

Appendix

Proof of proposition 3.1

Proof. Let $N = \sum_{i=1}^n N_i$. Since, conditioned on Λ , N_1, \dots, N_n are independent Poisson process with mean measure Λ , by the superposition proposition for Poisson processes, N is a Poisson process conditioned on Λ with mean measure $n\Lambda$. Since, $\mathbb{E}[e^{t_1 N(A) + t_2 N(B)} | \Lambda] = \mathbb{E}[e^{t_1 N(A)} | \Lambda(A)] \mathbb{E}[e^{t_2 N(B)} | \Lambda(B)]$, and $\Lambda(A)$ and $\Lambda(B)$ are independent, hence, $N(A)$ and $N(B)$ are also independent and therefore, N is a CRM. Hence, $N(dx) = \int_{\mathbb{R}^+} s \bar{N}(dz, dx)$, for a Poisson process \bar{N} on $S \times \mathbb{R}^+$. Moreover,

$$\begin{aligned} & \mathbb{E}[e^{-tN(A)}] \\ &= \mathbb{E} \left[\mathbb{E}[e^{-tN(A)} | \Lambda] \right] \\ &= \mathbb{E} \left[\exp(-n\Lambda(A)(1 - e^{-t})) \right] \\ &= \exp \left(-\mu(A) \int_{\mathbb{R}^+} (1 - e^{-n(1-e^{-t})z}) \rho(dz) \right) \\ &= \exp \left(-\mu(A) \int_{\mathbb{R}^+} \left(\sum_{k=0}^{\infty} \frac{e^{-nz} (nz)^k}{k!} \right. \right. \\ & \quad \left. \left. - \sum_{k=0}^{\infty} \frac{e^{-nz} e^{-kt} (nz)^k}{k!} \right) \rho(dz) \right) \end{aligned}$$

where, we have used the fact that $1 = \sum_{k=0}^{\infty} \frac{e^{-nz} (nz)^k}{k!}$. Rearranging the terms in the above equation, we get

$$\begin{aligned} & \mathbb{E}[e^{-tN(A)}] \\ &= \exp \left(-\mu(A) \sum_{k=1}^{\infty} (1 - e^{-kt}) \int_{\mathbb{R}^+} \frac{e^{-nz} (nz)^k}{k!} \rho(dz) \right). \end{aligned}$$

Hence, the Poisson intensity measure of N , when viewed as a CRM is given by

$$\bar{\nu}(dk, dx) = \mu(dx) \int_{\mathbb{R}^+} \frac{e^{-nz} (nz)^k}{k!} \rho(dz)$$

when $k \in \{1, 2, 3, \dots\}$, and 0 otherwise. The distinct points of N can be obtained by projecting N on S . Hence, by the mapping proposition for Poisson processes (Kingman, 1992), the distinct points of N form a Poisson process with mean measure $\mu^*(dx) = \bar{\nu}(f^{-1}(dx))$, where f is the projection map on S . Hence $f^{-1}(dx) = (\mathbb{R}^+, dx)$, and

$$\begin{aligned} \mu^*(dx) &= \bar{\nu}(\mathbb{R}^+, dx) \\ &= \mu(dx) \int_{\mathbb{R}^+} \sum_{k=1}^{\infty} \frac{e^{-nz} (nz)^k}{k!} \rho(dz) \\ &= \mu(dx) \int_{\mathbb{R}^+} (1 - e^{-nz}) \rho(dz). \end{aligned}$$

