

1. Appendix

1.1. Proof of Lemma 1

It is straight forward to see:

$$\begin{aligned}
\mathbb{E}\bar{H}_{t+1} &= \mathbb{E}\left[\frac{1}{n} \sum_{i=1}^n f_i^\delta(\phi_i^{t+1})\right] = \mathbb{E}\left[\frac{1}{n} \left(\sum_{i \in \mathcal{B}} f_i^\delta(w^t) + \sum_{i \notin \mathcal{B}} f_i^\delta(\phi_i^t) \right)\right] \\
&= \mathbb{E}\left[\mathbb{E}\left[\frac{1}{n} \left(\sum_{i \in \mathcal{B}} f_i^\delta(w^t) + \sum_{i \notin \mathcal{B}} f_i^\delta(\phi_i^t) \right) \middle| \mathcal{B} \right] \middle| |\mathcal{B}| = B\right] \\
&= \frac{1}{n} \left(\frac{\mathbb{E}|\mathcal{B}|}{n} \sum_{i=1}^n f_i^\delta(w^t) + \frac{n - \mathbb{E}|\mathcal{B}|}{n} \sum_{i=1}^n f_i^\delta(\phi_i^t) \right) \\
&= \frac{\mathbb{E}|\mathcal{B}|}{n} f^\delta(w^t) + \frac{n - \mathbb{E}|\mathcal{B}|}{n} \bar{H}_t
\end{aligned}$$

The second line of equality comes from the rule of total expectation, where the inner expectation is taken over the index set \mathcal{B} , and the outer expectation is taken over the set cardinality $|\mathcal{B}|$.

1.2. Proof of Lemma 2

The proof technique is similar to SAGA, as well as a useful inequality (Lemma 4 in (?)):

$$\begin{aligned}
f(x) &\geq f(y) + \langle f'(y), x - y \rangle + \frac{1}{2(L - \mu)} \|f'(x) - f'(y)\|^2 \\
&\quad + \frac{\mu L}{2(L - \mu)} \|x - y\|^2 - \frac{\mu}{(L - \mu)} \langle f'(x) - f'(y), x - y \rangle.
\end{aligned} \tag{A1}$$

First of all, by the update rule (2):

$$\begin{aligned}
\|w^{t+1} - w^*\|^2 &= \|\text{Prox}_{\gamma g}(w^t - \gamma G(w^t)) - \text{Prox}_{\gamma g}(w^* - \gamma f'(w^*))\|^2 \\
&\leq \|w^t - \gamma G(w^t) - w^* + \gamma f'(w^*)\|^2 \\
&= \|w^t - w^*\|^2 - 2\gamma \langle w^t - w^*, G(w^t) - f'(w^*) \rangle + \gamma^2 \|G(w^t) - f'(w^*)\|^2.
\end{aligned} \tag{A2}$$

The inequality follows from non-expansiveness of proximal operator, notice that our stochastic gradient $G(w^t)$ is unbiased, take the expectation to the second term and apply (A1) to each f_i and the average over all i will goes to:

$$\begin{aligned}
-\mathbb{E}[\langle w^t - w^*, G(w^t) - f'(w^*) \rangle] &= -\langle w^t - w^*, f'(w^t) - f'(w^*) \rangle \\
&\leq \langle w^t - w^*, f'(w^*) \rangle + \frac{L - \mu}{L} [f(w^*) - f(w^t)] - \frac{\mu}{2} \|w^* - w^t\|^2 \\
&\quad - \frac{1}{2Ln} \sum_{i=1}^n \|f'_i(w^*) - f'_i(w^t)\|^2 - \frac{\mu}{L} \langle f'(w^*), w^t - w^* \rangle.
\end{aligned} \tag{A3}$$

