

Mondrian Conformal Predictive Distributions

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Abstract

The distributions output by a standard (non-normalized) conformal predictive system all have the same shape but differ in location, while a normalized conformal predictive system outputs distributions that differ also in shape, through rescaling. An approach to further increasing the flexibility of the framework is proposed, called *Mondrian conformal predictive distributions*, which are (standard or normalized) conformal predictive distributions formed from multiple Mondrian categories. The effectiveness of the approach is demonstrated with an application to regression forests. By forming categories through binning of the predictions, it is shown that for this model class, the use of Mondrian conformal predictive distributions significantly outperforms the use of both standard and normalized conformal predictive distributions with respect to the continuous-ranked probability score. It is further shown that the use of Mondrian conformal predictive distributions results in as tight prediction intervals as produced by normalized conformal regressors, while improving upon the point predictions of the underlying regression forest.

Keywords: Conformal predictive systems, Conformal predictive distributions, Mondrian conformal predictive distributions, Continuous ranked probability score

1. Introduction

In contrast to a standard regression model, which outputs point predictions, and a conformal regressor, which outputs prediction intervals, a conformal predictive system for regression outputs cumulative probability distributions over the possible target values (Vovk et al., 2020). Following the terminology used in conjunction with conformal regressors (Johansson et al., 2014a), one may distinguish between *standard* and *normalized* conformal predictive systems, where the latter employ a quality estimate (σ_i) that is specific to each object (\mathbf{x}_i) for which a prediction is to be made, while the former use the same $\sigma_i = c$ for all objects, where c is some positive constant. As a consequence, distributions output by a standard conformal predictive system may differ only in their location and not in their shape. The distributions output by a normalized conformal predictive system may differ also in shape, as they are rescaled along the target value dimension using the quality estimates; the distribution for a high-quality prediction will be scaled down along this dimension,

resulting in a steeper increase of the cumulative probability distribution close to the point prediction, while the distribution of a low-quality prediction will be upscaled, resulting in a more uniform increase over the range of possible target values.

Conformal predictive systems can be seen as generalizations of both standard and conformal regressors, as they can easily be constrained to produce point predictions and prediction intervals, respectively. The latter can be obtained from a conformal predictive distribution through considering the lower and upper percentiles of interest, e.g., using the 2.5th and 97.5th percentiles for a 95% level of confidence. Conformal predictive systems can also be used for calibrating point predictions, e.g., by using the 50th percentile (median) or the mean of the distribution, rather than the prediction of the underlying model. The latter may be useful if the underlying model systematically under- or overestimates the target values, since using the median (or mean) of the distribution may correct for this by pushing the prediction upwards or downwards.

However, when the residuals (differences between actual and predicted values) are heteroscedastic, e.g., their distribution is not independent of the actual predictions, neither standard nor normalized predictive distributions may be very effective tools for calibrating the predictions. To see this, consider an underlying model that has a tendency to overestimate low actual values and underestimate high actual values. This means that the residuals for low predicted values will be negative on average, while they will be positive on average for high predicted values. A standard (non-normalized) conformal predictive distribution cannot correct for this, as the shape of the predictive distributions will be the same for all predictions. This means that the distribution can suggest moving the predictions at most in one direction, since the median (or mean) of the distribution will always be on the same side relative to the underlying prediction. Using a normalized conformal predictive distribution will not fix the problem, since the quality estimate used for scaling is always positive, and hence the direction of the correction suggested by the normalized conformal predictive distribution will be the same as for the standard distribution.

In order to overcome the above problem, an alternative approach to producing conformal predictive distributions is proposed, called *Mondrian conformal predictive distributions*. The approach borrows the idea of Mondrian conformal prediction (Vovk et al., 2005), which originally was proposed for allowing to control the error levels of objects falling into *a priori* defined categories. For example, by defining categories according to the class labels, one can guarantee the same error level for all classes. More recently, Boström and Johansson (2020) proposed *Mondrian conformal regressors*, where this idea was used to handle two problems of normalized conformal regressors; the prediction intervals may be several times larger (or smaller) than the largest (smallest) previously observed error, and the sizes of the intervals become less uniform with less informative quality (difficulty) estimates. By forming categories through binning of the quality estimates, it was shown that the size of the intervals is bounded by the size of the largest observed error and that non-informative quality estimates result in more uniformly sized intervals.

