

Supplementary Material of “Rebuilding Factorized Information Criterion: Asymptotically Accurate Marginal Likelihood”

May 16, 2015

1 Proofs

Proof of Proposition 4. If \mathbf{Z} is not degenerated, then Laplace’s method yields Eq. (10). By collecting from Eq. (10) the terms that depend on \mathbf{Z} , we obtain

$$p(\mathbf{Z} \mid \mathbf{X}, K) \propto p(\mathbf{Z}, \mathbf{X} \mid \hat{\boldsymbol{\Pi}}, K) |\mathbf{F}_{\hat{\boldsymbol{\Pi}}}|^{-1/2} (1 + O(N^{-1})). \quad (101)$$

If $p(\mathbf{Z} \mid \mathbf{X}, K)$ is degenerated, we consider the transformation (11). Here, the transformed prior $\tilde{p}(\boldsymbol{\Pi}_{K'} \mid K')$ would differ from the original prior $p(\boldsymbol{\Pi}_{K'} \mid K')$. However, since the mapping $\boldsymbol{\Pi} \rightarrow \tilde{\boldsymbol{\Pi}}_{K'}$ is onto $\mathbf{A1}$ and the prior is strictly positive in the whole space of $\boldsymbol{\Pi} \mathbf{A4}$, $\tilde{p}(\boldsymbol{\Pi} \mid K')$ is also strictly positive, including $\tilde{\boldsymbol{\Pi}}_{K'} = \operatorname{argmax}_{\boldsymbol{\Pi}_{K'}} \ln p(\mathbf{X}, \tilde{\mathbf{Z}}_{K'} \mid \boldsymbol{\Pi}_{K'}, K')$. Consequently, we can again use Laplace’s method for $\ln p(\mathbf{X}, \tilde{\mathbf{Z}}_{K'} \mid \hat{\boldsymbol{\Pi}}_{K'}, K')$, and by collecting the terms that depend on \mathbf{Z} , we obtain

$$p(\mathbf{X} \mid \mathbf{Z}, K) \propto p(\mathbf{X}, \tilde{\mathbf{Z}}_{K'} \mid \hat{\boldsymbol{\Pi}}_{K'}, K') |\mathbf{F}_{\hat{\boldsymbol{\Pi}}_{K'}}|^{-1/2} (1 + O(N^{-1})) \quad (102)$$

$$\propto p_{K'}(\tilde{\mathbf{Z}}_{K'}, K') (1 + O(N^{-1})). \quad (103)$$

This concludes the proof. □

Proof of Theorem 2. First, we prove the case that $p(\mathbf{Z} \mid \mathbf{X}, K)$ is not degenerated. In that case, Laplace’s approximation yields Eq. (10) in probability, and substituting Eq. (10) into (7) gives (8).

If $\kappa(p(\mathbf{Z} \mid \mathbf{X}, K)) = K' < K$, Proposition 4 gives us that $p(\mathbf{Z} \mid \mathbf{X}, K) = p_{K'}(\mathbf{Z}) (1 + O(N^{-1}))$. Since

$$\mathbb{E}_{p(\mathbf{z} \mid \mathbf{x}, K)} [\ln p(\mathbf{X}, \mathbf{Z} \mid K)] = \mathbb{E}_{p_{K'}} [\ln p(\mathbf{X}, \mathbf{Z} \mid K)] + O(1)$$

and

$$\begin{aligned} H(p(\mathbf{Z} \mid \mathbf{X}, K)) &= (1 + O(N^{-1})) H(p_{K'}) + (1 + O(N^{-1})) \ln(1 + O(N^{-1})) \\ &= H(p_{K'}) + O(1), \end{aligned}$$

$\ln p(\mathbf{X} | K)$ is rewritten by

$$\mathbb{E}_{p_{K'}}[\ln p(\mathbf{X}, \mathbf{Z} | K)] + H(p_{K'}) + O(1) \quad (104)$$

$$= \mathbb{E}_{p_{K'}}[\mathcal{L}(\hat{\mathbf{Z}}_{K'}, \hat{\mathbf{\Pi}}_{K'}, K')] + H(p_{K'}) + O(1) \quad (105)$$