Next we bound the last term in (A2):

$$\begin{aligned}
\mathbb{E} \|f'(w^*) - G(w^t)\|^2 &= \mathbb{E} \left\| f'(w^*) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(\phi_i^t) - \frac{1}{n} \sum_{i=1}^n f'_i(\phi_i^t) \right\|^2 \\
&= \mathbb{E} \left\| \left[\frac{1}{n} \sum_{i=1}^n f'_i(w^t) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^t) - f'(w^*) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^*) \right] \right. \\
&\quad \left. - \left[\frac{1}{n} \sum_{i=1}^n f'_i(\phi_i^t) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(\phi_i^t) - f'(w^*) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^*) \right] \right. \\
&\quad \left. + f'(w^*) - \frac{1}{n} \sum_{i=1}^n f'_i(w^t) \right\|^2 \tag{A4} \\
&\stackrel{*}{=} \mathbb{E} \left\| \left[\frac{1}{n} \sum_{i=1}^n f'_i(w^t) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^t) - f'(w^*) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^*) \right] \right. \\
&\quad \left. - \left[\frac{1}{n} \sum_{i=1}^n f'_i(\phi_i^t) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(\phi_i^t) - f'(w^*) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^*) \right] \right\|^2 \\
&\quad + \left\| f'(w^*) - \frac{1}{n} \sum_{i=1}^n f'_i(w^t) \right\|^2.
\end{aligned}$$

In equation $\stackrel{*}{=}$ we use the property that $\mathbb{E}[X^2] = \mathbb{E}[X - \mathbb{E}[X]]^2 + \mathbb{E}[X]^2$, now use the inequality $\|X + Y\|^2 \leq (1 + \beta)\|X\|^2 + (1 + \beta^{-1})\|Y\|^2$, $\beta > 0$ to the first term:

$$\begin{aligned}
\mathbb{E} \|f'(w^*) - G(w^t)\|^2 &\leq (1 + \beta) \mathbb{E} \left\| \frac{1}{n} \sum_{i=1}^n f'_i(w^t) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^t) - f'(w^*) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^*) \right\|^2 \\
&\quad + (1 + \beta^{-1}) \mathbb{E} \left\| \frac{1}{n} \sum_{i=1}^n f'_i(\phi_i^t) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(\phi_i^t) - f'(w^*) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^*) \right\|^2 \tag{A5} \\
&\quad + \beta \cdot \left\| f'(w^*) - \frac{1}{n} \sum_{i=1}^n f'_i(w^t) \right\|^2.
\end{aligned}$$

Next we bound the first and second terms again by variance decomposition, for simplicity we only take the first term as example:

$$\begin{aligned}
&\mathbb{E} \left\| \frac{1}{n} \sum_{i=1}^n f'_i(w^t) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^t) - f'(w^*) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^*) \right\|^2 \\
&= \mathbb{E} \left\| \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} (f'_i(w^*) - f'_i(w^t)) \right\|^2 - \left\| \frac{1}{n} \sum_{i=1}^n (f'_i(w^*) - f'_i(w^t)) \right\|^2 \\
&\stackrel{(1)}{\leq} \mathbb{E} \left(\frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} \|f'_i(w^*) - f'_i(w^t)\|^2 \right) - \left\| \frac{1}{n} \sum_{i=1}^n (f'_i(w^*) - f'_i(w^t)) \right\|^2 \tag{A6} \\
&= \frac{1}{n} \sum_{i=1}^n \|f'_i(w^*) - f'_i(w^t)\|^2 - \left\| \frac{1}{n} \sum_{i=1}^n (f'_i(w^*) - f'_i(w^t)) \right\|^2 \\
&\stackrel{(2)}{\leq} \frac{1}{n} \sum_{i=1}^n \|f'_i(w^*) - f'_i(w^t)\|^2,
\end{aligned}$$

(1) \leq is by RMS-AM inequality, and in (2) \leq we drop the negative term. Similarly,

$$\mathbb{E} \left\| \frac{1}{n} \sum_{i=1}^n f'_i(\phi_i^t) - \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(\phi_i^t) - f'(w^*) + \frac{1}{|\mathcal{B}|} \sum_{i \in \mathcal{B}} f'_i(w^*) \right\|^2 \leq \frac{1}{n} \sum_{i=1}^n \|f'_i(w^*) - f'_i(\phi_i^t)\|^2.$$