To the best of our knowledge, the Mondrian approach has however not been applied to conformal predictive systems. In this paper, we propose one such approach, which uses the categories to form multiple (standard or normalized) conformal predictive distributions. For the example above, we may consider using, say two, categories based on the predictions of the underlying model. For the first of these, which corresponds to low-valued predictions, the

conformal predictive distribution can correct for the systematic overestimations, while for the second category, which corresponds to high-valued predictions, the conformal predictive distribution can correct for the systematic underestimations. It should be noted that the prediction intervals output by conformal regressors do not directly allow for such corrections, as they are centered around the point predictions and provide no information on whether the true target can be expected to be higher or lower than the point prediction. In contrast to the previously proposed Mondrian conformal regressors, the categories for Mondrian conformal predictive distributions are here proposed to be formed using the predictions and not the quality estimates. The latter estimates may however still be used within the categories to form normalized conformal predictive distributions.

We will present results from a large-scale empirical evaluation of Mondrian conformal predictive systems, using regression forests as the underlying model. We will consider Mondrian distributions formed from categories corresponding to bins of the predicted values, where both standard and normalized conformal predictive distributions will be generated for the categories. The resulting Mondrian conformal predictive systems will be compared to standard and normalized conformal predictive systems, respectively, with respect to continuous ranked probability score. They will also be compared to standard and normalized conformal regressors on the tasks of generating efficient prediction intervals and accurate point predictions.

In the next section, we briefly describe conformal predictive distributions. In Section 3, we introduce the alternative Mondrian-based approach; Mondrian conformal predictive distributions. In Section 4, we first illustrate the approach and then present results from comparing the novel approach to standard and normalized conformal predictive systems, as well as to normalized (and standard) conformal regressors on a set of real-world datasets. Finally, in Section 5, we discuss the main findings and outline directions for future work.

2. Preliminaries

Conformal prediction was originally developed for the transductive case (Gammerman et al., 1998), requiring re-training of the underlying model for each new object to be predicted, something which often is computationally infeasible. Inductive conformal prediction (ICP) was proposed as a computationally less costly approach (Papadopoulos et al., 2002), requiring only one underlying model to be generated, at the cost of having to set aside part of the training examples for calibration, which leaves less examples to use for model building. A conformal predictive system utilizing the same idea is called a *split conformal predictive system* (Vovk et al., 2020).

A split conformal predictive system relies on an isotonic split conformity measure A to calculate a cumulative probability with respect to a label y_i , given some object \mathbf{x}_i and underlying model h .

Given a training set Z_{tr} , and a test object x , a *conformal predictive distribution* is constructed by a (*split*) *conformal predictive system* as follows:

1. Divide the training sequence Z_{tr} into two disjoint subsets; the proper training set Z_t and the calibration set $Z_c = \{(x_1, y_1), \dots, (x_q, y_q)\}$
2. Train the underlying model h using Z_t
3. Let $\hat{y}_i = h(x_i)$, for $i \in 1, \dots, q$
4. Obtain quality estimates for all the calibration examples $\hat{\sigma}_i$, for $i \in 1, \dots, q$
5. Make a point prediction for the test object $\hat{y} = h(x)$ and estimate its quality $\hat{\sigma}$
6. Produce a list of calibration scores using

$$C_i = \hat{y} + \frac{\hat{\sigma}}{\hat{\sigma}_i} (y_i - \hat{y}_i)$$

for $i \in 1, \dots, q$

7. Sort C_1, \dots, C_q in increasing order, resulting in $C_{(1)}, \dots, C_{(q)}$
8. Set $C_{(0)} = -\infty$ and $C_{(q+1)} = \infty$
9. Let $\tau \in U(0, 1)$
10. Return the predictive distribution:

