# A. Appendix

In this section, we provide the proofs of several technical results that are claimed or used in our main paper.

### A.1. Proof of Proposition 4.1

The proof follows via a reduction from the so-called SUBSETSUM problem, which is known to be NP-hard (Garey & Johnson, 1979). Recall that the SUBSETSUM decision problem is as follows: given n numbers,  $a_1, \ldots, a_n$  in  $\mathbb{R}$ , decide if there exists a partition  $S \subseteq [n]$  such that

$$\sum_{i \in S} a_i = \sum_{j \in S^c} a_j.$$

We show that if we can solve the mixed linear equations problem in polynomial time, then we can solve the SUBSETSUM problem, which would thus imply that P = NP.

Given  $\{a_1, \ldots, a_n\}$ , we must design a matrix X, and output variable y, such that if we could solve the mixed linear equation problem specified by (y, X), then we could decide the subset sum problem on  $\{a_1, \ldots, a_n\}$ . To this end, we define:

$$X = \begin{bmatrix} I_n & \\ I_n & \\ a_1 & \cdots & a_n \end{bmatrix}, \qquad \mathbf{y} = \begin{pmatrix} \mathbf{1}_{n \times 1} \\ \mathbf{0}_{n \times 1} \\ \sum_i a_i/2 \end{pmatrix}.$$

Here,  $I_n$  denotes the  $n \times n$  identity matrix,  $\mathbf{1}_{n \times 1}$  the  $n \times 1$  vector of 1's, and similarly,  $\mathbf{0}_{n \times 1}$  the  $n \times 1$  vector of 0's. Finding a solution to the mixed linear equations problem amounts to finding a subset  $S \subseteq [2n + 1]$  of the 2n + 1 constraints, and vectors  $\beta^{(1)}, \beta^{(2)} \in \mathbb{R}^n$ , so that  $\beta^{(1)}$  satisfies the equalities  $X_S\beta^{(1)} = \mathbf{y}_S$ , and  $\beta^{(2)}$  the equalities  $X_{S^c}\beta_2 = \mathbf{y}_{S^c}$ . Note that S cannot contain i and n + i, since these equalities are mutually exclusive. The consequence is that we have  $\beta_i^{(1)} \in \{0, 1\}$ , with  $\beta_i^{(1)} = 1 - \beta_i^{(2)}$ . Thus if the first 2n constraints are satisfied, the final constraint, therefore, can only be satisfied if we have

$$\sum_{i \in S} a_i = \sum_i a_i \beta_i^{(1)} = \sum_j a_j \beta_j^{(2)} = \sum_{j \in S^c} a_j,$$

thus proving the result.

#### A.2. Proof of Proposition 4.2

To show that our SVD initialization produces a good initial solution, requires two steps. Recall that Algorithm 5 finds the two dimensional subspace spanned by the top two eigenvectors of the matrix  $M = \frac{1}{|\mathcal{S}_*|} \sum_{i \in \mathcal{S}_*} y_i^2 \mathbf{x}_i \otimes \mathbf{x}_i$ , and then searches on a discretization of the circle in that subspace for two vectors that minimize the loss function,  $\mathcal{L}_+$  evaluated on the samples in  $\mathcal{S}_+$ .

We first show that the top eigenspace of M is indeed close to the top eigenspace of its expectation,  $p_1\beta_1^*\otimes\beta_1^*+p_2\beta_2^*\otimes\beta_2^*+I$ , i.e., it is close to span $\{\beta_1^*, \beta_2^*\}$ , and that some pair of elements of the discretization are close to  $(\beta_1^*, \beta_2^*)$ . This is the content of lemma A.1. We then show that our loss function  $\mathcal{L}_+$  is able to select good points from the discretization.

Our algorithm then uses the loss function  $\mathcal{L}_+$  (evaluated on new samples in  $\mathcal{S}_+$ ) to select good points from the grid G. Lemma A.2 shows that as long as the number  $\mathcal{S}_+$  of these new samples is large enough, we can upper *and* lower bound, with high probability, the empirically evaluated loss  $\mathcal{L}_+(\hat{\beta}_1, \hat{\beta}_2)$  of any candidate pair  $\hat{\beta}_1, \hat{\beta}_2$  by the true error err of that candidate pair. This provides the critical result allowing us to do the correct selection in the 1-d search phase.

