Supplementary Material: Scalable Bayesian Low-Rank Decomposition of Incomplete Multiway Tensors

In our MGP based CP decomposition, the tensor element x_i , with $i = [i_1, i_2, \dots, i_K]$ its K-dimensional index vector, can be concisely represented as:

$$x_{i} = \sum_{r=1}^{R} \lambda_r \prod_{k=1}^{K} u_{i_k r}^{(k)}$$

where

$$\begin{aligned} \lambda_r &\sim \mathcal{N}(0, \tau_r^{-1}), \quad 1 \le r \le R \\ \tau_r &= \prod_{l=1}^r \delta_l, \quad \delta_l \sim Ga(a_c, 1) \quad a_c > 1 \quad 1 < r \le R \\ \boldsymbol{u}_r^{(k)} &\sim \mathcal{N}(\boldsymbol{\mu}^{(k)}, \boldsymbol{\Sigma}^{(k)}), \ 1 < r \le R, \ 1 < k \le K \end{aligned}$$

Theorem 1. With $a_c > 1$, the sequence $\sum_{r=1}^R \lambda_r \prod_{k=1}^K u_{i_k r}^{(k)}$ converges in ℓ_2 as $R \to \infty$. **Proof:** To prove Theorem 1, we will make use of the following Lemma.

Lemma 1. A sequence of random variables $\{X_n\}_{n\geq 1}$ converges to a random variable X ($\mathbb{E}(X^2) < \infty$) in ℓ_2 iff $\mathbb{E}(X_{n+k} - X_n)^2 \to 0$ as $n, k \to \infty$.

Based on the Lemma 1, to prove Theorem 1, we need to show that $\sum_{r=R+1}^{\infty} \mathbb{E}(\lambda_r \prod_{k=1}^{K} u_{i_k r}^{(k)})^2 = 0.$

Note that, $\forall r$, in our MGP-CP construction λ_r and $\{u_{i_k r}^{(k)}\}_{k=1}^K$ are independent of each other, and $\mathbb{E}(\lambda_r) = 0$. Therefore we have

$$\sum_{r=R+1}^{\infty} \mathbb{E}(\lambda_r \prod_{k=1}^{K} u_{i_k r}^{(k)})^2 = \sum_{r=R+1}^{\infty} \mathbb{E}(\lambda_r)^2 \prod_{k=1}^{K} \mathbb{E}(u_{i_k r}^{(k)})^2$$
(1)

where $\mathbb{E}(\lambda_r)^2 = \mathbb{E}(\mathbb{E}(\lambda_r)^2 | \tau_r)) = \mathbb{E}(\frac{1}{\prod_{l=1}^r \delta_l}) = \frac{1}{a_c^r}$.

Under the assumption that the vectors $u_{i_k}^{(k)}$ have bounded variances, $\mathbb{E}(u_{i_k r}^{(k)})^2$ will be bounded by $\zeta_k < \infty$. Substituting these in Equation 1, we have

$$\sum_{k=R+1}^{\infty} \mathbb{E}(\lambda_r \prod_{k=1}^{K} u_{i_k r}^{(k)})^2 < \prod_{k=1}^{K} \zeta_k \sum_{r=R+1}^{\infty} \frac{1}{a_c^r} = \prod_{k=1}^{K} \zeta_k \frac{1}{a_c^R(a_c - 1)}$$

Taking the limit $R \to \infty$

$$\lim_{R \to \infty} \sum_{r=R+1}^{\infty} \mathbb{E}(\lambda_r \prod_{k=1}^{K} u_{i_k r}^{(k)})^2 = \lim_{R \to \infty} \frac{1}{a_c^R(a_c - 1)} \prod_{k=1}^{K} \zeta_k = 0$$

Thus the sequence $\sum_{r=1}^{R} \lambda_r \prod_{k=1}^{K} u_{i_k r}^{(k)}$ converges in ℓ_2 as $R \to \infty$, which completes the proof of Theorem 1.

Although the MGP construction is originally unbounded [1], a consequence of Theorem 1 is that even when a finite truncation level R is used, the approximation error due truncation R decreases exponentially fast with increasing R, as given by the following:

Theorem 2. Denote the residual by $M_{i_1 i_2 \dots i_k}^R = \sum_{r=R+1}^{\infty} \lambda_r \prod_{k=1}^K u_{i_k r}^{(k)}$. Then $\forall \epsilon > 0$ we have $P\{(M_{i_1 i_2 \dots i_k}^R)^2 > \epsilon\} < \frac{1}{\epsilon a_c^R(a_c-1)} \prod_{k=1}^K \zeta_k$

Proof: We have $P\{(M_{i_{1}i_{2}...i_{k}}^{R})^{2} \leq \epsilon\} = \mathbb{E}(P\{(M_{i_{1}i_{2}...i_{k}}^{R})^{2} \leq \epsilon | \tau\}) = 1 - \mathbb{E}(P\{(M_{i_{1}i_{2}...i_{k}}^{R})^{2} > \epsilon | \tau\}) = 1 - \frac{\mathbb{E}(\mathbb{E}((M_{i_{1}i_{2}...i_{k}}^{R})^{2} | \tau))}{\epsilon} \text{ where } \mathbb{E}(\mathbb{E}((M_{i_{1}i_{2}...i_{k}}^{R})^{2} | \tau)) < \frac{1}{a_{c}^{R}(a_{c}-1)} \prod_{k=1}^{K} \zeta_{k} \text{ as shown in the proof of Theorem 1.}$ Thus we have $P\{(M_{i_{1}i_{2}...i_{k}}^{R})^{2} > \epsilon\} < \frac{1}{\epsilon a_{c}^{R}(a_{c}-1)} \prod_{k=1}^{K} \zeta_{k}.$

References

[1] A. Bhattacharya and D. Dunson. Sparse bayesian infinite factor models. Biometrika, 2011.