

Appendix

Proof of Lemma 2.1. Since unitary transformations preserve dot-products, i.e., $\langle T(x), T(y) \rangle = \langle x, y \rangle$, we need to show that a group element acting on the image $I : \mathbb{R}^2 \mapsto \mathbb{R}$ as $T_g[I(x)] = |J_g|^{-1/2} I(T_g^{-1}(x))$, $\forall x$ is a unitary transformation.

Let J_g be the Jacobian of the transformation T_g , with determinant $|J_g|$. We have

$$\begin{aligned} \|I(T_g^{-1}(\cdot))\|^2 &= \int I^2(T_g^{-1}(x)) dx \\ &= \int I^2(z) |J_g| dz, \quad \text{substituting } z = T_g^{-1}(x) \Rightarrow dx = |J_g| dz \\ &= |J_g| \|I(\cdot)\|^2 \end{aligned}$$

Hence the transformation given as $T_g[I(\cdot)] = |J_g|^{-1/2} I(T_g^{-1}(\cdot))$ is unitary and thus $\langle T_g(I), T_g(I') \rangle = \langle I, I' \rangle$ for two images I and I' . \square

Proof of Theorem 3.1. We first define the notion of *U-statistics* [21].

U-statistics - Let $g : \mathbb{R}^2 \rightarrow \mathbb{R}$ be a symmetric function of its arguments. Given an i.i.d. sequence $X_1, X_2 \dots X_k$ of $k (\geq 2)$ random variables, the quantity $U := \frac{1}{n(n-1)} \sum_{i \neq j, i, j=1}^n g(X_i, X_j)$ is known as a pairwise **U-statistics**. If $\theta(P) = \mathbb{E}_{X_1, X_2 \sim P} g(X_1, X_2)$ then U is an unbiased estimate of $\theta(P)$.

Our goal is to bound

$$\sup_{x, y \in X} \left| \langle \psi_{RF}(x), \psi_{RF}(y) \rangle - k_{q,G}(x, y) \right|$$

where

$$\psi_{RF}(x) = \frac{1}{r} \sum_{i=1}^r z(g_i x), x \in X \subset \mathbb{R}^d.$$

We work with $z(\cdot) = \sqrt{2/s} [\cos(\langle \omega_1, \cdot \rangle + b_1), \dots, \cos(\langle \omega_s, \cdot \rangle + b_s)] \in \mathbb{R}^s$ with $b_i \sim \text{Unif}(0, 2\pi)$ as in [33].

Let $\widehat{k}_{q,G}(x, y) := \frac{1}{r^2} \sum_{i, j=1}^{r^2} k(g_i x, g_j y)$ and $\widetilde{k}_{q,G}(x, y) := \frac{1}{r(r-1)} \sum_{i \neq j, i, j=1}^{r^2} k(g_i x, g_j y)$.

Using the triangle inequality we have

$$\begin{aligned} \sup_{x, y \in X} \left| \langle \psi_{RF}(x), \psi_{RF}(y) \rangle - k_{q,G}(x, y) \right| &\leq \underbrace{\sup_{x, y \in X} \left| \langle \psi_{RF}(x), \psi_{RF}(y) \rangle - \widehat{k}_{q,G}(x, y) \right|}_A \\ &\quad + \underbrace{\sup_{x, y \in X} \left| \widehat{k}_{q,G}(x, y) - k_{q,G}(x, y) \right|}_B + \underbrace{\sup_{x, y \in X} \left| \widehat{k}_{q,G}(x, y) - \widetilde{k}_{q,G}(x, y) \right|}_C \end{aligned}$$

Bounding A.

$$A := \sup_{x, y \in X} \left| \frac{1}{r^2} \sum_{i, j} (\langle z(g_i x), z(g_j y) \rangle - k(g_i x, g_j y)) \right|$$

Let us define $f_{ij}(x, y) := \langle z(g_i x), z(g_j y) \rangle - k(g_i x, g_j y)$, and $f(x, y) = 1/r^2 \sum_{i, j} f_{ij}(x, y)$. Since each of the s independent random variables in the summand of $1/r^2 \sum_{i, j} \langle z(g_i x), z(g_j y) \rangle = \frac{1}{s} \sum_{k=1}^s \left(\frac{1}{r^2} \right.$

$\sum_{i,j} 2 \cos(\langle \omega_k, g_i x \rangle + b_k) \cos(\langle \omega_k, g_j y \rangle + b_k)$ is bounded by $[-2, 2]$, using Hoeffding's inequality for a given pair x, y , we have

$$\Pr[|f(x, y)| \geq \varepsilon/4] \leq 2 \exp(-s\varepsilon^2/128).$$

To obtain a uniform convergence guarantee over X , we follow the arguments in [33], relying on covering the space with an ε -net and Lipschitz continuity of the function $f(x, y)$.