$$Q(y) = \begin{cases} \frac{n+\tau}{q+1} & \text{if } y \in (C_{(n)}, C_{(n+1)}) \text{ for } n \in \{0, \dots, q\} \\ \frac{n'-1+(n''-n'+2)\tau}{q+1} & \text{if } y = C_{(n)} \text{ for } n \in \{1, \dots, q\} \end{cases}$$

where $n' = \min\{m | C_{(m)} = C_{(n)}\}$ and $n'' = \max\{m | C_{(m)} = C_{(n)}\}$

Note that for a sequence of test examples, only step 5 – 10 above need to be repeated. We adopt a similar terminology as for conformal regressors, and refer to conformal predictive systems as *standard conformal predictive systems*, in case $\hat{\sigma}_i = \hat{\sigma} = c$, for some constant $c > 0$, i.e., the quality estimate is independent of the predicted object. We refer to conformal predictive systems that are not standard as *normalized conformal predictive systems*. Approaches to estimating the quality include training a separate model to predict the size of the error, e.g., using kNN or ANN as in (Johansson et al., 2014a); others exploit properties of the underlying model, e.g., using disagreement (variance) of the trees in a random forest (Boström et al., 2017).

Prediction intervals for a chosen significance level ϵ can be defined by

$$[C_{Q''_{.\epsilon/2}}, C_{Q'_{.100-\epsilon/2}}] \tag{1}$$

where $Q''_p = \max\{m | Q_{(m)} < p/100\}$ and $Q'_p = \min\{m | Q_{(m)} > p/100\}$. A point prediction can similarly be defined by $C_{Q'_{.50}}$.

The above way of forming prediction intervals contrasts to (standard or normalized) conformal regressors by not enforcing that the underlying prediction is placed in the middle of the interval. In fact, when using conformal predictive systems to generate prediction intervals, it may very well be the case that the underlying prediction falls outside the prediction interval.

3. Mondrian Conformal Predictive Distributions

In Mondrian conformal predictors (Vovk et al., 2005), the available calibration examples are somehow divided into different categories, and then a valid conformal predictor is built for each category. The most common Mondrian conformal predictor is probably the *class-conditional conformal predictor* (Shi et al., 2013), where the categories represent the possible class labels, thus providing guarantees for each label, i.e., the errors will be evenly distributed over the classes. The problem space can also be divided w.r.t. to the feature space, e.g., for tree models, a very natural division is to regard each leaf (path) as a separate category, resulting in that each such leaf is independently valid, see e.g., (Johansson et al., 2014b). Recently, Mondrian conformal regressors were proposed in (Boström and Johansson, 2020). Until now, however, Mondrian conformal prediction has, to the best of our knowledge, not been applied to conformal predictive distributions.

A (*split*) *Mondrian conformal predictive system* produces a Mondrian conformal predictive distribution in the following way, given a training sequence Z_{tr} , and a test object x :

1. Divide the training sequence Z_{tr} into two disjoint subsets; the proper training set Z_t and the calibration set $Z_c = (x_1, y_1), \dots, (x_q, y_q)$
2. Train the underlying model h using Z_t
3. Divide Z_c into k disjoint subsets Z_{c1}, \dots, Z_{ck} , according to a Mondrian taxonomy κ with categories $\kappa_1, \dots, \kappa_k$
4. For each category κ_j , let $\hat{y}_{ji} = h(x_{ji})$ and let $\hat{\sigma}_{ji}$ be the corresponding quality estimate, for each $x_{ji} \in Z_{cj}$
5. Make a point prediction for the test object $\hat{y} = h(x)$ and estimate its quality $\hat{\sigma}$
6. Identify which category κ_j the test object belongs to and produce a list of calibration scores using

$$C_{ji} = \hat{y} + \frac{\hat{\sigma}}{\hat{\sigma}_{ji}} (y_{ji} - \hat{y}_{ji})$$

for $i \in 1, \dots, q$, where $q = |Z_{cj}|$

7. Sort C_{j1}, \dots, C_{jq} in increasing order, resulting in $C_{j(1)}, \dots, C_{j(q)}$
8. Set $C_{j(0)} = -\infty$ and $C_{j(q+1)} = \infty$.
9. Let $\tau \in U(0, 1)$

