be deterministic. Then, for any $\delta > 0$, with probability at least $1 - \delta$,*

$$\frac{(\sum_{t=1}^{\tau} \eta_t Z_t)^2}{V + \sum_{t=1}^{\tau} Z_t^2} \leq 2R^2 \ln \left(\frac{\sqrt{V + \sum_{t=1}^{\tau} Z_t^2}}{\delta \sqrt{V}} \right).$$

Proof. Pick $\lambda \in \mathbb{R}$ and let

$$\begin{aligned} D_t^\lambda &= \exp \left(\frac{\eta_t \lambda Z_t}{R} - \frac{1}{2} \lambda^2 Z_t^2 \right), \\ S_t &= \sum_{s=1}^t \eta_s \lambda Z_s, \\ M_t^\lambda &= \exp \left(\frac{\lambda S_t}{R} - \frac{1}{2} \lambda^2 \sum_{s=1}^t Z_s^2 \right). \end{aligned}$$

We claim that $\{M_t^\lambda\}_{t=1}^\infty$ is an $\{F_{t+1}\}_{t=1}^\infty$ -adapted supermartingale. That $M_t^\lambda \in F_{t+1}$ for $t = 1, 2, \dots$ is clear from the definitions. By sub-Gaussianity, $\mathbf{E}[D_t^\lambda | F_t] \leq 1$. Further,

$$\begin{aligned} \mathbf{E}[M_t^\lambda | F_t] &= \mathbf{E}[M_{t-1}^\lambda D_t^\lambda | F_t] \\ &= M_{t-1}^\lambda \mathbf{E}[D_t^\lambda | F_t] \leq M_{t-1}^\lambda, \end{aligned}$$

showing that $\{M_t^\lambda\}_{t=1}^\infty$ is indeed a supermartingale.

Next we show that M_τ^λ is always well-defined and $\mathbf{E}[M_\tau^\lambda] \leq 1$. First define $\tilde{M} = M_\tau^\lambda$ and note that $\tilde{M}(\omega) = M_{\tau(\omega)}^\lambda(\omega)$. Thus, when $\tau(\omega) = \infty$, we need to argue about $M_\infty^\lambda(\omega)$. By the convergence theorem for nonnegative supermartingales, $\lim_{t \rightarrow \infty} M_t^\lambda(\omega)$ is well-defined, which means M_τ^λ is well-defined, independently of whether $\tau < \infty$ holds or not. Now let $Q_t^\lambda = M_{\min\{\tau, t\}}^\lambda$ be a stopped version of M_t^λ . We proceed by using Fatou's Lemma to show that $\mathbf{E}[M_\tau^\lambda] = \mathbf{E}[\liminf_{t \rightarrow \infty} Q_t^\lambda] \leq \liminf_{t \rightarrow \infty} \mathbf{E}[Q_t^\lambda] \leq 1$.

Let F_∞ be the σ -algebra generated by $\{F_t\}_{t=1}^\infty$ i.e. the tail σ -algebra. Let Λ be a zero-mean Gaussian random variable with variance $1/V$ independent of F_∞ . Define $M_t = \mathbf{E}[M_t^\Lambda | F_\infty]$. Clearly, we still have $\mathbf{E}[M_\tau] = \mathbf{E}[M_\tau^\Lambda] = \mathbf{E}[\mathbf{E}[M_\tau^\Lambda | \Lambda]] \leq \mathbf{E}[1 | \Lambda] \leq 1$.

Let us calculate M_t . We will need the density λ which is $f(\lambda) = \frac{1}{\sqrt{2\pi/V}} e^{-V\lambda^2/2}$. Now, it is easy to write M_t explicitly

$$\begin{aligned} M_t &= \mathbf{E}[M_t^\Lambda | F_\infty] \\ &= \int_{-\infty}^{\infty} M_t^\lambda f(\lambda) d\lambda \\ &= \sqrt{\frac{V}{2\pi}} \int_{-\infty}^{\infty} \exp \left(\frac{\lambda S_t}{R} - \frac{\lambda^2}{2} \sum_{s=1}^t Z_s^2 \right) e^{-V\lambda^2/2} d\lambda \\ &= \exp \left(\frac{S_t^2}{2R^2(V + \sum_{s=1}^t Z_s^2)} \right) \sqrt{\frac{V}{V + \sum_{s=1}^t Z_s^2}}, \end{aligned}$$

where we have used that $\int_{-\infty}^{\infty} \exp(a\lambda - b\lambda^2) = \exp(a^2/(4b))\sqrt{\pi/b}$.