Now we are ready to prove the result. Suppose the conditions of lemma A.1 hold. Then we are guaranteed the existence of  $(\bar{\beta}_1, \bar{\beta}_2)$  in the grid G with  $\delta$ -resolution, such that  $\max_i \|\bar{\beta}_i - \beta_i^*\| < \delta$ . Next, let  $(\beta_1^{(0)}, \beta_2^{(0)})$  be the output of our SVD initialization, and let err denote their distance from  $(\beta_1^*, \beta_2^*)$ . By definition, the vectors  $(\beta_1^{(0)}, \beta_2^{(0)})$  minimize the loss function  $\mathcal{L}_+$  taken on inputs  $\mathcal{S}_+$ , and hence  $\mathcal{L}_+(\beta_1^{(0)}, \beta_2^{(0)}) \leq \mathcal{L}_+(\bar{\beta}_1, \bar{\beta}_2)$ . Using the lower bound from lemma A.2, applied to  $(\beta_1^{(0)}, \beta_2^{(0)})$  we have:

$$\frac{1}{5}\sqrt{\min\{p_1, p_2\}} \operatorname{err} \le \sqrt{\frac{\mathcal{L}_+(\beta_1^{(0)}, \beta_2^{(0)})}{|\mathcal{S}_+|}}.$$

From the upper bound applied to  $(\bar{\beta}_1, \bar{\beta}_2)$ , we have

$$\sqrt{\frac{\mathcal{L}_+(\bar{\beta}_1,\bar{\beta}_2)}{|\mathcal{S}_+|}} \le 1.1\delta$$

Recalling that  $\mathcal{L}_+(\beta_1^{(0)},\beta_2^{(0)}) \leq \mathcal{L}_+(\bar{\beta}_1,\bar{\beta}_2),$  and taking

$$\delta \leq \frac{2}{11} \widehat{c} \|\beta_1^* - \beta_2^*\|_2 \sqrt{\min\{p_1, p_2\}}^3,$$

we combine to finally obtain:

$$\operatorname{err} \leq \frac{11}{2} \frac{\delta}{\sqrt{\min\{p_1, p_2\}}} \\ \leq \widehat{c} \min\{p_1, p_2\} \|\beta_1^* - \beta_2^*\|_2.$$

where  $\hat{c}$  is as in the statement of proposition 4.2.

# A.3. Proof of Proposition 4.3

Using standard concentration results, in lemma A.1, we have shown if

$$|\mathcal{S}_*| > c(1/\widetilde{\delta})^2 k \log^2 k,$$

with probability at least  $1 - \frac{1}{k^2}$ ,

$$\|M - \mathbb{E}(M)\| < 3\delta$$

Hence, we have

$$\left|\left|\lambda_1^* - \lambda_2^*\right| - \left|\lambda_1 - \lambda_2\right|\right| \le 6\widetilde{\delta}.$$

The approximate error of  $\Delta_b^*$  can be bounded as:

$$\begin{aligned} 2p_b |\Delta_b^* - \Delta_b| &\leq 6\widetilde{\delta} + (p_b^2 - p_{-b}^2) \left[ \frac{1}{\lambda_{-b}^* - \lambda_b^*} - \frac{1}{\lambda_{-b} - \lambda_b} \right] \\ &\leq 6\widetilde{\delta} + |p_b^2 - p_{-b}^2| \frac{6\widetilde{\delta}}{(\lambda_{-b}^* - \lambda_b^*)(\lambda_{-b} - \lambda_b)} \\ &\leq 6\widetilde{\delta} + |p_b^2 - p_{-b}^2| \frac{6\widetilde{\delta}}{|\lambda_{-b}^* - \lambda_b^*|(|\lambda_{-b}^* - \lambda_b^*| - 6\widetilde{\delta})} \\ &\leq 6\widetilde{\delta} + |p_b^2 - p_{-b}^2| \frac{12\widetilde{\delta}}{|\lambda_{-b}^* - \lambda_b^*|^2} \end{aligned}$$

In the last inequality we use  $\widetilde{\delta} \leq \frac{|\lambda_1^* - \lambda_2^*|}{12}$ .