Plug (A6) into (A5) we get:

$$\mathbb{E} \|f'(w^*) - G(w^t)\|^2 \leq \frac{(1+\beta)}{n} \sum_{i=1}^n \|f'_i(w^*) - f'_i(w^t)\|^2 + \frac{(1+\beta^{-1})}{n} \sum_{i=1}^n \|f'_i(w^*) - f'_i(\phi_i^t)\|^2 - \beta \|f'(w^t) - f'(w^*)\|^2. \quad (\text{A7})$$

Combining (A2),(A3),(A7) becomes (5) immediately:

$$\begin{aligned} \|w^t - w^*\|^2 - \mathbb{E}\|w^{t+1} - w^*\|^2 &\geq \gamma\mu \|w^t - w^*\|^2 - (2\gamma^2 - \gamma/L)\mathbb{E}\|f'_i(w^t) - f'_i(w^*)\|^2 \\ &\quad + \gamma^2 \|f'(w^t) - f'(w^*)\|^2 + \frac{2\gamma(L-\mu)}{L} f^\delta(w^t) - 4\gamma^2 L \bar{H}_t. \end{aligned}$$

1.3. Proof of Theorem 3

It follows directly from Lemma 1 and 2:

$$\begin{aligned} \mathcal{L}_t - \mathbb{E}\mathcal{L}_{t+1} &= c(\bar{H}_t - \mathbb{E}\bar{H}_{t+1}) + (\|w^t - w^*\|^2 - \mathbb{E}\|w^{t+1} - w^*\|^2) \\ &\geq c \left(\frac{\mathbb{E}|\mathcal{B}|}{n} - \frac{2(1+\beta^{-1})\gamma^2 L}{c} \right) \bar{H}_t + \gamma\mu \|w^t - w^*\|^2 + (2\mu\beta\gamma^2 + \frac{2\gamma(L-\mu)}{L} - \frac{c \cdot \mathbb{E}|\mathcal{B}|}{n}) f^\delta(w^t) \\ &\quad + \left(\frac{\gamma}{L} - (1+\beta)\gamma^2 \right) \mathbb{E}\|f'_i(w^t) - f'(w^*)\|^2 \\ &\stackrel{?}{\geq} c \left(\frac{|\mathcal{B}|}{n} - \frac{2(1+\beta^{-1})\gamma^2 L}{c} \right) \bar{H}_t + \gamma\mu \|w^t - w^*\|^2 \\ &\geq \rho \mathcal{L}_t, \end{aligned} \quad (\text{A8})$$

where $\rho = \min(\frac{|\mathcal{B}|}{n} - \frac{2(1+\beta^{-1})\gamma^2 L}{c}, \gamma\mu)$, the last inequality $\stackrel{?}{\geq}$ comes with following condition:

$$\begin{aligned} 2\mu\beta\gamma^2 + \frac{2\gamma(L-\mu)}{L} - \frac{c|\mathcal{B}|}{n} &\geq 0 \\ \frac{\gamma}{L} - (1+\beta)\gamma^2 &\geq 0, \end{aligned} \quad (\text{A9})$$

furthermore, to keep our algorithm moving forward, i.e. $\|w^t - w^*\|^2$ decreasing, we should also make sure such condition hold:

$$\frac{|\mathcal{B}|}{n} - \frac{2(1+\beta^{-1})\gamma^2 L}{c} \geq 0. \quad (\text{A10})$$

1.4. Proof of Proposition 1

By plugging $\beta = 2$, $c = \frac{n}{3L\mathbb{E}|\mathcal{B}|}$ into (A9) it is easy to verify both inequalities hold.