To finish the proof, we use Markov's inequality and the fact that $\mathbf{E}[M_\tau] \leq 1$:

$$\begin{aligned}
 & \Pr \left[\frac{(\sum_{t=1}^{\tau} \eta_t Z_t)^2}{V + \sum_{t=1}^{\tau} Z_t^2} \geq 2R^2 \ln \left(\frac{\sqrt{V + \sum_{t=1}^{\tau} Z_t^2}}{\delta \sqrt{V}} \right) \right] \\
 &= \Pr \left[\frac{S_\tau^2}{2R^2(V + \sum_{t=1}^{\tau} Z_t^2)} \geq \ln \left(\frac{\sqrt{V + \sum_{t=1}^{\tau} Z_t^2}}{\delta \sqrt{V}} \right) \right] \\
 &= \Pr \left[\exp \left(\frac{S_\tau^2}{2R^2(V + \sum_{t=1}^{\tau} Z_t^2)} \right) \geq \frac{\sqrt{V + \sum_{t=1}^{\tau} Z_t^2}}{\delta \sqrt{V}} \right] \\
 &= \Pr \left[M_\tau \geq \frac{1}{\delta} \right] \\
 &\leq \delta.
 \end{aligned}$$

□

The theorem can be “bootstrapped” to a “stronger” statement (or at least one, that looks stronger at the first sight) that holds uniformly for all time steps t as opposed to only a particular (stopping) time τ . The idea of the proof goes back at least to Freedman (1975).

Corollary 8 (Uniform Bound). *Under the same assumptions as the previous theorem, for any $\delta > 0$, with probability at least $1 - \delta$, for all $n \geq 0$,*

$$\left| \sum_{t=1}^n \eta_t Z_t \right| \leq R \sqrt{2 \left(V + \sum_{t=1}^n Z_t^2 \right) \ln \left(\frac{\sqrt{V + \sum_{t=1}^n Z_t^2}}{\delta \sqrt{V}} \right)}.$$

Proof. Define the “bad” event

$$B_t(\delta) = \left\{ \omega \in \Omega : \frac{(\sum_{s=1}^t \eta_s Z_s)^2}{V + \sum_{s=1}^t Z_s^2} > 2R^2 \ln \left(\frac{\sqrt{V + \sum_{s=1}^t Z_s^2}}{\delta \sqrt{V}} \right) \right\}.$$

We are interested in bounding the probability that $\bigcup_{t \geq 0} B_t(\delta)$ happens. Define $\tau(\omega) = \min\{t \geq 0 : \omega \in B_t(\delta)\}$, with the convention that $\min \emptyset = \infty$. Then, τ is a stopping time. Further,

$$\bigcup_{t \geq 0} B_t(\delta) = \{\omega : \tau(\omega) < \infty\}.$$

Thus, by Theorem 7 it holds that

$$\begin{aligned}
 \Pr \left[\bigcup_{t \geq 0} B_t(\delta) \right] &= \Pr [\tau < \infty] \\
 &= \Pr \left[\frac{(\sum_{t=1}^{\tau} \eta_t Z_t)^2}{V + \sum_{t=1}^{\tau} Z_t^2} > 2R^2 \ln \left(\frac{\sqrt{V + \sum_{t=1}^{\tau} Z_t^2}}{\delta \sqrt{V}} \right) \text{ and } \tau < \infty \right] \\
 &= \Pr \left[\frac{(\sum_{t=1}^{\tau} \eta_t Z_t)^2}{V + \sum_{t=1}^{\tau} Z_t^2} > 2R^2 \ln \left(\frac{\sqrt{V + \sum_{t=1}^{\tau} Z_t^2}}{\delta \sqrt{V}} \right) \right] \\
 &\leq \delta.
 \end{aligned}$$

□

B Some Useful Tricks

Proposition 9 (Square-Root Trick). *Let $a, b \geq 0$. If $z^2 \leq a + bz$ then $z \leq b + \sqrt{a}$.*

Proof of the Proposition 9. Let $q(x) = x^2 - bx - a$. The condition $z^2 \leq a + bz$ can be expressed as $q(z) \leq 0$. The quadratic polynomial $q(x)$ has two roots

$$x_{1,2} = \frac{b \pm \sqrt{b^2 + 4a}}{2}.$$

The condition $q(z) \leq 0$ implies that $z \leq \max\{x_1, x_2\}$. Therefore,

$$z \leq \max\{x_1, x_2\} = \frac{b + \sqrt{b^2 + 4a}}{2} \leq b + \sqrt{a},$$

where we have used that $\sqrt{u+v} \leq \sqrt{u} + \sqrt{v}$ holds for any $u, v \geq 0$. □