Next, we calculate approximation error of eigenvectors. Note that  $\mathbb{E}(\frac{M-I}{2}) = p_1\beta_1^* \otimes \beta_1^* + p_2\beta_2^* \otimes \beta_2^*$ , we have

$$\{\lambda_1^*, \lambda_2^*\} = \{\frac{1+\kappa}{2}, \frac{1-\kappa}{2}\}$$

Using lemma A.3, we have,

$$\|\mathbf{v}_b - \mathbf{v}_b^*\|_2^2 \le \frac{6\widetilde{\delta}}{\kappa} + \frac{24\widetilde{\delta}}{1-\kappa} \le \frac{24\widetilde{\delta}}{\kappa(1-\kappa)}, \ b = 1, 2.$$

Then

$$\|\beta_{b}^{*} - \beta_{b}\|_{2} \leq \left|\sqrt{\frac{1 - \Delta_{b}^{*}}{2}} \mathbf{v}_{b}^{*} - \sqrt{\frac{1 - \Delta_{b}}{2}} \mathbf{v}_{b}\right| + \left|\sqrt{\frac{1 + \Delta_{b}^{*}}{2}} \mathbf{v}_{-b}^{*} - \sqrt{\frac{1 + \Delta_{b}}{2}} \mathbf{v}_{-b}\right|.$$
(14)

Note that

$$\begin{split} \left| \sqrt{\frac{1 - \Delta_b^*}{2}} \mathbf{v}_b^* - \sqrt{\frac{1 - \Delta_b}{2}} \mathbf{v}_b \right| &= \sqrt{\frac{1 - \Delta_b^*}{2}} \mathbf{v}_b^* - \sqrt{\frac{1 - \Delta_b^*}{2}} \mathbf{v}_b + \sqrt{\frac{1 - \Delta_b^*}{2}} \mathbf{v}_b - \sqrt{\frac{1 - \Delta_b}{2}} \mathbf{v}_b \right| \\ &\leq \sqrt{\frac{1 - \Delta_b^*}{2}} \| \mathbf{v}_b - \mathbf{v}_b^* \|_2 + \left| \sqrt{\frac{1 - \Delta_b^*}{2}} - \sqrt{\frac{1 - \Delta_b}{2}} \right| \| \mathbf{v}_b \|_2 \\ &\leq \| \mathbf{v}_b - \mathbf{v}_b^* \|_2 + \left| \sqrt{\frac{1 - \Delta_b^*}{2}} - \sqrt{\frac{1 - \Delta_b}{2}} \right| \\ &\leq \| \mathbf{v}_b - \mathbf{v}_b^* \|_2 + \sqrt{\frac{1 - \Delta_b^*}{2}} - \sqrt{\frac{1 - \Delta_b}{2}} \Big| . \end{split}$$

Plug the above result back to (14), we obtain

$$\begin{split} \|\beta_b^* - \beta_b\|_2 &\lesssim \sqrt{\left|\Delta_b - \Delta_b^*\right|} + \sum_b \|\mathbf{v}_b - \mathbf{v}_b^*\|_2 \\ &\lesssim \sqrt{\frac{\widetilde{\delta}}{\kappa(1-\kappa)}} + \frac{1}{\sqrt{\min\{p_1, p_2\}}} \sqrt{\widetilde{\delta} + \frac{\widetilde{\delta}}{\kappa^2}} \\ &\lesssim \sqrt{\frac{\widetilde{\delta}}{\min\{p_1, p_2\}}} \times \sqrt{\frac{1}{\kappa(1-\kappa)} + \frac{1}{\kappa^2}} \\ &= \sqrt{\frac{\widetilde{\delta}}{\min\{p_1, p_2\}}} \frac{1}{\kappa\sqrt{1-\kappa}}. \end{split}$$

By setting the above upper bound to be less than  $\hat{c} \min\{p_1, p_2\} \|\beta_1^* - \beta_2^*\|_2$ , we complete the proof.

## A.4. Proof of Proposition 4.5

It's equivalent to show that  $J_b = J_b^*, b = 1, 2$ . Let's consider b = 1, that is for all  $p_1 * |S_t|$  samples that are generated by  $y = \mathbf{x}^T \beta_1^*$ . For simplicity, let  $\beta_1, \beta_2$  denote  $\beta_1^{(t-1)}, \beta_2^{(t-1)}$ , we need

$$\left(\mathbf{x}^{T}(\beta_{1}^{*}-\beta_{1})\right)^{2} < \left(\mathbf{x}^{T}(\beta_{1}^{*}-\beta_{2})\right)^{2}$$