Thus, the result follows. \square

Proof of proposition 3.2

Proof. The proof relies on the simple fact, that conditioned on the number of points to be sampled, the points of a Poisson process are independent (Kingman, 1992). Thus, n point processes can be sampled from a measure Λ , by first sampling the number of points in each point process from a Poisson distribution with mean $\Lambda(S)$, and then sampling the points independently. Let $\Lambda = \sum_{i=1}^n \Delta_i \delta_{X_i}$. Let $(X_{l_1}, \dots, X_{l_k})$ be the features discovered by the n Poisson processes. Let the i^{th} point process N_i consist of m_{i1} occurrences of X_{l_1} , m_{i2} occurrences of X_{l_2} and m_{ik} occurrences of X_{l_k} . Then, the joint distribution of the n point processes conditioned on Λ is given by

$$\begin{aligned} & \mathbb{P}(N_1, \dots, N_n | \Lambda) \\ &= \prod_{i=1}^n \frac{\exp(-T) T^{\sum_{j=1}^k m_{ij}}}{(\sum_{j=1}^k m_{ij})!} \prod_{j=1}^k \left(\frac{\Delta_{l_j}}{T} \right)^{m_{ij}}, \end{aligned}$$

where $T = \Lambda(S) = \sum_{i=1}^{\infty} \Delta_i \delta_{X_i}(S) = \sum_{i=1}^{\infty} \Delta_i$. Readjusting the outermost product in the above equation, we get,

$$\mathbb{P}(N_1, \dots, N_n | \Lambda) = \frac{\exp(-nT)}{\prod_{i=1}^n (\sum_{j=1}^k m_{ij})!} \prod_{j=1}^k \Delta_{l_j}^{\sum_{i=1}^n m_{ij}}.$$

Since, we are not interested in the actual points X_{l_i} 's, but only the number of occurrences of the different points in the point processes, that is, $[m_{ij}]_{(n,k)}$, we can sum over every k -tuple of distinct atoms in the random measure Λ . Hence,

$$\begin{aligned} & P([m_{ij}]_{(n,k)} | \Lambda) \\ &= \frac{\exp(-nT)}{\prod_{i=1}^n (\sum_{j=1}^k m_{ij})!} \sum_{\Delta_{l_1} \neq \Delta_{l_2} \neq \dots \neq \Delta_{l_k}} \prod_{j=1}^k \Delta_{l_j}^{\sum_{i=1}^n m_{ij}}, \end{aligned}$$

where the sum is over all subsets of length k of the set $\{\Delta_1, \Delta_2, \dots\}$. Finally, in order to compute the result, we need to take expectation with respect to the distribution of Λ . Towards that end, we note that only the weights of Λ appear in the above equation. From section 2.2, we know that the weights of a CRM with Poisson intensity measure $\rho(dz)\mu(dx)$ form a Poisson process with mean measure $\mu(S)\rho(dz)$. Hence, it is enough to take the expectation with respect to the Poisson process.

$$P([m_{ij}]_{(n,k)}) \tag{32}$$

$$= \frac{1}{\prod_{i=1}^n (\sum_{j=1}^k m_{ij})!} \mathbb{E} \left[\exp(-nT) \tag{33}$$

$$\sum_{\Delta_{l_1} \neq \Delta_{l_2} \neq \dots \neq \Delta_{l_k}} \prod_{j=1}^k \Delta_{l_j}^{\sum_{i=1}^n m_{ij}} \right], \tag{34}$$

The expectation can further be simplified by applying Proposition 2.1 of (James, 2005).

Proposition 6.1 ((James, 2005)). *Let \mathcal{N} be the space of all σ -finite counting measures on \mathbb{R}^+ , equipped with an appropriate σ -field. Let $f : \mathbb{R}^+ \rightarrow \mathbb{R}^+$ and $g : \mathcal{N} \rightarrow \mathbb{R}^+$ be measurable with respect to their σ -fields. Then, for a Poisson process N with mean measure $\mathbb{E}[N(dx)] = \rho(dx)$,*

$$\mathbb{E} \left[g(N) e^{-\sum_{\Delta \in N} f(\Delta)} \right] = \mathbb{E} \left[e^{-\sum_{\Delta \in N} f(\Delta)} \right] \mathbb{E}[g(\bar{N})]$$

where \bar{N} is a Poisson process with mean measure $\mathbb{E}[N(dx)] = e^{-f(x)} \rho(dx)$.