10. Return the predictive distribution:

$$Q(y) = \begin{cases} \frac{n+\tau}{q+1} & \text{if } y \in (C_{j(n)}, C_{j(n+1)}) \text{ for } n \in \{0, \dots, q\} \\ \frac{n'-1+(n''-n'+2)\tau}{q+1} & \text{if } y = C_{j(n)} \text{ for } n \in \{1, \dots, q\} \end{cases}$$

where $n' = \min\{m | C_{j(m)} = C_{j(n)}\}$ and $n'' = \max\{m | C_{j(m)} = C_{j(n)}\}$

For Mondrian conformal classification, the possible class labels are often used to define the categories, while for Mondrian conformal regression (Boström and Johansson, 2020), the categories are defined using the quality (difficulty) estimate σ . We here propose to form the categories of the Mondrian taxonomy through binning of the predictions, using equal-sized bins, similar to what has been proposed for classification problems in the context of Venn prediction (Zhou et al., 2014). To handle the special case when the number of identical predictions is larger than the bin size, we assume that a very small random number ξ is added to each prediction, which allows for that approximately the same number of examples will fall into each bin (category). The number of bins (categories) is hence a parameter of the approach, and it should be chosen in a way such that the size of each partition of the calibration set is sufficiently large, e.g., to allow for prediction intervals with specified level of confidence to be extracted.

4. Experiments

First, we illustrate the proposed approach using the bank8fm dataset. After that, we present results from using different strategies to forming predictive systems on real-world datasets.

4.1. Illustration using the bank8fm dataset

In this demonstration, half of the data is used as a proper training set and the other half for calibration. The underlying model is a regression forest with 500 trees. Fig. 1a plots the predictions vs. the residuals. Interestingly enough, there is a clear trend showing that the residuals are larger for the higher predictions. With this in mind, we divide the predictions into five bins, with an equal number of examples in each bin, and then form a standard (i.e., non-normalized) conformal predictive distribution (CPD), given the mean prediction of each bin. Fig. 1b – Fig. 1f, show the CPD for each bin, specifically comparing the mean prediction of each bin (indicated by an orange solid line) to the median of the CPD (indicated by a green dashed line).

If we would consider using the medians of the CPDs as point predictions, instead of the original mean predictions, we can see that they would be substantially modified. For the first three bins, the predictions would be lowered, but for the last two, they would instead be increased. In particular, in the first bin, the prediction from the underlying model is actually outside of the 95%-confidence interval (indicated by the yellow dashed lines), see Fig. 1b, showing the potential of a CPD to improve predictions by calibration.