Since X is compact, we can find an ε -net that covers X with $N_X = \left(\frac{2 \text{diam}(X)}{\eta}\right)^d$ balls of radius η [12]. Let $\{c_k\}_{k=1}^{N_X}$ be the centers of these balls, and let L_f denote the Lipschitz constant of $f(\cdot, \cdot)$, i.e., $|f(x, y) - f(c_k, c_l)| \leq L_f \|(x, y) - (c_k, c_l)\|$ for all $x, y, c_k, c_l \in X$. For any $x, y \in X$, there exists a pair of centers c_k, c_l such that $\|(x, y) - (c_k, c_l)\| < \sqrt{2}\eta$. We will have $|f(x, y)| < \varepsilon/2$ for all x, y if (i) $|f(c_k, c_l)| < \frac{\varepsilon}{4}$, $\forall c_k, c_l$, and (ii) $L_f < \frac{\varepsilon}{4\sqrt{2}\eta}$.

We immediately get the following by applying union bound for all the center pairs (c_k, c_l)

$$\Pr[\cup_{k,l} |f(c_k, c_l)| \geq \varepsilon/4] \leq 2N_X^2 \exp(-s\varepsilon^2/128). \quad (9)$$

We use Markov inequality to bound the Lipschitz constant of f . By definition, we have $L_f = \sup_{x,y} \|\nabla_{x,y} f(x, y)\| = \|\nabla_{x,y} f(x^*, y^*)\|$, where $\nabla_{x,y} f(x, y) = \begin{pmatrix} \nabla_x f(x, y) \\ \nabla_y f(x, y) \end{pmatrix}$. We also have $\mathbb{E}_{\omega \sim p} \nabla_{x,y} \langle z(g_i x), z(g_j y) \rangle = \nabla_{x,y} k(g_i x, g_j y)$. It follows that

$$\begin{aligned} \mathbb{E}_{\omega \sim p} \|\nabla_{x,y} f(x^*, y^*)\|^2 &= \mathbb{E}_{\omega \sim p} \left\| \frac{1}{r^2} \sum_{i,j=1}^r \nabla_{x,y} \langle z(g_i x^*), z(g_j y^*) \rangle \right\|^2 - \left\| \frac{1}{r^2} \sum_{i,j=1}^r \nabla_{x,y} k(g_i x^*, g_j y^*) \right\|^2 \\ &\leq \mathbb{E}_{\omega \sim p} \left\| \frac{1}{r^2} \sum_{i,j=1}^r \nabla_{x,y} \langle z(g_i x^*), z(g_j y^*) \rangle \right\|^2 \\ &\leq \mathbb{E}_{\omega \sim p} \left(\frac{1}{r^2} \sum_{i,j=1}^r \|\nabla_{x,y} \langle z(g_i x^*), z(g_j y^*) \rangle\| \right)^2 \\ &\leq 2 \mathbb{E}_{\omega \sim p} \sup_{x,y,g_i,g_j} \|\nabla_x \langle z(g_i x), z(g_j y) \rangle\|^2 \\ &\leq 2 \mathbb{E}_{\omega \sim p} \sup_{x,g} \left(\frac{1}{s} \sum_{k=1}^s \|\nabla_x T_g(x) \omega_k\| \right)^2 \\ &\leq 2 \mathbb{E}_{\omega \sim p} \sup_{x,g} \left(\frac{1}{s} \sum_{k=1}^s \|\nabla_x T_g(x)\|_2 \|\omega_k\| \right)^2 \\ &= 2 \mathbb{E}_{\omega \sim p} \sup_{x,g} \|\nabla_x T_g(x)\|_2^2 \frac{1}{s^2} \sum_{k=1}^s \sum_{l=1}^s \|\omega_k\| \|\omega_l\| \\ &= 2 \sup_{x,g} \|\nabla_x T_g(x)\|_2^2 \frac{1}{s^2} \sum_{k=1}^s \sum_{l=1}^s \mathbb{E}_{\omega \sim p} \|\omega_k\| \|\omega_l\| \\ &= 2 \sup_{x,g} \|\nabla_x T_g(x)\|_2^2 \frac{1}{s^2} \left(s \mathbb{E}_{\omega \sim p} \|\omega\|^2 + \sum_{k,l=1, k \neq l}^s (\mathbb{E}_{\omega \sim p} \|\omega\|)^2 \right) \quad (\omega_k \text{ i.i.d.}) \\ &\leq 2 \sup_{x,g} \|\nabla_x T_g(x)\|_2^2 \frac{1}{s^2} \left(s \mathbb{E}_{\omega \sim p} \|\omega\|^2 + \sum_{k,l=1, k \neq l}^s \mathbb{E}_{\omega \sim p} \|\omega\|^2 \right) \quad (\text{Jensen's inequality}) \end{aligned}$$