1.5. Proof of Proposition 2

In this case we choose $\beta = 1$. From Theorem 3 we know that with a suitable step size γ and c , we have:

$$\mathbb{E}\|w^t - w^*\|^2 \leq \mathbb{E}\mathcal{L}_t \leq (1-\rho)^t \mathcal{L}_0 = (1-\rho)^t [\|w^0 - w^*\|^2 + c\bar{H}_0].$$

For the optimal convergence rate, we try to maximize the geometric factor $\rho = \min(\frac{\mathbb{E}|\mathcal{B}|}{n} - \frac{4\gamma^2 L}{c}, \gamma\mu)$. Denote γ_0 as the solution of: $\frac{\mathbb{E}|\mathcal{B}|}{n} - \frac{4\gamma_0^2 L}{c} = \gamma_0\mu$. Notice that $\rho(\gamma) = \gamma\mu$ is increasing with γ when $\gamma \leq \gamma_0 = \frac{c}{8\kappa} \left(\sqrt{1 + \frac{16\kappa\mathbb{E}|\mathcal{B}|}{cn\mu}} - 1 \right)$

and $\rho(\gamma) = \frac{\mathbb{E}|\mathcal{B}|}{n} - \frac{4\gamma^2 L}{c}$ is decreasing when $\gamma > \gamma_0$. So the optimal step size should be $\gamma = \gamma_0$. However we should also verify that this step size indeed satisfies the condition in (A9). First of all:

$$\gamma_0 = \frac{c}{8\kappa} \left(\sqrt{1 + \frac{16\kappa\mathbb{E}|\mathcal{B}|}{cn\mu}} - 1 \right) \stackrel{(1)}{\leq} \frac{c}{8\kappa} \sqrt{\frac{16\kappa\mathbb{E}|\mathcal{B}|}{cn\mu}} = \sqrt{\frac{c\mathbb{E}|\mathcal{B}|}{4nL}} \stackrel{(2)}{\leq} \frac{1}{2L}. \quad (\text{A11})$$

(1) comes from the fact that $\sqrt{1+x} - 1 \leq \sqrt{x}$, (2) holds by choosing $c = \frac{\tau n}{L\mathbb{E}|\mathcal{B}|}$, where $\tau < 1$ is a small constant. These two inequalities together ensure the upper bound part of (A9). As to the lower bound, we have $\sqrt{1+x} - 1 > \sqrt{x} - 1$, so:

$$\gamma_0 > \frac{c}{8\kappa} \left(\sqrt{\frac{16\kappa\mathbb{E}|\mathcal{B}|}{cn\mu}} - 1 \right) \geq \frac{c\mathbb{E}|\mathcal{B}|L}{2n(L-\mu)} \implies \tau \leq \left(\frac{1}{\frac{L}{L-\mu} + \frac{n}{4\kappa\mathbb{E}|\mathcal{B}|}} \right)^2 < 1.$$

So if we choose τ properly, both sides of (A9) can be satisfied.

1.6. Proof of Corollary 1, 2

Following (7) we take a derivative to $\mathbb{E}|\mathcal{B}|$:

$$\frac{\partial f(\mathbb{E}|\mathcal{B}|)}{\partial \mathbb{E}|\mathcal{B}|} = \frac{(\alpha\mathbb{E}|\mathcal{B}| - 1)^2 \mathbb{E}|\mathcal{B}|}{\sqrt{1 + \alpha^2 \mathbb{E}|\mathcal{B}|^2} (\sqrt{1 + \alpha^2 \mathbb{E}|\mathcal{B}|^2} - 1)^2} \geq 0, \quad (\text{A12})$$

where $\alpha = \frac{4\kappa}{\sqrt{\tau n}}$, so there is no optimal batch size, and since we always want to access one data point, i.e. $|\mathcal{B}| \geq 1$ and SAGA style update is optimal.

For Corollary 2, it is easy to see for our algorithm, which choose $|\mathcal{B}| = n$ with probability $p \ll 1$ and $|\mathcal{B}| = 1$ with probability $1 - p$, has average batch size $\mathbb{E}|\mathcal{B}| = np + 1 - p \approx np + 1$. For each update, it takes on average time $\tau = n\eta\tau p + (1-p)\tau \approx (1 + n\eta)\tau$. If we want to get a ϵ -suboptimal solution, the total iteration will be $N = \frac{\log(\epsilon/\epsilon_0)}{\log(1-\rho)} \propto 1/\rho$, So the running time will be:

$$\begin{aligned} T &\propto \frac{1 + n\eta}{\sqrt{\frac{1}{\mathbb{E}|\mathcal{B}|^2} + \frac{16\kappa^2}{\tau n^2} - \frac{1}{\mathbb{E}|\mathcal{B}|}}} \\ &\approx \frac{(\mathbb{E}|\mathcal{B}|^2 - \mathbb{E}|\mathcal{B}|)\eta + \mathbb{E}|\mathcal{B}|}{\sqrt{1 + \alpha^2 \mathbb{E}|\mathcal{B}|^2} - 1}. \end{aligned} \quad (\text{A13})$$

