

# Robust Multi-task Regression

## – Supplementary Material

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We provide in this supplementary material proofs that are omitted in the submission.

### A Proof of Lemma 5

*Proof.* By Condition 2 we have

$$\sqrt{\phi_{\min}(S + S')} \geq 4\sqrt{(\phi_{\max}(S'))(S/S')},$$

which leads to

$$\begin{aligned} \kappa_1 &= \frac{\sqrt{\phi_{\min}(S + S')} - 3\sqrt{\phi_{\max}(S')S/S'}}{1 + 3\sqrt{S/S'}} \\ &\geq \frac{\sqrt{\phi_{\min}(S + S')}}{16} \geq \frac{\sqrt{(\phi_{\max}(S'))(S/S')}}{4}. \end{aligned}$$

This established the first claim. Next, note that

$$\begin{aligned} \kappa_2/\kappa_1 &= \frac{\frac{\sqrt{\phi_{\min}(S+S')} - 3\sqrt{\phi_{\max}(S')S/S'}}{\sqrt{S'} + 3\sqrt{S}} + \sqrt{\phi_{\max}(S')/S'}}{\frac{\sqrt{\phi_{\min}(S+S')} - 3\sqrt{\phi_{\max}(S')S/S'}}{1 + 3\sqrt{S/S'}}} \\ &= \frac{1}{\sqrt{S'}} + \frac{(1 + 3\sqrt{S/S'})\sqrt{\phi_{\max}(S')/S'}}{\sqrt{\phi_{\min}(S + S')} - 3\sqrt{\phi_{\max}(S')S/S'}} \\ &\leq \frac{1}{\sqrt{S'}} + \frac{(1 + 3\sqrt{S/S'})\sqrt{\phi_{\max}(S')/S'}}{\sqrt{\phi_{\max}(S')S/S'}} \\ &= \frac{4}{\sqrt{S'}} + \frac{1}{\sqrt{S}} \leq \frac{5}{\sqrt{S}}, \end{aligned}$$

where the last inequality holds from  $S' \geq S$ . □

## B Proof of Lemma 6

*Proof.* To prove the lemma, we need the following two results. Lemma B.1 is adapted from Theorem 2.13 of [1], and Lemma B.2 is from [2].

**Lemma B.1.** *Let  $\Gamma$  be an  $n \times q$  matrix, whose entries are all i.i.d.  $\mathcal{N}(0, 1)$  Gaussian variables; then*

$$\Pr(\|\Gamma\|_{op} > \sqrt{n} + \sqrt{q} + \epsilon) \leq \exp(-\epsilon^2/2).$$

**Lemma B.2.** *Let  $U$  be a  $\chi^2$  statistic with  $D$  degrees of freedom. For any positive  $x$ ,*

$$\Pr(U - D \geq 2\sqrt{Dx} + 2x) \leq \exp(-x).$$

We now proceed to prove Lemma 6. Note that  $\|X^\top W\|_{\infty, 2} \leq \max_{i=1, \dots, p} \|X_i\|_2 \|W\|_{op} \leq \|W\|_{op}$ . Note that Lemma B.1 implies that for  $W$  with i.i.d.  $\mathcal{N}(0, \sigma^2/n)$  entries, we have with probability

$$\Pr(\|W\|_{op} \geq 7\sigma(1 + \sqrt{q/n})) \leq \exp(-6(n+q)/2) \leq 1/(4n^3), \quad (\text{B.1})$$

where the last inequality holds when  $n \geq 2$ . On the other hand, fix a  $X_i^\top$ ,  $X_i^\top W_j \sim \mathcal{N}(0, \|X_i\|_2^2 \sigma^2/n)$ , and hence

$$\frac{\sqrt{n}}{\sigma \|X_i\|_2} \|X_i' W\|_2 \sim \chi_{q-1}.$$

By Lemma B.2, we have

$$\begin{aligned} & \Pr\left(\|X_i' W\|_2 \geq \frac{\sigma(\sqrt{q} + \sqrt{8 \log(pn)})}{\sqrt{n}}\right) \\ & \stackrel{(a)}{\leq} \Pr\left(\|X_i' W\|_2 \geq \frac{\sigma \|X_i\|_2 (\sqrt{q} + \sqrt{2 \log(p^4 n^4)})}{\sqrt{n}}\right) \\ & = \Pr\left(\frac{n \|X_i' W\|_2^2}{\sigma^2 \|X_i\|_2^2} \geq q + 2\sqrt{2q \log p^4 n^4} + 2 \log(p^4 n^4)\right) \\ & \leq \Pr\left(\frac{n \|X_i' W\|_2^2}{\sigma^2 \|X_i\|_2^2} - q \geq 2\sqrt{q \log(p^4 n^4)} + 2 \log(p^4 n^4)\right) \\ & \stackrel{(b)}{\leq} 1/(n^4 p^4) \leq 1/(4n^3 p), \end{aligned}$$