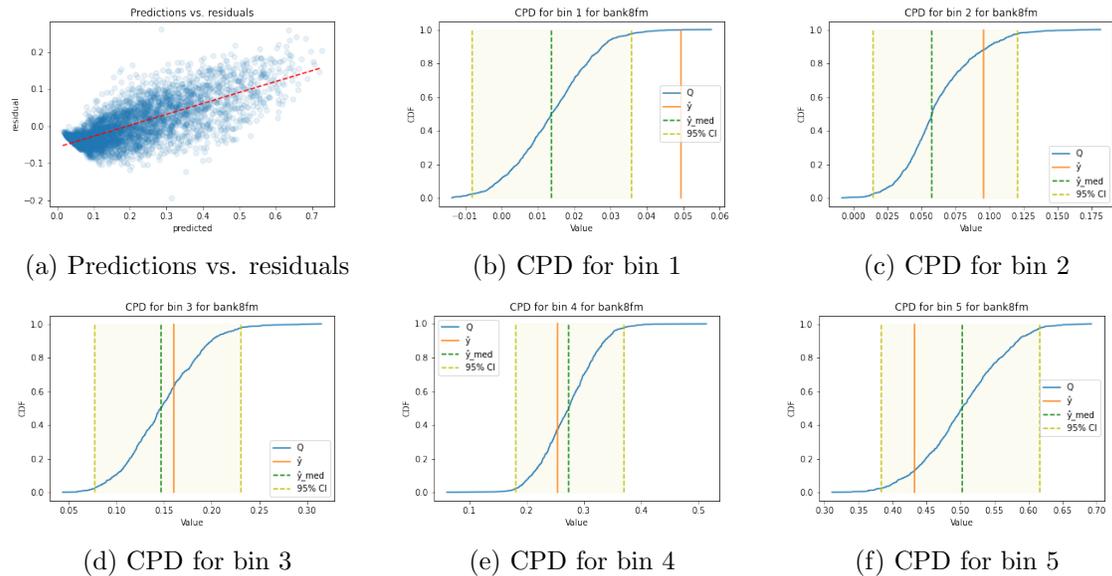


Figure 1: Illustration using the bank8fm dataset

4.2. Real-world datasets

4.2.1. EXPERIMENTAL SETUP

In the experimentation, we utilize random forests with 500 binary regression trees as the underlying models. Here, 1/3 of the available features were randomly selected for consideration in each split optimisation during the generation. Following the findings in (Werner et al., 2020), we have opted for using out-of-bag-calibration, thus making all training examples available for both training the underlying model and obtaining calibration scores. As a quality estimate, the variance of the predictions from the individual trees in the forest was used in all normalised setups, as originally proposed for conformal regressors in (Boström et al., 2017) and also investigated for conformal predictive systems in (Werner et al., 2020).

For the evaluation, we compare the suggested Mondrian approaches to both conformal regression setups and standard predictive conformal distributions. All-in-all, six different setups are evaluated:

- CR: Standard conformal regressor as described in, for instance, (Boström et al., 2017), i.e., using the absolute error as the non-conformity measure, and utilizing out-of-bag estimates for calibration.
- CRn: Normalized conformal regressor using the variance of the predictions from the individual trees as the difficulty function, as described in (Boström et al., 2017).
- CPS: Standard conformal predictive system as described in Section 2.
- CPSn: Normalized conformal predictive system as described in Section 2.
- MCPS: Mondrian conformal predictive system using five bins as described in Section 3.

Table 1: Real-world datasets

Dataset	#Examples	#Features	Dataset	#Examples	#Features
abalone	4177	8	kin8nh	8192	8
anacalt	4052	7	kin8nm	8192	8
bank8fh	8192	8	laser	993	4
bank8fm	8192	8	mg	1385	6
bank8nh	8192	8	mortage	1048	15
bank8nm	8192	8	plastic	1649	2
boston	506	13	puma8fh	8192	8
comp	8192	12	puma8fm	8192	8
concrete	1030	8	puma8nh	8192	8
cooling	768	8	puma8nm	8192	8
deltaA	7129	5	quakes	2178	3
deltaE	9517	6	stock	950	9
friedm	1200	5	treasury	1048	15
heating	768	8	wineRed	1599	11
istanbul	536	7	wineWhite	4898	11
kin8fh	8192	8	wizmir	1461	9
kin8fm	8192	8			

- MCPSn: Normalized Mondrian conformal predictive system using five bins as described in Section 3.

All-in-all 33 publicly available data sets, previously used in e.g., (Johansson et al., 2014a; Boström et al., 2017; Boström and Johansson, 2020; Werner et al., 2020), were employed; for characteristics of these data sets, see Table 1. For the actual evaluation, standard 10-fold cross-validation was used.