Here, since the projection $\mathbf{T}_{K'} : \mathbf{Z} \rightarrow \tilde{\mathbf{Z}}_{K'}$ is continuous and onto (**A1**), we can describe $p_{K'}(\mathbf{Z})$ as the density of $\mathbf{Z}_{K'}$ by using a change of variables, which we denote by $\tilde{p}_{K'}(\mathbf{Z}_{K'})$. Now, we can rewrite the first term as the integral over $\mathbf{Z}_{K'}$, i.e.,

$$\mathbb{E}_{p_{K'}}[\mathcal{L}(\tilde{\mathbf{Z}}_{K'}, \hat{\mathbf{\Pi}}_{K'}, K')] = \int \mathcal{L}(\mathbf{T}_{K'}(\mathbf{Z}), \hat{\mathbf{\Pi}}_{K'}, K') p_{K'}(\mathbf{T}_{K'}(\mathbf{Z})) \mathbf{Z} \quad (106)$$

$$= \int \mathcal{L}(\mathbf{Z}_{K'}, \hat{\mathbf{\Pi}}_{K'}, K') \tilde{p}_{K'}(\mathbf{Z}_{K'}) \mathbf{Z}_{K'}. \quad (107)$$

Similarly, $\text{gFIC}(K')$ is rewritten using Proposition 4 as

$$\text{gFIC}(K') = \mathbb{E}_{p_{K'}}[\mathcal{L}(\mathbf{Z}_{K'}, \hat{\mathbf{\Pi}}_{K'}, K')] + H(p_{K'}) + O(1) \quad (108)$$

Again, the first term is written as

$$\mathbb{E}_{p_{K'}}[\mathcal{L}(\mathbf{Z}_{K'}, \hat{\mathbf{\Pi}}_{K'}, K')] = \int \mathcal{L}(\mathbf{Z}_{K'}, \hat{\mathbf{\Pi}}_{K'}, K') p_{K'}(\mathbf{T}_{K'}(\mathbf{Z})) \mathbf{Z}_{K'} \quad (109)$$

$$= \int \mathcal{L}(\mathbf{Z}_{K'}, \hat{\mathbf{\Pi}}_{K'}, K') \tilde{p}_{K'}(\mathbf{Z}_{K'}) \mathbf{Z}_{K'} \quad (110)$$

Since Eq. (107) and (110) are the same, this concludes Eq. (8). \square

Proof of Proposition 6. Proposition 4 shows that, if \mathbf{Z} is non-degenerated,

$$p(\mathbf{Z} | \mathbf{X}, K) \propto p(\mathbf{X}, \mathbf{Z} | \hat{\mathbf{\Pi}}) |\mathbf{F}_{\hat{\mathbf{\Pi}}}|^{-1/2} \quad (111)$$

$$\propto \prod_n p(\mathbf{x}_n, \mathbf{z}_n | \hat{\mathbf{\Pi}}) |\mathbf{F}_{\hat{\mathbf{\Pi}}}|^{-1/2N} \quad (112)$$

Since $\ln |\mathbf{F}_{\hat{\mathbf{\Pi}}}| = O(1)$, $|\mathbf{F}_{\hat{\mathbf{\Pi}}}|^{-1/2N}$ quickly diminishes to 1 for $N \rightarrow \infty$. \square

Proof of Proposition 7. For technical reasons, we redefine the estimators as follows:

$$\hat{\mathbf{\Pi}} \equiv \operatorname{argmax}_{\mathbf{\Pi}} g_N(\mathbf{\Pi}) = \operatorname{argmax}_{\mathbf{\Pi}} \frac{1}{N} \ln p(\mathbf{X}, \mathbf{Z} | \mathbf{\Pi}), \quad (113)$$

$$\bar{\mathbf{\Pi}} \equiv \operatorname{argmax}_{\mathbf{\Pi}} G_N(\mathbf{\Pi}) = \operatorname{argmax}_{\mathbf{\Pi}} \mathbb{E}_q \left[\frac{1}{N} \ln p(\mathbf{X}, \mathbf{Z} | \mathbf{\Pi}) \right]. \quad (114)$$

According to **A5**, $g_N(\mathbf{\Pi})$ is continuous and concave, and it uniformly converges to $G_N(\mathbf{\Pi})$, i.e.,

$$\sup_{\mathbf{\Pi} \in \mathcal{P}} |g_N(\mathbf{\Pi}) - G_N(\mathbf{\Pi})| \xrightarrow{P} 0. \quad (115)$$