From lemma 5.1,

$$\mathbb{P}\left[\left(\mathbf{x}^{T}(\beta_{1}^{*}-\beta_{1})\right)^{2} < \left(\mathbf{x}^{T}(\beta_{1}^{*}-\beta_{2})\right)^{2}\right] \ge 1 - \frac{\|\beta_{1}^{*}-\beta_{1}\|_{2}}{\|\beta_{1}^{*}-\beta_{2}\|_{2}}$$
(15)

$$\geq 1 - 2 \frac{\|\beta_1^* - \beta_1\|_2}{\|\beta_1^* - \beta_2^*\|_2} \tag{16}$$

$$\geq 1 - \frac{2c_1}{k^2}.$$
 (17)

Then we use union bound for  $p_1 * |S_t|$  samples in  $J_1^*$ ,

$$\mathbb{P}\left[\left(\mathbf{x}_{i}^{T}(\beta_{1}^{*}-\beta_{1})\right)^{2} < \left(\mathbf{x}_{i}^{T}(\beta_{1}^{*}-\beta_{2})\right)^{2}, \text{ for all } i \in J_{1}^{*}\right] \geq 1 - p_{1}c_{2}k \times \frac{2c_{1}}{k^{2}} \geq 1 - \frac{c'}{k}.$$

So all samples are correctly clustered with high probability.

As  $\frac{1}{\min(p_1,p_2)}k < |\mathcal{S}_t|$ , number of samples in  $J_1$  and  $J_2$  are both greater than k. Therefore, least square solution reveals the ground truth. In other words,  $err^{(t)} = 0$ .

#### A.5. Proof of Lemma 5.1

(1)

Without loss of generality, we assume  $T\{u, v\} = T\{\mathbf{e}_1, \mathbf{e}_2\}$ . Let  $x_1, x_2$  denote  $\mathbf{x}^T \mathbf{e}_1, \mathbf{x}^T \mathbf{e}_2$ . As  $x_1, x_2$  are independent Gaussian random variables, we have  $x_1 = A \cos \theta, x_2 = A \sin \theta$ , where A is Rayleigh random variable, and  $\theta$  is uniformly distributed over  $[0, 2\pi)$ . Conditioning on  $(\mathbf{x}^T u)^2 > (\mathbf{x}^T v)^2$ , the range of  $\theta$  is truncated to be  $[\theta_0, \theta_0 + \alpha_{(u,v)}] \cup [\theta_0 + \pi, \theta_0 + \pi + \alpha_{(u,v)}]$  for some  $\theta_0$ . It is not hard to see the eigenvalues of covariance matrix of  $(x_1, x_2)$  are  $1 + \frac{\sin \alpha_{(u,v)}}{\alpha_{(u,v)}}, 1 - \frac{\sin \alpha_{(u,v)}}{\alpha_{(u,v)}}$ . As the rest if the eigenvalues of  $\Sigma$  are 1, this completes the proof.

Note that

$$\mathbb{P}\left[(\mathbf{x}^T u)^2 > (\mathbf{x}^T v)^2\right] = \frac{\alpha_{(u,v)}}{\pi}.$$

If  $||u||_2 > ||v||_2$ ,  $\alpha_{(u,v)} > \frac{\pi}{2}$ , when  $||u||_2 < ||v||_2$ ,

$$\cos \alpha_{(u,v)} \ge \frac{\|v\|_2^2 - \|u\|_2^2}{\|u\|_2^2 + \|v\|_2^2}.$$

Note that for any  $\alpha \in [0, \pi/2], \alpha \leq \frac{\pi}{2} \sin \alpha$ . We have

$$\mathbb{P}\left[ (\mathbf{x}^T u)^2 > (\mathbf{x}^T v)^2 \right] \le \frac{1}{2} \sin \alpha_{(u,v)} \le \frac{\|u\|_2 \|v\|_2}{\|u\|_2^2 + \|v\|_2^2} \le \frac{\|u\|_2}{\|v\|_2}.$$

### A.6. Supporting Lemmas

**Lemma A.1.** For any given  $\delta > 0$ , let G denote the grid points, at resolution  $\delta$ , of the unit circle on the subspace spanned by the top two eigenvectors of M, formed with  $|S_*|$  samples. Then, there exists an absolute constant c such that if

$$|S_*| \ge c(1/\tilde{\delta})^2 k \log^2 k,$$

where

$$\tilde{\delta} = \frac{\delta^2}{384} (1 - \sqrt{1 - 4(1 - \langle \beta_1^*, \beta_2^* \rangle^2) p_1 p_2}),$$

then

$$\min_{\mathbf{a}\in G} \|\beta_i^* - \mathbf{a}\| \le \delta, i = 1, 2,$$

with probability at least  $1 - O\left(\frac{1}{k^2}\right)$ .