For simplicity we denote $B = \mathbb{E}|\mathcal{B}|$. By taking the partial derivative and set it to zero $\partial T / \partial B = 0$ can solve the best batch size:

$$((2B - 1)\eta + 1)(\sqrt{1 + \alpha^2 B^2} - 1) = ((B^2 - B)\eta + B) \frac{\alpha^2 B}{\sqrt{1 + \alpha^2 B^2}}, \quad (\text{A14})$$

after solving the above equation we get:

$$B = \left(\frac{1}{\eta} - 1 \right) \left(\frac{\xi - 1}{2 - \xi} \right), \quad \xi = \frac{\alpha^2 B^2}{1 + \alpha^2 B^2 - \sqrt{1 + \alpha^2 B^2}}. \quad (\text{A15})$$

By showing the second order derivative $\partial^2 T / \partial B^2 \geq 0$ it's easy to verify that this solution is actually a global minimum.

1.7. Proof of Lemma 4

We begin with non-expansiveness of proximal operation:

$$\begin{aligned} \|w^{t+1} - w^*\|^2 &= \|\text{Prox}_{\gamma g}(w^t - \gamma G(w^t)) - \text{Prox}_{\gamma g}(w^* - \gamma f'(w^*))\|^2 \\ &\leq \|w^t - \gamma G(w^t) - w^* + \gamma f'(w^*)\|^2 \\ &= \|w^t - w^*\|^2 - 2\gamma \langle w^t - w^*, G(w^t) - f'(w^*) \rangle + \gamma^2 \|G(w^t) - f'(w^*)\|^2, \end{aligned} \quad (\text{A16})$$

where $f(w) = \frac{1}{n} \sum_{i=1}^n f_i(w)$ By taking expectation on each side and notice $G(w^t)$ is a unbiased estimation of $f'(w^t)$:

$$\mathbb{E}\|w^{t+1} - w^*\|^2 = \|w^t - w^*\|^2 - 2\gamma \langle w^t - w^*, f'(w^t) - f'(w^*) \rangle + \gamma^2 \mathbb{E}\|G(w^t) - f'(w^*)\|^2, \quad (\text{A17})$$

and then apply the following bounds for strongly convex function f :

$$\begin{aligned} \langle w^t - w^*, f'(w^t) - f'(w^*) \rangle &\geq \mu \|w^t - w^*\|^2 \\ \langle w^t - w^*, f'(w^t) - f'(w^*) \rangle &\geq \frac{1}{L} \|f'(w^t) - f'(w^*)\|^2, \end{aligned} \quad (\text{A18})$$

so the inner product term have a composite upper bound:

$$-2\gamma \langle w^t - w^*, f'(w^t) - f'(w^*) \rangle \leq -\gamma(\mu \|w^t - w^*\|^2 + \frac{1}{L} \|f'(w^t) - f'(w^*)\|^2) \quad (\text{A19})$$

on the other hand, we can bound $\mathbb{E}\|G(w^t) - f'(w^*)\|^2$ as (A5) but we only need to care about one sample in a batch case, since we are comparing SAGA with SVRG update style:

$$\mathbb{E}\|G(w^t) - f'(w^*)\|^2 \leq 2\mathbb{E}\|f'_i(\phi_i^t) - f'_i(w^*)\|^2 + 2\mathbb{E}\|f'_i(w^t) - f'_i(w^*)\|^2 - \|f'(w^t) - f'(w^*)\|^2. \quad (\text{A20})$$