Applying the above proposition to (??), we get

$$P([m_{ij}]_{(n,k)}) = \frac{\mathbb{E} \left[e^{-\sum_{i=1}^n n \Delta_i} \right]}{\prod_{i=1}^n (\sum_{j=1}^k m_{ij})!} \\ \times \mathbb{E} \left[\sum_{\Delta_{i_1} \neq \Delta_{i_2} \neq \dots \neq \Delta_{i_k} \in \bar{N}} \prod_{j=1}^k \Delta_{l_j}^{\sum_{i=1}^n m_{ij}} \right]$$

where \bar{N} is a Poisson process with mean measure $\mathbb{E}[N(dz)] = e^{-nz} \rho(dz)\theta$. The first expectation can be evaluated using Campbell's proposition and is given by $\exp(-\theta \int_{\mathbb{R}^+} (1 - e^{-nz}) \rho(dz))$. In order to evaluate the second expectation, we construct a new point process from N^* on \mathbb{R}^{+k} by concatenating every set of k distinct points in \bar{N} . The expression in the second expectation can then be rewritten as

$$\sum_{(\Delta_{i_1}, \dots, \Delta_{i_k}) \in N^*} \prod_{j=1}^k \Delta_{l_j}^{\sum_{i=1}^n m_{ij}}$$

By Campbell's proposition for point processes,

$$\mathbb{E} \left[\sum_{\Delta \in N} f(\Delta) \right] = \int_{z \in \mathbb{R}^+} f(z) \rho(dz),$$

where $\rho(dz) = \mathbb{E}[N(dz)]$. Moreover, since the point process N^* is obtained by concatenating distinct points in N , $\mathbb{E}[N^*(dz_1, \dots, dz_k)] = \prod_{j=1}^k \mathbb{E}[\bar{N}(dz_j)] = \prod_{j=1}^k \theta e^{-nz} \rho(dz_j)$, whenever z_j 's are distinct. Hence,

$$\mathbb{E} \left[\sum_{\Delta_{i_1} \neq \Delta_{i_2} \neq \dots \neq \Delta_{i_k} \in \bar{N}} \prod_{j=1}^k \Delta_{l_j}^{\sum_{i=1}^n m_{ij}} \right] \\ = \prod_{j=1}^k \int_{z \in \mathbb{R}^+} \theta e^{-nz} z^{\sum_{i=1}^n m_{ij}} \rho(dz).$$

Hence, the final expression for the marginal distribution of the set of counts for each latent feature is given by

$$P([m_{ij}]_{(n,k)}) = \frac{\exp(-\theta \int_{\mathbb{R}^+} (1 - e^{-nz}) \rho(dz))}{\prod_{i=1}^n (\sum_{j=1}^k m_{ij})!} \\ \times \prod_{j=1}^k \int_{z \in \mathbb{R}^+} \theta e^{-nz} z^{\sum_{i=1}^n m_{ij}} \rho(dz)$$

The above expression can be simplified by letting $\psi(t) = \theta \int_{\mathbb{R}^+} (1 - e^{-tz}) \rho(dz)$. Hence, $\psi^{(l)}(t) = (-1)^{l-1} \int_{\mathbb{R}^+} \theta e^{-tz} z^l \rho(dz)$. Hence, the above expression can be rewritten as

$$P([m_{ij}]_{(n,k)}) = (-1)^{\sum_{i=1}^n \sum_{j=1}^k m_{ij} - k} \frac{\theta^k e^{-\theta \psi(n)}}{\prod_{i=1}^n (\sum_{j=1}^k m_{ij})!} \\ \times \prod_{j=1}^k \psi^{(\sum_{i=1}^n m_{ij})}(n)$$