*Proof.* In order to prove the result, we make use of standard concentration results.

Let  $\Sigma = \mathbb{E}[M]$ . We observe that  $\mathbb{P}[|y| > \sqrt{2\alpha \log k}] \le n^{-\alpha}$ ,  $\mathbb{P}[\|\mathbf{x}\|_2^2 \ge 3k] \le e^{-k/3}$ . Suppose N is much less than  $O(k^{10})$ , where the constant is arbitrarily chosen here. Set  $\alpha = 12$ . Then with probability at least  $1 - O(\frac{1}{k^2})$ , The vectors  $y_i \mathbf{x}_i$  are all supported in a ball with radius  $\sqrt{72k \log k}$ . Directly following theorem 5.44 in (Vershynin, 2010), we claim that when  $N > C(1/\delta)^2 k \log^2 k$ ,

$$\|M - \Sigma\| \le \tilde{\delta} \|\Sigma\| \le 3\tilde{\delta}.$$

We use  $\sigma_i(A)$  to denote the *i*'th biggest eigenvalue of the positive semidefinite matrix A. By simple algebraic calculation we get  $\sigma_1(\Sigma) = 2 + \kappa$ ,  $\sigma_2(\Sigma) = 2 - \kappa$ , where  $\kappa = \sqrt{1 - 4(1 - \langle \beta_1^*, \beta_2^* \rangle^2)p_1p_2}$ . The top two eigenvectors of  $\Sigma$  are denoted as  $\mathbf{v}_1^*, \mathbf{v}_2^*$ . We use  $\mathbf{v}_1, \mathbf{v}_2$  to denote the top two eigenvectors of M. Lemma A.3 yields that

$$\|\mathbf{v}_{i}^{*} - \mathcal{P}_{T(\mathbf{v}_{1},\mathbf{v}_{2})}\mathbf{v}_{i}^{*}\|_{2}^{2} \leq \frac{12\delta}{\sigma_{2}(M) - \sigma_{3}(M)}$$
$$\leq \frac{12\tilde{\delta}}{\sigma_{2}(\Sigma) - \sigma_{3}(\Sigma) - 6\tilde{\delta}}$$
$$= \frac{12\tilde{\delta}}{1 - \kappa - 6\tilde{\delta}}$$
$$= \frac{24\tilde{\delta}}{1 - \kappa}, i = 1, 2.$$

The last inequality holds when  $\tilde{\delta} \leq \frac{1-\kappa}{12}$ . Using the fact that for any two vectors  $\mathbf{a}, \mathbf{b}, \|\mathbf{a} + \mathbf{b}\|_2^2 \leq 2\|\mathbf{a}\|_2^2 + 2\|\mathbf{b}\|_2^2$ , we conclude that

$$\|\beta_i^* - \mathcal{P}_{T(\mathbf{v}_1, \mathbf{v}_2)}\beta_i^*\|_2^2 \le \frac{48\delta}{1-\kappa}, i = 1, 2.$$

Let  $w = \|\beta_i^* - \mathcal{P}_{T(\mathbf{u},\mathbf{v})}\beta_i^*\|_2$ . Then, by simple geometric relation,

$$\begin{split} \min_{\mathbf{a}\in\mathbb{S}^{k-1}\cap T_{(\mathbf{u},\mathbf{v})}} \|\mathbf{a}-\beta_i^*\|_2^2 &\leq 2-2\sqrt{1-w^2}\\ &\leq 2w^2\\ &\leq (\frac{\epsilon}{2})^2, i=1,2. \end{split}$$

Consider the  $\delta$ -resolution grid G. We observe that for any point in  $\mathbb{S}^{k-1} \cap T_{(\mathbf{u},\mathbf{v})}$ , there exists a point in G that is within  $\delta/2$  away from it. By triangle inequality, we end up with

$$\min_{\mathbf{a}\in W} \|\mathbf{a} - \beta_i^*\|_2 \le \delta.$$
(18)