$$\leq 2\sigma_p^2 \sup_{x \in X, g \in G} \|\nabla_x T_g(x)\|_2^2,$$

where $\sigma_p^2 = \mathbb{E}(\omega^\top \omega)$, and $T_g(x) = gx$ denotes the transformation corresponding to the group action. If we assume the group action to be linear, i.e., $T_g(x + y) = T_g(x) + T_g(y)$ and $T_g(\alpha x) = \alpha T_g(x)$, which holds for all group transformations considered in this work (e.g., rotation, translation, scaling or general affine transformations on image x ; permutations of x), we can bound $\|\nabla_x T_g(x)\|_2$ as

$$\begin{aligned} \|\nabla_x T_g(x)\|_2 &= \sup_{u: \|u\|=1} \|\nabla_x T_g(x)u\| \\ &= \sup_{u: \|u\|=1} \left\| \lim_{h \rightarrow 0} \frac{T_g(x + hu) - T_g(x)}{h} \right\| \quad (\text{directional derivative of vector valued function } T_g(\cdot)) \\ &= \sup_{u: \|u\|=1} \|T_g(u)\| = 1 \\ &\quad (\text{since } T_g(\cdot) \text{ is either unitary or is converted to unitary by construction (see Lemma 2.1)}) \end{aligned}$$

Using Markov inequality, $\Pr[L_f^2 \geq \varepsilon] \leq \mathbb{E}(L_f^2)/\varepsilon$, hence we get

$$\Pr \left[L_f \geq \frac{\varepsilon}{4\sqrt{2}\eta} \right] \leq \frac{64\sigma_p^2\eta^2}{\varepsilon^2}.$$

Combining Eq. (9) with the above result on Lipschitz continuity, we get

$$\Pr \left[\sup_{x,y} |f(x,y)| \leq \varepsilon/2 \right] \geq 1 - 2N_X^2 \exp(-s\varepsilon^2/128) - \frac{64\sigma_p^2\eta^2}{\varepsilon^2}. \quad (10)$$

Bounding B.

As defined earlier, $\tilde{k}_{q,G}(x,y) := \frac{1}{r(r-1)} \sum_{i \neq j, i,j=1}^r k(g_i x, g_j y)$. From the result of U-statistics literature [21], it is easy to see that $\mathbb{E}(\tilde{k}_{q,G}(x,y)) = k_{q,G}(x,y)$.

Since $g_1, g_2 \dots g_r$ are i.i.d samples, we can consider $\tilde{k}_{q,G}(x,y)$ as function of r random variables $(g_1, g_2, \dots g_r)$. Denote $\tilde{k}_{q,G}(x,y)$ as $f(g_1, g_2, \dots g_r)$. Now if a variable g_p is changed to g'_p then we can bound the absolute difference of the changed and the original function. For the rbf kernel, $|k(g_p x, g_j y) - k(g'_p x, g_j y)| \leq 1$

$$\begin{aligned} |f(g_1, g_2, \dots g_p, \dots g_r) - f(g_1, \dots g_{p-1}, g'_p, g_{p+1} \dots g_r)| &= \frac{1}{r(r-1)} \left| \sum_{j=1, j \neq p}^r k(g_p x, g_j y) - k(g'_p x, g_j y) \right| \\ &\leq \frac{1}{r(r-1)} \sum_{j=1, j \neq p}^r |k(g_p x, g_j y) - k(g'_p x, g_j y)| \\ &\leq \frac{(r-1)}{r(r-1)} = \frac{1}{r} \end{aligned}$$