Proposition 10 (Logarithmic Trick). *Let $c \geq 1$, $f > 0$, $\delta \in (0, 1/4]$. If $z \geq 1$ and $z \leq c + f\sqrt{\ln(z/\delta)}$ then $z \leq c + f\sqrt{2\ln\left(\frac{c+f}{\delta}\right)}$.*

Proof of the Proposition 10. Let $g(x) = x - c - f\sqrt{\ln(x/\delta)}$ for any $x \geq 1$. The condition $z \leq c + f\sqrt{\ln(z/\delta)}$ can be expressed as $g(z) \leq 0$. For large enough x , the function $g(x)$ is increasing. This is easy to see, since $g'(x) = 1 - \frac{f}{2x\sqrt{\ln(x/\delta)}}$. Namely, it is not hard to see $g(x)$ is increasing for $x \geq \max\{1, f/2\}$ since for any such x , $g'(x)$ is positive.

Clearly, $c + f\sqrt{2\ln\left(\frac{c+f}{\delta}\right)} \geq \max\{1, f/2\}$ since $c \geq 1$ and $\delta \in (0, 1/4]$. Therefore, it suffices to show that

$$g\left(c + f\sqrt{2\ln\left(\frac{c+f}{\delta}\right)}\right) \geq 0.$$

This is verified by the following calculation

$$\begin{aligned} g\left(c + f\sqrt{2\ln\left(\frac{c+f}{\delta}\right)}\right) &= c + f\sqrt{2\ln\left(\frac{c+f}{\delta}\right)} - c - f\sqrt{\ln\left(\frac{c + f\sqrt{2\ln((c+f)/\delta)}}{\delta}\right)} \\ &= f\sqrt{2\ln\left(\frac{c+f}{\delta}\right)} - f\sqrt{\ln\left(\frac{c + f\sqrt{2\ln((c+f)/\delta)}}{\delta}\right)} \\ &= f\sqrt{\ln\left(\frac{c+f}{\delta}\right)^2} - f\sqrt{\ln\left(\frac{c + f\sqrt{2\ln((c+f)/\delta)}}{\delta}\right)} \\ &\geq f\sqrt{\ln\left(\frac{c+f}{\delta}\right)^2} - f\sqrt{\ln\left(\frac{(c+f)\sqrt{2\ln((c+f)/\delta)}}{\delta}\right)} \\ &= f\sqrt{\ln(A^2)} - f\sqrt{\ln(A\sqrt{2\ln A})} \\ &\geq 0, \end{aligned}$$

where we have defined $A = (c+f)/\delta$ and the last inequality follows from that $A^2 \geq A\sqrt{2\ln A}$ for any $A > 0$. □

C Proof of Theorem 3

In this section we will need the following notation. For a given positive definite matrix $A \in \mathbb{R}^{d \times d}$ we denote by $\langle x, y \rangle_A = x^\top A y$ the inner product between two vectors $x, y \in \mathbb{R}^d$ induced by A . We denote by $\|x\|_A = \sqrt{\langle x, x \rangle_A} = \sqrt{x^\top A x}$ the corresponding norm.

The following lemma is from Dani et al. (2008). We reproduce the proof for completeness.

Lemma 11 (Elliptical Potential). *Let $x_1, x_2, \dots, x_n \in \mathbb{R}^d$ and let $V_t = I + \sum_{s=1}^t x_s^\top x_s$ for $t = 0, 1, 2, \dots, n$. Then it holds that*

$$\sum_{t=1}^n \min \left\{ 1, \|x_t\|_{V_{t-1}}^2 \right\} \leq 2 \ln(\det(V_n)) .$$

Furthermore, if $\|x_t\|_2 \leq X$ for all $t = 1, 2, \dots, n$ then

$$\ln(\det(V_n)) \leq d \ln \left(1 + \frac{nX^2}{d} \right) .$$

Proof of Lemma 11 . We use the inequality $x \leq 2 \ln(1 + x)$ valid for all $x \in [0, 1]$:

$$\sum_{t=1}^n \min \left\{ 1, \|x_t\|_{V_{t-1}}^2 \right\} \leq \sum_{t=1}^n 2 \ln \left(1 + \|x_t\|_{V_{t-1}}^2 \right) = 2 \ln \left(\prod_{t=1}^n \left(1 + \|x_t\|_{V_{t-1}}^2 \right) \right) .$$