4.2.2. RESULTS

We first look at the the quality of the predictions from the four variants of conformal predictive systems using the loss function *continuous ranked probability score* (CRPS) (Vovk et al., 2020) in Table 2. Here, the two Mondrian approaches clearly outperform both standard and normalized conformal predictive systems. This is confirmed by a Friedman test (Friedman, 1937), followed by a Nemenyi post-hoc test (Nemenyi, 1963), showing the differences to be significant at $\alpha=0.05$, see Fig. 2. This is of course a very strong result for the suggested Mondrian conformal predictive system and consequently a key finding of this study.

Turning to the efficiency results in Table. 3, we see from the mean interval sizes that the normalized versions produce the most informative predictions. Specifically, the Mondrian normalized version obtained the tightest intervals, on average. Looking at the results of the statistical testing in Fig. 3, the only significant difference is, however, that the standard conformal regressor and standard conformal predictive system are significantly worse than the other four alternatives.

Considering median interval sizes instead, the above results are confirmed; see Table. 4 and Fig. 4.

Table 2: CRPS

	CPS	CPSn	MCPS	MCPSn		CPS	CPSn	MCPS	MCPSn
abalone	.040	.039	.039	.039	kin8nh	.072	.071	.068	.068
anacalt	.007	.007	.007	.006	kin8nm	.060	.059	.051	.050
bank8fh	.053	.053	.048	.048	laser	.012	.011	.011	.010
bank8fm	.030	.030	.021	.020	mg	.039	.035	.038	.034
bank8nh	.056	.056	.053	.053	mortgage	.004	.004	.003	.003
bank8nm	.025	.025	.020	.019	plastic	.094	.096	.093	.096
boston	.036	.034	.033	.033	puma8fh	.082	.081	.078	.078
comp	.015	.015	.014	.014	puma8fm	.042	.042	.034	.034
concrete	.036	.035	.033	.032	puma8nh	.081	.080	.075	.074
cooling	.021	.020	.019	.018	puma8nm	.052	.052	.039	.038
deltaA	.020	.019	.019	.019	quakes	.095	.096	.091	.094
deltaE	.029	.029	.029	.028	stock	.013	.013	.013	.013
friedm	.042	.042	.031	.031	treasury	.004	.004	.004	.004
heating	.009	.009	.009	.009	wineRed	.060	.058	.054	.053
istanbul	.044	.044	.044	.044	wineWhite	.052	.050	.048	.046
kin8fh	.042	.042	.040	.039	wizmir	.011	.011	.011	.011
kin8fm	.025	.025	.018	.018	Mean	.039	.039	.036	.036
					Ranks	3.70	3.09	2.00	1.21

Table 3: Efficiency (mean)

	CR	CRn	CPS	CPSn	MCPS	MCPSn		CR	CRn	CPS	CPSn	MCPS	MCPSn
abalone	.320	.283	.325	.284	.291	.288	kin8nh	.492	.482	.491	.481	.477	.470
anacalt	.074	.050	.076	.052	.087	.066	kin8nm	.417	.395	.414	.394	.372	.347
bank8fh	.392	.341	.368	.353	.349	.352	laser	.090	.066	.092	.066	.117	.063
bank8fm	.228	.176	.210	.176	.144	.140	mg	.358	.210	.358	.211	.349	.221
bank8nh	.461	.413	.432	.414	.403	.412	mortgage	.037	.033	.036	.033	.030	.029
bank8nm	.238	.150	.234	.151	.146	.134	plastic	.657	.803	.679	.788	.678	.833
boston	.294	.267	.282	.265	.285	.294	puma8fh	.562	.522	.560	.521	.543	.534
comp	.114	.107	.116	.108	.108	.106	puma8fm	.281	.273	.280	.275	.245	.241
concrete	.274	.246	.274	.246	.281	.246	puma8nh	.554	.525	.552	.523	.540	.534
cooling	.187	.142	.190	.142	.140	.126	puma8nm	.330	.319	.331	.318	.293	.296
deltaA	.155	.145	.155	.146	.155	.145	quakes	.709	.848	.693	.820	.633	.757
deltaE	.215	.214	.215	.214	.217	.214	stock	.095	.089	.095	.090	.098	.090
friedm	.299	.315	.301	.316	.239	.247	treasury	.043	.038	.043	.038	.040	.036
heating	.074	.068	.070	.066	.073	.064	wineRed	.500	.450	.498	.456	.495	.464
istanbul	.321	.333	.314	.334	.369	.379	wineWhite	.417	.371	.418	.374	.400	.369
kin8fh	.297	.290	.303	.291	.280	.278	wizmir	.080	.078	.079	.077	.079	.079
kin8fm	.183	.177	.184	.177	.131	.126	Mean	.295	.279	.293	.279	.275	.272
							Ranks	4.97	2.58	4.82	3.00	3.33	2.30