Using bounded difference inequality

$$\Pr \left[\left| f(g_1, g_2, \dots g_r) - \mathbb{E}[f(g_1, g_2 \dots g_r)] \right| \geq \frac{\varepsilon}{2} \right] \leq 2 \exp \left(\frac{-r\varepsilon^2}{2} \right).$$

The above bound holds for a given pair x, y . Similar to the earlier segment for bounding the first term A , we use the ε -net covering of X and Lipschitz continuity arguments to get a uniform convergence guarantee. Using a union bound on all pairs of centers, we have

$$\Pr \left[\bigcup_{k,\ell=1}^{N_X} \left| \mathbb{E}[k(gc_k, g'c_\ell)] - \frac{1}{r(r-1)} \sum_{i,j=1, i \neq j}^r k(g_i c_k, g_j c_\ell) \right| > \frac{\varepsilon}{2} \right] \leq 2N_X^2 \exp \left(\frac{-r\varepsilon^2}{2} \right). \quad (11)$$

In order to extend the bound from the centers c_i to all $x \in X$, we use the Lipschitz continuity argument. Let

$$h(x, y) = k_{q,G}(x, y) - \tilde{k}_{q,G}(x, y).$$

Let L_h denote the Lipschitz constant of $h(\cdot, \cdot)$, i.e., $|h(x, y) - h(c_k, c_l)| \leq L_h \|(x, y) - (c_k, c_l)\|$ for all $x, y, c_k, c_l \in X$. By the definition of ε -net, for any $x, y \in X$, there exists a pair of centers c_k, c_l such that $\|(x, y) - (c_k, c_l)\| < \sqrt{2}\eta$. We will have $|h(x, y)| < \varepsilon/2$ for all x, y if (i) $|h(c_k, c_l)| < \frac{\varepsilon}{4}$, $\forall c_k, c_l$, and (ii) $L_h < \frac{\varepsilon}{4\sqrt{2}\eta}$.

We will again use Markov inequality to bound the Lipschitz constant of h . By definition, we have $L_h = \sup_{x,y} \|\nabla_{x,y} h(x, y)\| = \|\nabla_{x,y} h(x^*, y^*)\|$, where $\nabla_{x,y} h(x, y) = \begin{pmatrix} \nabla_x h(x, y) \\ \nabla_y h(x, y) \end{pmatrix}$. We also have $\mathbb{E}_{\omega \sim p} \nabla_{x,y} \tilde{k}_{q,G}(x, y) = \nabla_{x,y} k_{q,G}(x, y)$. It follows that

$$\begin{aligned} \mathbb{E}_{g_1, \dots, g_r} \|\nabla_{x,y} h(x^*, y^*)\|^2 &= \mathbb{E}_{g_1, \dots, g_r} \|\nabla_{x,y} \tilde{k}_{q,G}(x^*, y^*)\|^2 - \|\nabla_{x,y} k_{q,G}(x^*, y^*)\|^2 \\ &\leq \mathbb{E}_{g_1, \dots, g_r} \|\nabla_{x,y} \tilde{k}_{q,G}(x^*, y^*)\|^2 \\ &= \mathbb{E}_{g_1, \dots, g_r} \left\| \frac{1}{r(r-1)} \sum_{i \neq j} \nabla_{x,y} k(g_i x^*, g_j y^*) \right\|^2. \end{aligned}$$

Noting $T_{g_i}(x) = g_i x$, and $k(x, y) = \exp -\frac{1}{2\sigma^2} \|x - y\|^2$, we have

$$\begin{aligned} \nabla_x k(g_i x, g_j y) &= \nabla_x k(T_{g_i}(x), T_{g_j}(y)) \\ &= -\frac{1}{\sigma^2} \nabla_x T_{g_i}(x) (g_i x - g_j y) \exp \left(-\frac{1}{2\sigma^2} \|g_i x - g_j y\|^2 \right). \end{aligned}$$