which by union bound leads to

$$\Pr(\|X^\top W\|_{\infty, 2} \geq \frac{\sigma(\sqrt{q} + \sqrt{2 \log(pn)})}{\sqrt{n}}) \leq 1/(4n^3). \quad (\text{B.2})$$

Here, (a) holds since  $\|X_i\|_2 \leq 1$  and (b) holds by applying to Lemma B.2 to  $\chi^2$  random variable  $\frac{n \|X_i' W\|_2^2}{\sigma^2 \|X_i\|_2^2}$ . Lemma 6 follows by a union bound on Equation (B.1) and (B.2).  $\square$

## C Proof of Example 1

In this appendix we show the validity of the claim in Example 1. Recall the following theorem from [1] (Thm. 2.13).

**Theorem C.1.** *Given  $m, n \in \mathbb{N}$  with  $m \leq n$ , put  $\Theta = m/n$  and consider the  $n \times m$  random matrix  $\Gamma$  whose entries are real, independent Gaussian random variable following  $\mathcal{N}(0, 1/n)$ . Let the singular value be  $\sigma_1(\Gamma) \geq \dots \geq \sigma_m(\Gamma)$ . Then for any  $t > 0$ ,*

$$Pr(\sigma_1(\Gamma) \geq 1 + \sqrt{\Theta} + t) \leq \exp(-nt^2/2); \quad Pr(\sigma_m(\Gamma) \leq 1 - \sqrt{\Theta} - t) \leq \exp(-nt^2/2).$$

Then, fix any  $I \subset [1 : p]$  with  $|I| = S + S'$  and denote  $M = X_I$  we have that  $M$  is a  $(S + S') \times n$  random matrix with IID entries following  $\mathcal{N}(0, 1/n)$ . Thus, for any  $t > 0$ ,

$$\begin{aligned} & Pr(\phi_{\min}(S + S') < 1 - \sqrt{(S + S')/n} - t) \\ & \leq \binom{p}{S + S'} \times Pr(\sigma_{S+S'}(M) < 1 - \sqrt{(S + S')/n} - t) \\ & \leq p^{S+S'} \exp(-nt^2/2) = \exp\left(\frac{2(S + S') \log p - nt^2}{2}\right), \end{aligned}$$

where the first inequality follows from a union bound on all subsets with cardinality  $S + S'$ . Similarly

$$Pr(\phi_{\max}(S') > 1 + \sqrt{S'/n} + t) \leq \exp\left(\frac{2S' \log p - nt^2}{2}\right)$$

Set  $t = \sqrt{4(S + S') \log p/n}$  we have that with probability at least  $1 - 1/p^2$ ,

$$\begin{aligned} \phi_{\min}(S + S') & \geq 1 - \sqrt{\frac{S + S'}{n}} - \sqrt{\frac{4(S + S') \log p}{n}} \geq 1 - \sqrt{\frac{9(S + S') \log p}{n}}; \\ \phi_{\max}(S') & \leq 1 + \sqrt{\frac{S'}{n}} + \sqrt{\frac{4(S + S') \log p}{n}} \leq 1 + \sqrt{\frac{9(S + S') \log p}{n}}. \end{aligned}$$

When  $S \log p \leq n/1764$  and  $S' = 48S$ , we have that  $(S + S') \log p/n \leq 1/36$  which implies that  $\phi_{\min}(S + S') \geq 1/2$  and  $\phi_{\max}(S') \leq 3/2$ . Hence

$$S' \phi_{\min}(S + S') \geq 16S \phi_{\max}(S'),$$

as claimed.

## References

- [1] K. Davidson and S. Szarek. Local operator theory, random matrices and banach spaces. In W. Johnson and J. Lindenstrauss, editors, *Handbook on the Geometry of Banach Spaces*, pages 317–366. Elsevier, 2001.
- [2] B. Laurent and P. Massart. Adaptive estimation of a quadratic functional of by model selection. *Annals of Statistics*, 28:1302–1338, 2000.