□

Proof of Corollary 3.3

Proof. From proposition 3.1, the distinct points in the point processes $N_i, 1 \leq i \leq n$, form a Poisson process with mean measure $\frac{\mu(dx)}{\mu(S)} \psi(n)$. Hence, the total number of distinct points k is distributed as Poisson($\psi(n)$). Hence, conditioning equation (5) with respect to k , we get the desired result. □

Proof of Proposition 3.4

Proof. Let $N = \sum_{i=1}^n N_i$. From the arguments of proposition 3.1, N is a CRM, and hence, can be written as $N(dx) = \int_{\mathbb{R}^+} z \bar{N}(dz, dx)$ for some Poisson process \bar{N} . Let Π be the random collection of points corresponding to \bar{N} . Now define a map $f : \mathbb{R}^+ \times S \rightarrow S$ as the projection map on S , that is, $f(x, y) = y$ and $M = f(\Pi) = \{\{f(x, y) : (x, y) \in \Pi\}\}$, where the double brackets indicate that M is a multiset. The rest of the arguments remain same as in proposition 3.1 and proposition 3.2. □

Proof of Lemma 4.1

Proof. Using Proposition 3.4 to marginalize Λ_i from 8, we get that $[m_{ij}]_{1 \leq j \leq r_i}$ is distributed as CRM-Poisson($\Phi(S), \rho, 1$), that is,

$$P([m_{ij}]_{1 \leq j \leq r_i} | \Phi(S)) \\ = \frac{\exp(-\Phi(S) \int_{\mathbb{R}^+} (1 - e^{-z}) \bar{\rho}(dz))}{(\sum_{j=1}^{r_i} m_{ij})!} \\ \times \prod_{j=1}^{r_i} \int_{z \in \mathbb{R}^+} \Phi(S) e^{-z} z^{m_{ij}} \bar{\rho}(dz) \quad (35)$$

Let $m_{i\cdot} = \sum_{j=1}^{r_{i\cdot}} m_{ij}$. Taking expectation with respect to $\Phi(S)$, we get the marginal distribution of $[m_{ij}]_{1 \leq j \leq r_{i\cdot}}$, where $r_{i\cdot}$ is also random.

$$\begin{aligned} & P([m_{ij}]_{1 \leq j \leq r_{i\cdot}}) \\ &= \mathbb{E} \left[\exp \left(-\Phi(S) \int_{\mathbb{R}^+} (1 - e^{-z}) \bar{\rho}(dz) \right) \Phi(S)^{r_{i\cdot}} \right] \\ & \quad \times \frac{\prod_{j=1}^{r_{i\cdot}} \int_{z \in \mathbb{R}^+} e^{-z} z^{m_{ij}} \bar{\rho}(dz)}{m_{i\cdot}!} \end{aligned} \quad (36)$$

It is given that

$$\begin{aligned} h(u) &= \mathbb{E}[e^{-u\Phi(S)}] \\ \bar{\psi}(u) &= \int_{\mathbb{R}^+} (1 - e^{-uz}) \bar{\rho}(dz), \end{aligned}$$

Hence

$$\begin{aligned} \frac{d^{r_{i\cdot}}}{du^{r_{i\cdot}}} h(u) &= (-1)^{r_{i\cdot}} \mathbb{E} \left[\Phi(S)^{r_{i\cdot}} e^{-u\Phi(S)} \right] \\ \bar{\psi}^{(m_{ij})}(u) &= (-1)^{m_{ij}-1} \int_{\mathbb{R}^+} e^{-uz} z^{m_{ij}} \bar{\rho}(dz) \end{aligned}$$

Using the above results with $u = \bar{\psi}(1)$, equation (??) can be rewritten as

$$\begin{aligned} & P([m_{ij}]_{1 \leq j \leq r_{i\cdot}}) \\ &= (-1)^{m_{i\cdot}} h^{(r_{i\cdot})}(\bar{\psi}(1)) \frac{\prod_{j=1}^{r_{i\cdot}} \bar{\psi}^{(m_{ij})}(1)}{m_{i\cdot}!} \end{aligned} \quad (37)$$

□