We now show that $\det(V_n) = \prod_{t=1}^n (1 + \|x_t\|_{V_{t-1}}^2)$:

$$\begin{aligned} \det(V_n) &= \det(V_{n-1} + x_n x_n^\top) \\ &= \det \left(V_{n-1} (I + (V_{n-1}^{-1/2} x_n) (V_{n-1}^{-1/2} x_n)^\top) \right) \\ &= \det(V_{n-1}) \det \left(I + (V_{n-1}^{-1/2} x_n) (V_{n-1}^{-1/2} x_n)^\top \right) \\ &= \det(V_{n-1}) \cdot \left(1 + \|x_n\|_{V_{n-1}}^2 \right) \\ &= \dots \\ &= \prod_{t=1}^n (1 + \|x_t\|_{V_{t-1}}^2) . \end{aligned} \quad (\text{since } V_0 = I)$$

In the above calculation we have used that $\det(I + z z^\top) = 1 + \|z\|_2^2$ since all but one eigenvalue of $I + z z^\top$ equals to 1 and the remaining eigenvalue is $1 + \|z\|_2^2$ with associated eigenvector z .

To prove the second part, consider the eigenvalues $\alpha_1, \alpha_2, \dots, \alpha_d$ of V_n . Since V_n is positive definite, the eigenvalues are positive. Recall that $\det(V_n) = \prod_{i=1}^d \alpha_i$. The bound on $\|x_t\| \leq X$ implies a bound on the trace of V_n :

$$\text{Trace } V_n = \text{Trace}(I) + \sum_{t=1}^n \text{Trace}(x_t x_t^\top) = d + \sum_{t=1}^n \|x_t\|_2^2 \leq d + nX^2 .$$

Recalling that $\text{Trace}(V_n) = \sum_{i=1}^d \alpha_i$ we can apply the AM-GM inequality:

$$\sqrt[d]{\alpha_1 \alpha_2 \dots \alpha_d} \leq \frac{\alpha_1 + \alpha_2 + \dots + \alpha_d}{d} = \frac{\text{Trace}(V_n)}{d} ,$$

from which the second inequality follows by taking logarithm and multiplying by d . □

Proof of Theorem 3. Consider the event A when $\theta_* \in \bigcap_{t=0}^\infty C_t$. By Corollary 2, the event A occurs with probability at least $1 - \delta$.

The set C_{t-1} is an ellipsoid underlying the covariance matrix $V_{t-1} = I + \sum_{s=1}^{t-1} X_s^\top X_s$ and center

$$\hat{\theta}_t = \underset{\theta \in \mathbb{R}^d}{\text{argmin}} \left(\|\theta\|_2^2 + \sum_{s=1}^{t-1} (\hat{Y}_s - \langle \theta, X_s \rangle)^2 \right) .$$

The ellipsoid C_{t-1} is non-empty since θ_* lies in it (on the event A). Therefore $\hat{\theta}_t \in C_{t-1}$. We can thus express the ellipsoid as

$$C_{t-1} = \left\{ \theta \in \mathbb{R}^d : (\theta - \hat{\theta}_t)^\top V_{t-1} (\theta - \hat{\theta}_t) + \|\hat{\theta}_t\|_2^2 + \sum_{s=1}^{t-1} \left(\hat{Y}_s - \langle \hat{\theta}_t, X_s \rangle \right)^2 \leq \beta_{t-1}(\delta) \right\} .$$

The ellipsoid is contained in a larger ellipsoid

$$C_{t-1} \subseteq \left\{ \theta \in \mathbb{R}^d : (\theta - \hat{\theta}_t)^\top V_{t-1} (\theta - \hat{\theta}_t) \leq \beta_{t-1}(\delta) \right\} = \left\{ \theta \in \mathbb{R}^d : \|\theta - \hat{\theta}_t\|_{V_{t-1}} \leq \sqrt{\beta_{t-1}(\delta)} \right\}.$$