Continuing

$$\begin{aligned} \left\| \frac{1}{r(r-1)} \sum_{i \neq j} \nabla_{x,y} k(g_i x, g_j y) \right\| &\leq \frac{1}{r(r-1)} \sum_{i \neq j} \left\| \nabla_{x,y} k(g_i x, g_j y) \right\| \\ &\leq \frac{\sqrt{2}}{r(r-1)} \sup_x \sum_{i \neq j} \left\| \nabla_x k(g_i x, g_j y) \right\| \quad (\text{using symmetry of } k(\cdot, \cdot)) \\ &= \frac{\sqrt{2}}{r(r-1)\sigma^2} \sup_x \sum_{i \neq j} k(g_i x, g_j y) \left\| \nabla_x T_{g_i}(x) (g_i x - g_j y) \right\| \\ &\leq \frac{\sqrt{2}}{r(r-1)\sigma^2} \sum_{i \neq j} k(g_i x, g_j y) \|\nabla_x T_{g_i}(x)\|_2 \|g_i x - g_j y\| \\ &\leq \frac{\sqrt{2}e^{-1/2}}{\sigma} \sup_{x \in X, g \in G} \|\nabla_x T_g(x)\|_2 \quad (\text{using } \sup_{z \geq 0} z e^{-z^2/(2\sigma^2)} = \sigma e^{-1/2}) \\ &\leq \frac{\sqrt{2}e^{-1/2}}{\sigma} \quad (\text{using linearity and unitarity of } T_g(\cdot) \text{ as before}) \end{aligned}$$

It follows that

$$\mathbb{E}(L_h^2) \leq \frac{2}{\sigma^2 e}.$$

Now using Markov inequality we have

$$\mathbb{P} \left[L_h > \sqrt{t} \right] \leq \frac{\mathbb{E}(L_h^2)}{t},$$

Hence we have for $t = \left(\frac{\varepsilon}{4\sqrt{2}\eta}\right)^2$,

$$\mathbb{P}\left[L_h > \frac{\varepsilon}{4\sqrt{2}\eta}\right] \leq \frac{32\eta^2\mathbb{E}((L_h)^2)}{\varepsilon^2} \leq \frac{64\eta^2}{e\sigma^2\varepsilon^2},$$

Hence

$$\Pr[B \leq \varepsilon/2] \geq 1 - 2(N_X)^2 \exp\left(\frac{-r\varepsilon^2}{2}\right) - \frac{64\eta^2}{e\sigma^2\varepsilon^2}.$$

Bounding C.

$$\begin{aligned} \left|\tilde{k}_{q,G}(x,y) - \widehat{k}_{q,G}(x,y)\right| &= \left|\frac{1}{r(r-1)} \sum_{i,j=1,i \neq j}^r k(g_i x, g_j y) - \frac{1}{r^2} \sum_{i,j=1}^r k(g_i x, g_j y)\right| \\ &= \left|\left(\frac{1}{r(r-1)} - \frac{1}{r^2}\right) \sum_{i,j=1,i \neq j}^r k(g_i x, g_j y) - \frac{1}{r^2} \sum_{i,j=1,i=j}^r k(g_i x, g_j y)\right| \\ &\leq \max\left(\frac{1}{r^2(r-1)} \sum_{i,j=1,i \neq j}^r k(g_i x, g_j y), \frac{1}{r^2} \sum_{i,j=1,i=j}^r k(g_i x, g_j y)\right) \quad (\text{since } k(\cdot, \cdot) \geq 0) \\ &\leq \frac{1}{r} \quad (\text{as Gaussian kernel } k(\cdot, \cdot) \leq 1) \end{aligned}$$

Finally we have

$$\sup_{x,y \in X} \left| \langle \psi_{RF}(x), \psi_{RF}(y) \rangle - k_{q,G}(x,y) \right| \leq A + B + C \leq \varepsilon + \frac{1}{r}$$

with a probability at least $1 - 2N_X^2 \exp\left(\frac{-s\varepsilon^2}{128}\right) - 2N_X^2 \exp\left(\frac{-r\varepsilon^2}{2}\right) - \left(\frac{64\eta^2 d}{\varepsilon^2 \sigma^2}\right) - \left(\frac{64\eta^2}{e\varepsilon^2 \sigma^2}\right)$, noting that $\sigma_p^2 = d/\sigma^2$ for the Gaussian kernel $k(x,y) = e^{-\frac{\|x-y\|^2}{2\sigma^2}}$.