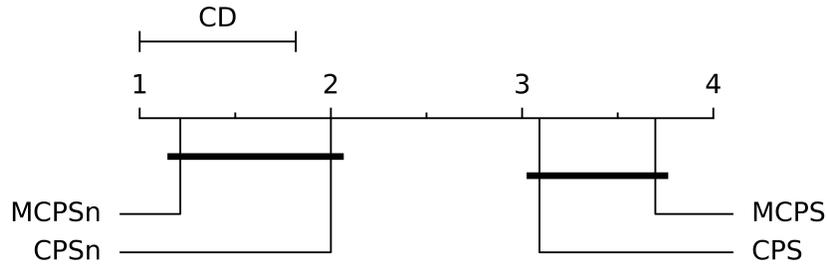


Figure 2: CRPS ranks

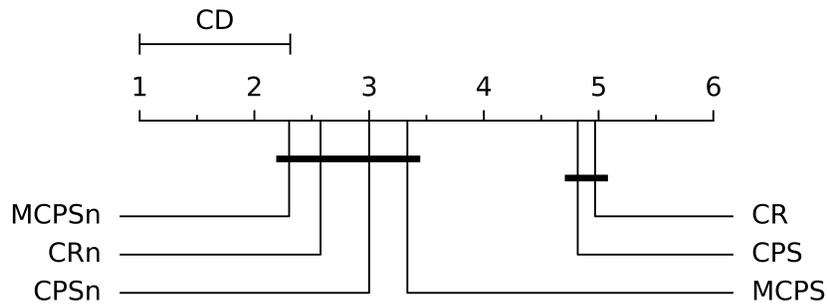


Figure 3: Efficiency ranks

Turning to the predictive performance, Table. 5 shows the mean absolute errors. Here, the two Mondrian approaches clearly outperform the other setups. In fact, statistical testing shows that the differences are significant at $\alpha = 0.05$, see Fig. 5. So, based on these results considering a large number of real-world data sets, the Mondrian approach leads to better predictors than the alternatives.

When considering mean squared error instead of mean absolute error, the differences between the setups are actually very small in absolute numbers. Still, when looking at the mean ranks over all data sets, the normalized Mondrian conformal predictive systems are the most accurate, see Table. 6. Here, however, as seen in Fig. 6, the statistical tests identify fewer significant differences.

Summarising the main experiment, we see that the novel Mondrian conformal predictive systems outperformed standard and normalized conformal predictive systems with regard to the continuous ranked probability score loss metric. In addition, the informativeness, as measured using the prediction interval sizes, was generally better for the normalized versions compared to standard conformal regressors and conformal predictive systems. Finally, also when evaluating the predictive performance in terms of mean absolute or mean-squared error, the Mondrian variants were the most accurate, thus demonstrating their inherent capability of improving the predictions of the underlying models.