First, we bound the instantaneous regret using that $(X_t, \tilde{\theta}_t) = \operatorname{argmax}_{(x, \theta) \in D_t \times C_{t-1}} \langle x, \theta \rangle$:

$$\begin{aligned} \langle x_* - X_t, \theta_* \rangle &= \langle x_*, \theta_* \rangle - \langle X_t, \theta_* \rangle \\ &\leq \langle X_t, \tilde{\theta}_t \rangle - \langle X_t, \theta_* \rangle \\ &= \langle X_t, \tilde{\theta}_t - \theta_* \rangle \\ &= \langle X_t, \tilde{\theta}_t - \hat{\theta}_t \rangle - \langle X_t, \hat{\theta}_t - \theta_* \rangle \\ &\leq \left| \langle X_t, \tilde{\theta}_t - \hat{\theta}_t \rangle \right| + \left| \langle X_t, \hat{\theta}_t - \theta_* \rangle \right| \\ &\leq \|X_t\|_{V_{t-1}^{-1}} \left\| \tilde{\theta}_t - \hat{\theta}_t \right\|_{V_{t-1}} + \|X_t\|_{V_{t-1}^{-1}} \left\| \hat{\theta}_t - \theta_* \right\|_{V_{t-1}} && \text{(Cauchy-Schwarz)} \\ &\leq 2\sqrt{\beta_{t-1}(\delta)} \cdot \|X_t\|_{V_{t-1}^{-1}}. && \text{(because } \tilde{\theta}_t, \theta_* \in C_{t-1}) \end{aligned}$$

Since we assume that $|\langle x, \theta_* \rangle| \leq G$ for any $x \in D_t$ and any $t = 1, 2, \dots, n$, we can upper bound $\langle x_* - X_t, \theta_* \rangle \leq 2 \min\{G, \sqrt{\beta_{t-1}(\delta)} \cdot \|X_t\|_{V_{t-1}^{-1}}\}$. Summing over all t we upper bound regret

$$\begin{aligned} R_n &= \sum_{t=1}^n \langle x_* - X_t, \theta_* \rangle \\ &\leq 2 \sum_{t=1}^n \min \left\{ G, \sqrt{\beta_{t-1}(\delta)} \cdot \|X_t\|_{V_{t-1}^{-1}} \right\} \\ &\leq 2 \sum_{t=1}^n \sqrt{\beta_{t-1}(\delta)} \cdot \min \left\{ G, \|X_t\|_{V_{t-1}^{-1}} \right\} && \text{(since } \beta_{t-1}(\delta) \geq 1) \\ &\leq 2 \left(\max_{0 \leq t < n} \sqrt{\beta_t(\delta)} \right) \sum_{t=1}^n \min \left\{ G, \|X_t\|_{V_{t-1}^{-1}} \right\} \\ &\leq 2 \left(\max_{0 \leq t < n} \sqrt{\beta_t(\delta)} \right) \max\{1, G\} \sum_{t=1}^n \min \left\{ 1, \|X_t\|_{V_{t-1}^{-1}} \right\} \\ &\leq 2 \left(\max_{0 \leq t < n} \sqrt{\beta_t(\delta)} \right) \max\{1, G\} \times \sqrt{n \sum_{t=1}^n \min \left\{ 1, \|X_t\|_{V_{t-1}^{-1}}^2 \right\}} && \text{(Cauchy-Schwarz)} \\ &\leq 2 \max\{1, G\} \sqrt{2nd \log \left(1 + \frac{nX^2}{d} \right) \max_{0 \leq t < n} \beta_t(\delta)}, \end{aligned}$$

where the last inequality follows from Lemma 11. □

Proof of Theorem 4. Summing over all t we upper bound regret

$$R_n = \sum_{t=1}^n \langle x_* - X_t, \theta_* \rangle \leq \frac{1}{\Delta} \sum_{t=1}^n \langle x_* - X_t, \theta_* \rangle^2,$$

where the last inequality follows from the fact that either $\langle x_* - X_t, \theta_* \rangle = 0$ or $\langle x_* - X_t, \theta_* \rangle > \Delta$. Then we take

similar steps as in the proof of Theorem 3 to obtain

$$\begin{aligned}
 R_n &\leq \frac{1}{\Delta} \sum_{t=1}^n \langle x^* - X_t, \theta_* \rangle^2 \\
 &\leq \frac{4}{\Delta} \left(\max_{0 \leq t < n} \beta_t(\delta) \right) \max\{1, G^2\} \sum_{t=1}^n \min \left\{ 1, \|X_t\|_{V_{t-1}^{-1}}^2 \right\} \\
 &\leq \frac{8d}{\Delta} \left(\max_{0 \leq t < n} \beta_t(\delta) \right) \max\{1, G^2\} \log \left(1 + \frac{nX^2}{d} \right),
 \end{aligned}$$

finishing the proof of the problem dependent bound. □