**Lemma A.2.** Let  $\hat{\beta}_1, \hat{\beta}_2$  be any two given vectors with error defined by  $\operatorname{err} := \max_{i=1,2} \|\hat{\beta}_i - \beta_i^*\|$ . There exist constants  $c_1, c_2 > 0$  such that as long as we have enough testing samples,

$$|\mathcal{S}_+| \ge c_1 k / \min\{p_1, p_2\},\$$

then with probability at least  $1 - O(e^{-c_2k})$ 

$$\sqrt{\frac{\mathcal{L}_{+}(\hat{\beta}_{1},\hat{\beta}_{2})}{|\mathcal{S}_{+}|}} \le 1.1 \,\mathrm{err}$$

and

$$\sqrt{\frac{\mathcal{L}_{+}(\hat{\beta}_{1},\hat{\beta}_{2})}{|\mathcal{S}_{+}|}} \geq \frac{1}{5}\sqrt{\min\{p_{1},p_{2}\}}\min\left\{\mathrm{err},\frac{1}{2}\|\beta_{1}^{*}-\beta_{2}^{*}\|_{2}\right\}.$$

*Proof.* Our notation here, namely,  $J_1, J_2, J_1^*, J_2^*$ , is consistent with proof of Theorem 4.4. Note that we have:

$$\mathcal{L}(\beta_1, \beta_2) = \sum_i \min_{z_i} z_i (y_i - \mathbf{x}_i^T \beta_1)^2 + (1 - z_i)(y_i - \mathbf{x}_i^T \beta_2)^2.$$

For the upper bound, we assign label  $z_i$  as the true label. Then,

$$\mathcal{L} \leq \sum_{i \in J_1^*} (\mathbf{x}_i^T (\beta_1^* - \beta_1))^2 + \sum_{i \in J_2^*} (\mathbf{x}_i^T (\beta_2^* - \beta_2))^2.$$

When  $|S_+| \ge C \frac{k}{\min\{p_1, p_2\}}$ , then the number of samples in set  $J_1^*, J_2^*$  is also greater than Ck. Following standard concentration results, there exist constants  $C, c_1$ , such that with probability greater than  $1 - e^{-c_1k}$ , we have

$$\left\|\frac{1}{p_j|\mathcal{S}_+|}\sum_{i\in J_j^*} (\mathbf{x}_i \mathbf{x}_i^T) - I\right\| \le 0.21, j = 1, 2.$$

We have

$$\mathcal{L} \le 1.21 p_1 |\mathcal{S}_+| ||\beta_1 - \beta_1^*||_2^2 + 1.21 p_2 |\mathcal{S}_+| ||\beta_2 - \beta_2^*||_2^2$$
  
$$\le 1.21 |\mathcal{S}_+| \operatorname{err}^2.$$

For the lower bound, we observe that

$$\mathcal{L} = \underbrace{\sum_{i \in J_1 \cap J_1^*} (\mathbf{x}_i^T (\beta_1 - \beta_1^*))^2 + \sum_{i \in J_2 \cap J_1^*} (\mathbf{x}_i^T (\beta_2 - \beta_1^*))^2}_{A_1} + \underbrace{\sum_{i \in J_1 \cap J_2^*} (\mathbf{x}_i^T (\beta_1 - \beta_2^*))^2 + \sum_{i \in J_2 \cap J_2^*} (\mathbf{x}_i^T (\beta_2 - \beta_2^*))^2}_{A_2}.$$

First we consider the first term, A1. Note a simple fact that  $\|\beta_1 - \beta_1^*\|_2 < \|\beta_2 - \beta_1^*\|_2$  or  $\|\beta_1 - \beta_1^*\|_2 > \|\beta_2 - \beta_1^*\|_2$ . In the first case, from Lemma 5.1,  $\mathbb{E}[|J_1 \cap J_1^*|] \ge \frac{1}{2}p_1|\mathcal{S}_+|$ . From Hoeffding's inequality and concentration result (see proof of Lemma 5.1 for similar techniques), for any  $\delta \in (0, 1 - \frac{2}{\pi})$ , there exist constants  $C', c'_1$ , such that when  $N \ge C'k/p_1$ , with probability at least  $1 - e^{-c'_1k}$ ,

$$\sum_{i \in J_1 \cap J_1^*} (\mathbf{x}_i^T (\beta_1 - \beta_1^*))^2 \ge \frac{1}{4} p_1 |\mathcal{S}_+| (1 - \frac{1}{\pi} - \delta) ||\beta_1 - \beta_1^* ||_2^2$$

In the second case, we have a similar result:

$$\sum_{i \in J_2 \cap J_1^*} (\mathbf{x}_i^T (\beta_2 - \beta_1^*))^2 \ge \frac{1}{4} p_1 |\mathcal{S}_+| (1 - \frac{1}{\pi} - \delta) ||\beta_2 - \beta_1^* ||_2^2.$$