Let

$$\begin{aligned} p &= 1 - 2N_X^2 \exp\left(\frac{-s\varepsilon^2}{128}\right) - 2N_X^2 \exp\left(\frac{-r\varepsilon^2}{2}\right) - \left(\frac{64\eta^2 d}{\varepsilon^2 \sigma^2}\right) - \left(\frac{64\eta^2}{e\varepsilon^2 \sigma^2}\right) \\ &= 1 - 2\left(\frac{2\text{diam}(X)}{\eta}\right)^{2d} \exp\left(\frac{-s\varepsilon^2}{128}\right) - 2\left(\frac{2\text{diam}(X)}{\eta}\right)^{2d} \exp\left(\frac{-r\varepsilon^2}{2}\right) - \left(\frac{64\eta^2 d}{\varepsilon^2 \sigma^2}\right) - \left(\frac{64\eta^2}{e\varepsilon^2 \sigma^2}\right) \\ &\geq 1 - 2\eta^{-2d} \left((2\text{diam}(X))^{2d} \exp\left(\frac{-r\varepsilon^2}{2}\right) + (2\text{diam}(X))^{2d} \exp\left(\frac{-s\varepsilon^2}{128}\right) \right) - \eta^2 \left(\frac{64(d+1)}{\varepsilon^2 \sigma^2}\right). \end{aligned}$$

The above probability is of the form of $1 - (\kappa_1 + \kappa_2)\eta^{-2d} - \kappa_3\eta^2$ where $\kappa_1 = 2(2\text{diam}(X))^{2d} \exp\left(\frac{-r\varepsilon^2}{2}\right)$, $\kappa_2 = 2(2\text{diam}(X))^{2d} \exp\left(\frac{-s\varepsilon^2}{128}\right)$ and $\kappa_3 = \left(\frac{64(d+1)}{\varepsilon^2 \sigma^2}\right)$. Choose $\eta = \left(\frac{\kappa_1 + \kappa_2}{\kappa_3}\right)^{\frac{1}{2(d+1)}}$

Hence $p \geq 1 - 2(\kappa_1 + \kappa_2)^{\frac{1}{d+1}} \kappa_3^{\frac{d}{d+1}}$.

For given $\delta_1, \delta_2 \in (0, 1)$, we conclude that for fixed constants C_1, C_2 , for

$$r \geq \frac{C_1 d}{\varepsilon^2} \log(\text{diam}(X)/\delta_1),$$

$$s \geq \frac{C_2 d}{\varepsilon^2} (\log(\text{diam}(X)/\delta_2)),$$

we have

$$\sup_{x,y \in X} \left| \langle \psi_{RF}(x), \psi_{RF}(y) \rangle - k_{q,G}(x,y) \right| \leq \varepsilon + \frac{1}{r},$$

with probability $1 - \left(\frac{64(d+1)}{\varepsilon^2 \sigma^2}\right)^{\frac{d}{d+1}} (\delta_1 + \delta_2)^{\frac{2d}{d+1}}$.

Proof of Theorem 3.2. We give here the proof of Theorem 3.2.

Lemma A.1 (Lemma 4 [35]). - Let $\mathbf{X} = \{x_1, x_2 \dots x_K\}$ be iid random variables in a ball \mathcal{H} of radius M centered around the origin in a Hilbert space. Denote their average by $\bar{\mathbf{X}} = \frac{1}{K} \sum_{i=1}^K x_i$. Then for any $\delta > 0$, with probability at least $1 - \delta$,

$$\|\bar{\mathbf{X}} - \mathbb{E}\bar{\mathbf{X}}\| \leq \frac{M}{\sqrt{K}} \left(1 + \sqrt{2 \log \frac{1}{\delta}}\right)$$

Proof. For proof, see [35]. □

Now consider a space of functions,

$$\mathcal{F}_p \equiv \left\{ f(x) = \int_{\Omega} \alpha(\omega) \int_G \phi(gx, \omega) q(g) d\nu(g) d\omega \mid |\alpha(\omega)| \leq Cp(\omega) \right\},$$

and also consider another space of functions,

$$\hat{\mathcal{F}}_p \equiv \left\{ \hat{f}(x) = \sum_{k=1}^s \alpha_k \frac{1}{r} \sum_{i=1}^r \phi(g_i x, \omega_k) \mid |\alpha_k| \leq \frac{C}{s} \right\},$$

where $\phi(gx, \omega) = e^{-i\langle gx, \omega \rangle}$.