Remember we have proved above formula in (A7), for $\mathbb{E}\|f'_i(w^t) - f'_i(w^*)\|^2$ we have:

$$\begin{aligned} \mathbb{E}\|f'_i(w^t) - f'_i(w^*)\|^2 &\leq \frac{2L}{n} \sum_{i=1}^n f_i(w^t) - f_i(w^*) - f'_i(w^*)^\top (w^t - w^*) \\ &= 2L(f(w^t) - f(w^*) - f'(w^*)^\top (w^t - w^*)). \end{aligned} \quad (\text{A21})$$

Similarly, for $\|f'(w^t) - f'(w^*)\|^2$ we recall f is a μ -strongly convex function:

$$\|f'(w^t) - f'(w^*)\|^2 \geq 2\mu(f(w^t) - f(w^*) - f'(w^*)^\top (w^t - w^*)). \quad (\text{A22})$$

Add those inequalities together:

$$\mathbb{E}\|w^{t+1} - w^*\|^2 \leq (1 - \gamma\mu)\|w^t - w^*\|^2 + (4L\gamma^2 - \frac{2\mu\gamma}{L} - 2\mu\gamma^2)f^\delta(w^t) + 2\gamma^2\mathbb{E}\|f'_i(\phi_i^t) - f'_i(w^*)\|^2. \quad (\text{A23})$$

1.8. Proof of Lemma 5

Since we know the distribution of random variable τ , also denote t_s as the index of the latest gradient snapshot so for SVRG/SAGA++ $t_s = kT$ where k is the number of outer iteration and T is the length of inner iteration, for SAGA $t_s = 0$ so in either method we have $t_s \geq 0$ then by conditional expectation relationship:

$$\begin{aligned} \mathbb{E}[\|\alpha_i - f'_i(w^*)\|^2 | \mathcal{F}_0] &\stackrel{(1)}{=} \frac{1}{n} \sum_{k=1}^n \mathbb{E}[\|\alpha_k - f'_k(w^*)\|^2 | \mathcal{F}_{t_s}] \\ &\stackrel{(2)}{=} \frac{1}{n} \sum_{k=1}^n \sum_{l=t_s}^t p_l \|f'_k(w^l) - f'_k(w^*)\|^2 \\ &= \sum_{l=t_s}^t p_l \frac{1}{n} \sum_{k=1}^n \|f'_k(w^l) - f'_k(w^*)\|^2 \\ &\leq 2L \sum_{l=t_s}^t p_l (f(w^l) - f(w^*) - f'(w^*)^\top (w^l - w^*)), \end{aligned} \quad (\text{A24})$$

$\stackrel{(1)}{=}$ is taken over the choices of i , while $\stackrel{(2)}{=}$ is taken over the random variable τ in $\alpha_k = f'_k(w^\tau)$. Because the regularization function $g(w)$ is convex, and from optimal condition we know: $-f'(w^*) \in \partial g(w^*)$, we have:

$$\begin{aligned} f(w^l) - f(w^*) - f'(w^*)^\top (w^l - w^*) &= f(w^l) - f(w^*) + v^l (w^l - w^*) \\ &\leq f(w^l) - f(w^*) + g(w^l) - g(w^*) \\ &= F(w^l) - F(w^*), \end{aligned} \quad (\text{A25})$$

where $v^l \in \partial g(w^l)$. Finally we have $\mathbb{E}[\|\alpha_i - f'_i(w^*)\|^2 | \mathcal{F}_0] \leq 2L \sum_{l=t_s}^t p_l (F(w^l) - F(w^*)).$

1.9. Proof of Proposition 3

Recall the quadratic upper bound of L -Lipschitz function:

$$f(w^t - \gamma G(w^t)) \leq f(w^t) - \gamma \nabla f^\top(w^t) G(w^t) + \frac{L\gamma^2}{2} \|G(w^t)\|^2. \quad (\text{A26})$$