Let  $1 - \frac{2}{\pi} - \delta = 0.3$  and choose  $C', c'_1$  to let the above results also hold for A2. We then conclude that when  $N > C' \frac{k}{\min\{p_1, p_2\}}$ ,

$$\mathcal{L} \geq \frac{0.3}{4} p_1 |\mathcal{S}_+| \min\{\|\beta_1 - \beta_1^*\|_2^2, \|\beta_2 - \beta_1^*\|_2^2\} + \frac{0.3}{4} p_2 |\mathcal{S}_+| \min\{\|\beta_1 - \beta_2^*\|_2^2, \|\beta_2 - \beta_2^*\|_2^2\}.$$
(19)

When  $\|\beta_1 - \beta_1^*\|_2 < \|\beta_2 - \beta_1^*\|_2$  and  $\|\beta_2 - \beta_2^*\|_2 < \|\beta_1 - \beta_2^*\|_2$ , (19) implies

$$\mathcal{L} \ge \frac{1}{25} \min\{p_1, p_2\} |\mathcal{S}_+| \text{err}^2.$$
 (20)

When  $\|\beta_1 - \beta_1^*\|_2 > \|\beta_2 - \beta_1^*\|_2$  and  $\|\beta_2 - \beta_2^*\|_2 < \|\beta_1 - \beta_2^*\|_2$ , we have

$$\mathcal{L} \ge \frac{1}{25} \min\{p_1, p_2\} |\mathcal{S}_+| (\|\beta_2 - \beta_1^*\|_2^2 + \|\beta_2 - \beta_2^*\|_2^2)$$
(21)

$$\geq \frac{1}{25} \min\{p_1, p_2\} |\mathcal{S}_+| \frac{1}{4} \|\beta_1^* - \beta_2^* \|_2^2.$$
<sup>(22)</sup>

Note that it is impossible for  $\|\beta_1 - \beta_1^*\|_2 > \|\beta_2 - \beta_1^*\|_2$  and  $\|\beta_2 - \beta_2^*\|_2 > \|\beta_1 - \beta_2^*\|_2$  both to be true. Otherwise, we could switch the subscripts of the two  $\beta$ 's. Putting (20) and (22) together, we complete the proof.

**Lemma A.3.** Suppose symmetric matrix  $\Sigma \in \mathbb{R}^{n \times n}$  has eigenvalues  $\lambda_1 \geq \lambda_2 > \lambda_3$ ... with corresponding normalized eigenvectors denoted as  $u_1, u_2, u_3, \ldots$  Let M be another symmetric matrix with eigenvalues:  $\tilde{\lambda}_1 \geq \tilde{\lambda}_2 > \tilde{\lambda}_3$ ... and eigenvectors  $\tilde{u}_1, \tilde{u}_2, \tilde{u}_3, \ldots$  (a) Let  $span\{u_1, u_2\}$  denote the hyperplane spanned by  $u_1 u_2$ . If  $||M - \Sigma||_2 \leq \varepsilon$ , for  $\varepsilon < \frac{\lambda_2 - \lambda_3}{2}$  we have

$$\|\tilde{u}_{i} - \mathcal{P}_{T(u_{1}, u_{2})}\tilde{u}_{i}\|_{2}^{2} \leq \frac{4\varepsilon}{\lambda_{2} - \lambda_{3}}, i = 1, 2.$$
(23)

*Moreover, if*  $\lambda_1 \neq \lambda_2$ *,* 

$$\|u_1 - \tilde{u}_1\|_2^2 \le \frac{4\epsilon}{\lambda_1 - \lambda_2} \tag{24}$$

$$\|u_2 - \tilde{u}_2\|_2^2 \le \frac{4\epsilon}{\lambda_1 - \lambda_2} + \frac{8\epsilon}{\lambda_2 - \lambda_3}$$

$$\tag{25}$$

*Proof.* Suppose  $\tilde{u}_1 = \alpha_1 u_1 + \beta_1 u_2 + \gamma_1 w$ ,  $\tilde{u}_2 = \alpha_2 u_1 + \beta_2 u_2 + \gamma_2 v$ , where w, v are vector orthogonal to  $span\{u_1, u_2\}$ . We have  $\alpha_1^2 + \beta_1^2 + \gamma_1^2 = \alpha_2^2 + \beta_2^2 + \gamma_2^2 = 1$ . Since  $||M - \Sigma||_2 \le \varepsilon$ ,