Lemma A.2. Let μ be a measure defined on X , and f^* a function in \mathcal{F}_p . If $\omega_1, \omega_2 \dots \omega_s$ are iid samples from $p(\omega)$, then for $\delta_1, \delta_2 > 0$, there exists a function $\hat{f} \in \hat{\mathcal{F}}_p$ such that

$$\|f^* - \hat{f}\|_{\mathcal{L}_2(X, \mu)} \leq \frac{C}{\sqrt{s}} \left(1 + \sqrt{2 \log \frac{1}{\delta_1}}\right) + \frac{C}{\sqrt{r}} \left(1 + \sqrt{2 \log \frac{1}{\delta_2}}\right),$$

with probability at least $1 - \delta_1 - \delta_2$.

Proof. Consider $\psi(x; \omega_k) = \int_G \phi(gx, \omega_k) q(g) d\nu(g)$. Let $\tilde{f}_k = \beta_k \psi(\cdot; \omega_k)$, $k = 1 \dots s$, with $\beta_k = \frac{\alpha(\omega_k)}{p(\omega_k)}$. Hence $\mathbb{E}_{\omega_k \sim p} \tilde{f}_k = f^*$.

Define $\tilde{f}(x) = \frac{1}{s} \sum_{k=1}^s \tilde{f}_k$. Let $\hat{f}_k(x) = \beta_k \hat{\psi}(x; \omega_k)$, where $\hat{\psi}(x; \omega_k) = \frac{1}{r} \sum_{i=1}^r \phi(g_i x, \omega_k)$ is the empirical estimate of $\psi(x; \omega_k)$. Define $\hat{f}(x) = \frac{1}{s} \sum_{k=1}^s \hat{f}_k(x)$. We have $\mathbb{E}_{g_i \sim q} \hat{f}(x) = \tilde{f}(x)$.

$$\|f^* - \hat{f}\|_{\mathcal{L}_2(X, \mu)} \leq \|f^* - \tilde{f}\|_{\mathcal{L}_2(X, \mu)} + \|\tilde{f} - \hat{f}\|_{\mathcal{L}_2(X, \mu)}$$

From Lemma 1 of [35], with probability $1 - \delta_1$,

$$\|f^* - \tilde{f}\|_{\mathcal{L}_2(X, \mu)} \leq \frac{C}{\sqrt{s}} \left(1 + \sqrt{2 \log \frac{1}{\delta_1}}\right).$$

Since $\hat{f}(x) = \frac{1}{r} \sum_{i=1}^r \sum_{k=1}^s \frac{\beta_k}{s} \phi(g_i x, \omega_k)$ and $\mathbb{E}_{g_i \sim q} \hat{f}(x) = \tilde{f}(x)$ with g_i iid (and $\{\omega_k\}_{k=1}^s$ fixed beforehand), we can apply Lemma A.1 with

$$M = \left\| \sum_{k=1}^s \frac{\beta_k}{s} \phi(g_i x, \omega_k) \right\| \leq \sum_{k=1}^s \left| \frac{\beta_k}{s} \right| \|\phi(g_i x, \omega_k)\| \leq \sum_{k=1}^s \left| \frac{\beta_k}{s} \right| \leq C.$$

We conclude that with a probability at least $1 - \delta_2$,

$$\|\tilde{f} - \hat{f}\|_{\mathcal{L}_2(X, \mu)} \leq \frac{C}{\sqrt{r}} \left(1 + \sqrt{2 \log \frac{1}{\delta_2}}\right).$$

Hence, with probability at least $1 - \delta_1 - \delta_2$, we have

$$\|f^* - \hat{f}\|_{\mathcal{L}_2(X, \mu)} \leq \frac{C}{\sqrt{s}} \left(1 + \sqrt{2 \log \frac{1}{\delta_1}}\right) + \frac{C}{\sqrt{r}} \left(1 + \sqrt{2 \log \frac{1}{\delta_2}}\right)$$