By taking the expectation,

$$\begin{aligned} \mathbb{E}[f(w^t - \gamma G(w^t)) | \mathcal{F}_t] &\leq f(w^t) - \gamma \| \nabla f(w^t) \|^2 + \frac{L\gamma^2}{2} \mathbb{E}[\|G(w^t)\|^2 | \mathcal{F}_t] \\ &\leq f(w^t) - (\gamma - \frac{L\gamma^2}{2}) \| \nabla f(w^t) \|^2 + \frac{L\gamma^2}{2} \text{Var}[G(w^t)]. \end{aligned} \quad (\text{A27})$$

On the other hand, for μ -strongly convex f , we have:

$$\| \nabla f(w^t) \|^2 \geq 2\mu(f(w^t) - f^*), \quad (\text{A28})$$

so if $\text{Var}[G(w^t)]$ also converges to zero at the order of $f^{\text{sub}}(w^t) = f(w^t) - f^*$ then γ can keep to a small constant rather than damping like SGD. In fact (?) (Corollary 3) already proved it for SVRG, here we prove a similar result for SAGA style update:

$$\begin{aligned} \text{Var}[G(w^t) | \mathcal{F}_s] &= \mathbb{E} \left[\left\| \nabla f_{i_k}(w^t) - \nabla f_{i_k}(\phi_{i_k}^t) - \frac{1}{n} \sum_{j=1}^n (\nabla f_j(w^t) - \nabla f_j(\phi_j^t)) \right\|^2 \middle| \mathcal{F}_s \right] \\ &= \mathbb{E} \left[\left\| \nabla f_{i_k}(w^t) - \nabla f_{i_k}(\phi_{i_k}^t) \right\|^2 \middle| \mathcal{F}_s \right] - \left\| \frac{1}{n} \sum_{j=1}^n (\nabla f_j(w^t) - \nabla f_j(\phi_j^t)) \right\|^2 \\ &\leq \mathbb{E} \left[\mathbb{E} \left[\left\| \nabla f_{i_k}(w^t) - \nabla f_{i_k}(\phi_{i_k}^t) \right\|^2 \middle| \mathcal{F}_t, \mathcal{F}_s \right] \right] \\ &= \frac{2}{n} \sum_{j=1}^n \mathbb{E} [\| \nabla f_j(w^t) - \nabla f_j(w^*) \|^2 | \mathcal{F}_s] + \frac{2}{n} \sum_{j=1}^n \mathbb{E} [\| \nabla f_j(\phi_j^t) - \nabla f_j(w^*) \|^2 | \mathcal{F}_s] \\ &\leq 4L(\mathbb{E}[f(w^t) | \mathcal{F}_s] - f(w^*)) + \frac{4L}{n} \sum_{j=1}^n \sum_{\tau=s}^t p_\tau (\mathbb{E}[f_j(w_\tau) | \mathcal{F}_s] - f_j(w^*)) \\ &= 4L(\mathbb{E}[f(w^t) | \mathcal{F}_s] - f(w^*)) + 4L \sum_{\tau=s}^t p_\tau (\mathbb{E}[f(w_\tau) | \mathcal{F}_s] - f(w^*)), \end{aligned} \quad (\text{A29})$$

here $\{\mathcal{F}_t\}_{t \geq 0}$ is the filtered probability space, $t - T \leq s \leq t$ (recall T is the length of inner iteration) is the latest available full gradient time stamp, p_τ is the probability distribution of stored gradient discussed in (10). Since $t - s$ is upper bounded (this is true for SVRG/SAGA++, as to SAGA, the expectation is $n \log n$ by ‘‘Coupon collection problem’’), together with linear convergence, we know the second term is close to the first term up to a constant.