$$\tilde{u}_1^T M \tilde{u}_1 \ge \lambda_1 - \varepsilon \tag{26}$$

$$\tilde{u}_1^T M \tilde{u}_1 \leq \tilde{u}_1^T (M - \Sigma) \tilde{u}_1 + \tilde{u}_1^T \Sigma \tilde{u}_1$$
(27)

$$\leq \varepsilon + \tilde{u}_1^T \Sigma \tilde{u}_1. \tag{28}$$

Combining (26) and (27), using  $\tilde{u}_1^T \Sigma \tilde{u}_1 = \alpha_1^2 \lambda_1 + \beta_1^2 \lambda_2 + \gamma_1^2 \lambda_3$ , we get

$$\alpha_1^2 \lambda_1 + \beta_1^2 \lambda_2 + \gamma_1^2 \lambda_3 \ge \lambda_1 - 2\varepsilon \tag{29}$$

Since  $\alpha_1^2 \lambda_1 + \beta_1^2 \lambda_2 + \gamma_1^2 \lambda_3 \leq (1 - \gamma_1^2) \lambda_1 + \gamma_1^2 \lambda_3$ , it implies that

$$\gamma_1^2 \le \frac{2\varepsilon}{\lambda_1 - \lambda_3} \le \frac{2\varepsilon}{\lambda_2 - \lambda_3}.$$
(30)

We assume  $\lambda_1 \neq \lambda_2$ . Otherwise, the above inequality also holds for  $\tilde{u}_2$ , then the proof of (23) is completed. By using another upper bound  $\alpha_1^2 \lambda_1 + \beta_1^2 \lambda_2 + \gamma_1^2 \lambda_3 \leq \alpha_1^2 \lambda_1 + (1 - \alpha_1^2) \lambda_2$ , the following inequality  $\alpha_1^2$  holds

$$\alpha_1^2 \ge 1 - \frac{2\varepsilon}{\lambda_1 - \lambda_2}.\tag{31}$$

Note  $\|\tilde{u}_2 - \mathcal{P}_{T(u_1, u_2)}\tilde{u}_2\|_2^2 = \gamma_1^2$ , we get the distance bound of  $u_1$ . Next, we show the bound for  $\tilde{u}_2$ . Similar to (29),

$$\alpha_2^2 \lambda_1 + \beta_2^2 \lambda_2 + \gamma_2^2 \lambda_3 \ge \lambda_2 - 2\varepsilon.$$
(32)

Again, by using  $\alpha_2^2 \lambda_1 + \beta_2^2 \lambda_2 + \gamma_2^2 \lambda_3 \le \alpha_2^2 \lambda_1 + (1 - \alpha_2^2) \lambda_2$ , we get

$$\gamma_2^2 \le \frac{2\varepsilon + \alpha_2^2(\lambda_1 - \lambda_2)}{\lambda_2 - \lambda_3}.$$
(33)

We use the condition that  $\tilde{u}_1 \tilde{u}_2$  are orthogonal. Hence,  $\alpha_1^2 \alpha_2^2 \le (1 - \alpha_1^2)(1 - \alpha_2^2)$ . It is easy to see  $\alpha_1^2 + \alpha_2^2 \le 1$ . Plugging it into (33) and using (31) result in

$$\gamma_2^2 \le \frac{4\varepsilon}{\lambda_2 - \lambda_3}.\tag{34}$$

Through (30) and (34), we complete the proof of (23).

Using some intermediate results, we derive the bounds for eigenvectors in the case  $\lambda_1 \neq \lambda_2$ .

$$\begin{aligned} \|u_1 - \tilde{u}_1\|_2^2 &= (1 - \alpha_1)^2 + \beta_1^2 + \gamma_1^2 \\ &= (1 - \alpha_1)^2 + 1 - \alpha_1^2 \\ &\leq 2(1 - \alpha_1^2) \\ &\leq \frac{4\epsilon}{\lambda_1 - \lambda_2}. \end{aligned}$$

The last inequality follows from (31).

Similarly,

$$\begin{aligned} \|u_2 - \tilde{u}_2\|_2^2 &\leq 2(1 - \beta_2^2) \\ &= 2(\alpha_2^2 + \gamma_2^2) \\ &\leq 2(1 - \alpha_1^2 + \gamma_2^2) \\ &\leq \frac{4\epsilon}{\lambda_1 - \lambda_2} + \frac{8\epsilon}{\lambda_2 - \lambda_3} \end{aligned}$$

We obtain the last inequality from (31) and (34).