□

Theorem A.3 (Estimation error [35]). *Let \mathcal{F} be a bounded class of functions, $\sup_{x \in X} |f(x)| \leq C$ for all $f \in \mathcal{F}$. Let $V(y_i f(x_i))$ be an L -Lipschitz loss. Then with probability $1 - \delta$, with respect to training samples $\{x_i, y_i\}_{i=1,2,\dots,N}$ ($iid \sim P$), every f satisfies*

$$\mathcal{E}_V(f) \leq \hat{\mathcal{E}}_V(f) + 4LR_N(\mathcal{F}) + \frac{2|V(0)|}{\sqrt{N}} + LC\sqrt{\frac{1}{2N} \log \frac{1}{\delta}},$$

where $\mathcal{R}_N(\mathcal{F})$ is the Rademacher complexity of the class \mathcal{F} :

$$\mathcal{R}_N(\mathcal{F}) = \mathbb{E}_{x, \sigma} \left[\sup_{f \in \mathcal{F}} \left| \frac{1}{N} \sum_{i=1}^N \sigma_i f(x_i) \right| \right],$$

and σ_i are iid symmetric Bernoulli random variables taking value in $\{-1, 1\}$, with equal probability and are independent from x_i .

Proof. See in [35].

□

Let $f \in \mathcal{F}_p$ and $\hat{f} \in \hat{\mathcal{F}}_p$ then the approximation error is bounded as

$$\begin{aligned} \mathcal{E}_V(\hat{f}) - \mathcal{E}_V(f) &\leq \mathbb{E}_{(x,y) \sim P} \left| V(y\hat{f}(x)) - V(yf(x)) \right| \\ &\leq L\mathbb{E}|\hat{f}(x) - f(x)| \\ &\leq L\sqrt{\mathbb{E}(\hat{f}(x) - f(x))^2} \quad (\text{Jensen's inequality for } \sqrt{\cdot} \text{ concave function}) \\ &\leq LC \left(\frac{1}{\sqrt{s}} \left(1 + \sqrt{2 \log \frac{1}{\delta_1}}\right) + \frac{1}{\sqrt{r}} \left(1 + \sqrt{2 \log \frac{1}{\delta_2}}\right) \right), \end{aligned}$$

with probability at least $1 - \delta_1 - \delta_2$. Now let $f_N^* = \arg \min_{f \in \hat{\mathcal{F}}_p} \hat{\mathcal{E}}_V(f)$ and $\tilde{f} = \arg \min_{f \in \hat{\mathcal{F}}_p} \mathcal{E}_V(f)$. We have

$$\begin{aligned} \mathcal{E}_V(f_N^*) - \min_{f \in \hat{\mathcal{F}}_p} \mathcal{E}_V(f) &= \hat{\mathcal{E}}_V(f_N^*) - \mathcal{E}_V(\tilde{f}) + \mathcal{E}_V(\tilde{f}) - \min_{f \in \hat{\mathcal{F}}_p} \mathcal{E}_V(f) \\ &\leq 2 \sup_{\tilde{f} \in \hat{\mathcal{F}}_p} \left| \mathcal{E}_V(\tilde{f}) - \hat{\mathcal{E}}_V(\tilde{f}) \right| + L \left(\frac{C}{\sqrt{s}} \left(1 + \sqrt{2 \log \frac{1}{\delta_1}}\right) + \frac{C}{\sqrt{r}} \left(1 + \sqrt{2 \log \frac{1}{\delta_2}}\right) \right) \\ &\leq 2 \left(4LR_N(\mathcal{F}) + \frac{2|V(0)|}{\sqrt{N}} + LC\sqrt{\frac{1}{2N} \log \frac{1}{\delta}} \right) + \\ &\quad LC \left(\frac{1}{\sqrt{s}} \left(1 + \sqrt{2 \log \frac{1}{\delta_1}}\right) + \frac{1}{\sqrt{r}} \left(1 + \sqrt{2 \log \frac{1}{\delta_2}}\right) \right), \end{aligned}$$

with probability at least $1 - \delta - \delta_1 - \delta_2$. It is easy to show that $\mathcal{R}_N(\mathcal{F}) \leq \frac{C}{\sqrt{N}}$. Taking $\delta = \delta_1 = \delta_2$ yields the statement of the theorem.

□