This suffices to show the consistency (for example, see Theorem 5.7 in van der Vaart (1998).) \square

2 The gFAB Algorithm of BPCA

Let \mathcal{Q} be restricted as Gaussian with mean field. Then, the form of q is determined as $q(\mathbf{Z}) = \prod_n N(\mathbf{z}_n | \boldsymbol{\mu}_n, \boldsymbol{\Omega})$. By substituting the BPCA's likelihood into Eq. (17), we obtain

$$\begin{aligned}
& \max_{q \in \mathcal{Q}} \mathbb{E}_q \left[-\frac{\lambda}{2} \|\mathbf{X} - \mathbf{Z}\mathbf{W}^\top\|_{\text{Fro}}^2 - \frac{1}{2} \|\mathbf{Z}\|_{\text{Fro}}^2 - \frac{1}{2} \ln |\mathbf{F}| \right] + \frac{ND}{2} \ln \lambda + H(q) + \text{const.} \\
&= \max_{q \in \mathcal{Q}} \mathbb{E}_q \left[-\frac{\lambda}{2} \|\mathbf{X} - \mathbf{Z}\mathbf{W}^\top\|_{\text{Fro}}^2 - \frac{1}{2} \|\mathbf{Z}\|_{\text{Fro}}^2 - \frac{D}{2} \left(\ln \left| \frac{1}{N} \mathbf{Z}^\top \mathbf{Z} \right| + K \ln \lambda \right) \right] + \frac{ND}{2} \ln \lambda + \frac{N}{2} \ln |\boldsymbol{\Omega}| + \text{const.} \\
&\geq \max_{q \in \mathcal{Q}} \mathbb{E}_q \left[-\frac{\lambda}{2} \|\mathbf{X} - \mathbf{Z}\mathbf{W}^\top\|_{\text{Fro}}^2 - \frac{1}{2} \|\mathbf{Z}\|_{\text{Fro}}^2 \right] - \frac{D}{2} \ln \left| \frac{1}{N} \mathbf{S} \right| + \frac{(N-K)D}{2} \ln \lambda + \frac{N}{2} \ln |\boldsymbol{\Omega}| + \text{const.} \\
&= \max_{q \in \mathcal{Q}} -\frac{\lambda}{2} (\text{tr}(\mathbf{W}^\top \mathbf{W} \mathbf{S} - \mathbf{X} \mathbf{W} \mathbb{E}[\mathbf{Z}^\top])) - \frac{1}{2} \text{tr}(\mathbf{S}) - \frac{D}{2} \ln |\mathbf{S}| + \frac{(N-K)D}{2} \ln \lambda + \frac{N}{2} \ln |\boldsymbol{\Omega}| + \text{const.}
\end{aligned}$$

where $\mathbf{S} \equiv \mathbb{E}[\mathbf{Z}^\top \mathbf{Z}] = N\boldsymbol{\Omega} + \sum_n \boldsymbol{\mu}_n \boldsymbol{\mu}_n^\top$. Note that the equality in the second-to-last line holds at $N \rightarrow \infty$.

Update q By setting the derivatives to zero, we obtain the following update rules:

$$\boldsymbol{\mu}_n^{\text{new}} = \lambda \mathbf{x}_n \mathbf{W} \boldsymbol{\Omega}, \quad (116)$$

$$\boldsymbol{\Omega}^{\text{new}} = (\mathbf{I} + \lambda \mathbf{W}^\top \mathbf{W} + D(\mathbf{S} + \mathbf{I})^{-1})^{-1}. \quad (117)$$

Note that we use \mathbf{S} as an auxiliary variable.

Update Π Update rules of \mathbf{W} and λ are almost the same as those of the EM algorithm, which are:

$$\mathbf{W}^{\text{new}} = \mathbf{X}^\top [\boldsymbol{\mu}_1, \dots, \boldsymbol{\mu}_N] \mathbf{S}^{-1}, \quad (118)$$

$$\lambda^{\text{new}} = \frac{(N-K)D}{\mathbb{E}[\|\mathbf{X} - \mathbf{Z}\mathbf{W}^\top\|_{\text{Fro}}^2]}. \quad (119)$$

References

van der Vaart, A. W. *Asymptotic statistics*. Cambridge University Press, 1998.