1.10. Proof of Theorem 6

First of all, we have the following recursive formula:

$$\begin{aligned} P_g(x, \eta, c, n) &= \text{Prox}_g(P_g(x, \eta, c, n - 1) - c) \\ &= \begin{cases} P(x, \eta, c, n - 1) - c - \eta, & \text{if } P(x, \eta, c, n - 1) \geq c + \eta \\ 0, & \text{if } c - \eta \leq P(x, \eta, c, n - 1) \leq c + \eta \\ P(x, \eta, c, n - 1) - c + \eta, & \text{if } P(x, \eta, c, n - 1) \leq c - \eta \end{cases} \end{aligned} \quad (\text{A30})$$

Because c can be either positive or negative but η is always positive, we consider about following cases:

- ($c < -\eta$) In this case $0 > c + \eta > c - \eta$, if:

-
1. $x \geq c + \eta$, then $P(x, \eta, c, n) = x - n(c + \eta)$;
 2. $x < c + \eta$, then suppose $x = q(c - \eta) + \epsilon$, $q \in \mathbb{N}$, $\epsilon \in [c - \eta, c + \eta]$, if $q \geq n$ then $P(x, \eta, c, n) = x - n(c - \eta)$; else $P(x, \eta, c, q) = \epsilon$, $P(x, \eta, c, q + 1) = 0$, $P(x, \eta, c, n) = -(n - q - 1)(c + \eta)$.

• ($c > \eta$) In this case $0 < c - \eta < c + \eta$ which is symmetric to previous case, if:

1. $x \leq c - \eta$, then $P(x, \eta, c, n) = x - n(c - \eta)$;
2. $x > c - \eta$, then suppose $x = q(c + \eta) + \epsilon$, $q \in \mathbb{N}$, $\epsilon \in [c - \eta, c + \eta]$, if $q \geq n$ then $P(x, \eta, c, n) = x - n(c - \eta)$; else $P(x, \eta, c, q) = \epsilon$, $P(x, \eta, c, q + 1) = 0$, $P(x, \eta, c, n) = -(n - q - 1)(c - \eta)$.

• ($-\eta \leq c \leq \eta$) finally, $c - \eta \leq 0 \leq c + \eta$, if:

1. $x \geq n(c + \eta)$, then $P(x, \eta, c, n) = x - n(c + \eta)$;
2. $x \leq n(c - \eta)$, then $P(x, \eta, c, n) = x + n(c - \eta)$;
3. otherwise, $\lfloor \frac{x}{c + \eta} \rfloor < n$ or $\lfloor \frac{-x}{-c + \eta} \rfloor < n$ then we know it will eventually be zero: $P(x, \eta, c, n) = 0$.

Clearly this is a piecewise linear function with tangent either 1 or 0.

1.11. ℓ_2 Logistic Regression Experiment

In this supplemental experiment, we conduct the ℓ_2 logistic regression experiment, formulated as follows

$$w^* = \arg \min_w \frac{1}{n} \sum_{i=1}^n \log(1 + \exp(y_i x_i^T w)) + \frac{\lambda}{2} \|w\|_2^2. \quad (\text{A31})$$

The datasets and settings are the same as ℓ_1 experiment discussed in the main text. The experiment result is exhibited in Figure 1.

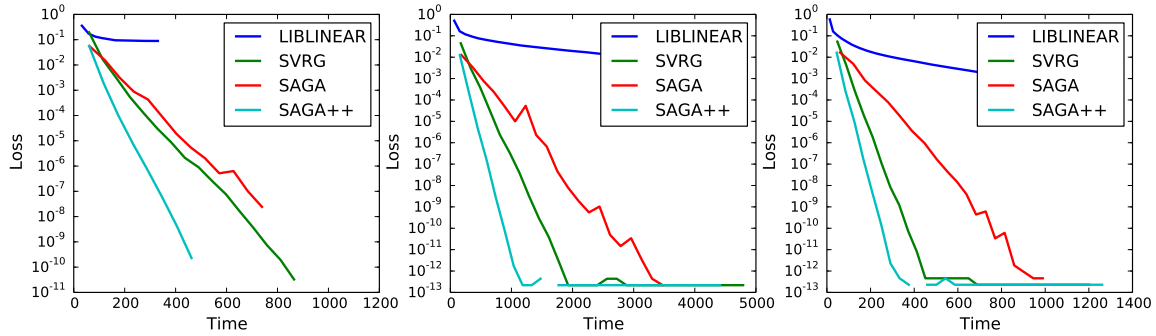


Figure 1. Running time comparison among different data ($\lambda = 1.0 \times 10^{-7}$ for all data).