Table 4: Efficiency (median)

	CR	CR _n	CPS	CPS _n	MCPS	MCPS _n		CR	CR _n	CPS	CPS _n	MCPS	MCPS _n
abalone	.320	.258	.325	.259	.273	.261	kin8nh	.492	.478	.491	.477	.489	.465
anacalt	.074	.044	.076	.046	.000	.000	kin8nm	.417	.385	.414	.384	.369	.336
bank8fh	.392	.322	.368	.334	.344	.345	laser	.090	.055	.092	.054	.117	.050
bank8fm	.228	.165	.210	.166	.155	.146	mg	.358	.162	.358	.162	.381	.168
bank8nh	.461	.381	.432	.384	.387	.389	mortgage	.037	.032	.036	.032	.023	.023
bank8nm	.238	.120	.234	.121	.095	.087	plastic	.657	.694	.679	.689	.651	.725
boston	.294	.206	.282	.208	.257	.242	puma8fh	.562	.515	.560	.515	.549	.527
comp	.114	.103	.116	.103	.105	.102	puma8fm	.281	.274	.280	.277	.271	.249
concrete	.274	.219	.274	.222	.244	.213	puma8nh	.554	.492	.552	.491	.569	.495
cooling	.187	.124	.190	.124	.134	.110	puma8nm	.330	.301	.331	.301	.318	.278
deltaA	.155	.138	.155	.140	.128	.125	quakes	.709	.758	.693	.733	.635	.705
deltaE	.215	.208	.215	.209	.216	.205	stock	.095	.084	.095	.085	.092	.086
friedm	.299	.309	.301	.311	.234	.240	treasury	.043	.037	.043	.037	.033	.030
heating	.074	.064	.070	.062	.076	.063	wineRed	.500	.432	.498	.437	.479	.443
istanbul	.321	.319	.314	.322	.371	.358	wineWhite	.417	.364	.418	.369	.382	.364
kin8fh	.297	.281	.303	.283	.272	.267	wizmir	.080	.075	.079	.075	.078	.078
kin8fm	.183	.172	.184	.173	.118	.117	Mean	.295	.260	.293	.260	.268	.251
							Ranks	5.06	2.58	4.97	2.91	3.39	2.09

Table 5: Mean absolute errors

	CR	CR _n	CPS	CPS _n	MCPS	MCPS _n		CR	CR _n	CPS	CPS _n	MCPS	MCPS _n
abalone	.055	.054	.053	.053	.053	.053	kin8nh	.102	.102	.102	.102	.097	.097
anacalt	.008	.008	.008	.008	.008	.008	kin8nm	.085	.085	.085	.085	.072	.071
bank8fh	.076	.076	.074	.076	.067	.067	laser	.014	.014	.014	.014	.014	.013
bank8fm	.043	.043	.042	.045	.029	.029	mg	.051	.051	.051	.051	.049	.048
bank8nh	.082	.082	.075	.078	.073	.072	mortgage	.005	.005	.005	.005	.005	.005
bank8nm	.034	.034	.031	.034	.028	.027	plastic	.132	.131	.129	.129	.128	.128
boston	.048	.048	.048	.048	.046	.045	puma8fh	.117	.117	.117	.117	.111	.111
comp	.020	.020	.020	.020	.020	.019	puma8fm	.060	.060	.060	.060	.048	.048
concrete	.049	.049	.049	.049	.045	.044	puma8nh	.117	.117	.117	.117	.105	.105
cooling	.026	.026	.027	.026	.026	.026	puma8nm	.077	.077	.078	.077	.054	.052
deltaA	.026	.026	.026	.026	.026	.026	quakes	.137	.138	.134	.132	.129	.131
deltaE	.040	.040	.040	.040	.040	.040	stock	.019	.019	.019	.019	.018	.018
friedm	.058	.059	.059	.059	.043	.043	treasury	.005	.005	.005	.005	.005	.005
heating	.013	.013	.013	.013	.012	.012	wineRed	.081	.081	.081	.081	.072	.071
istanbul	.060	.060	.061	.060	.061	.061	wineWhite	.069	.069	.069	.069	.064	.063
kin8fh	.059	.059	.059	.059	.056	.056	wizmir	.015	.015	.015	.015	.015	.015
kin8fm	.035	.035	.035	.035	.025	.025	Mean	.055	.055	.055	.055	.050	.049
							Ranks	4.64	4.70	3.88	4.15	2.15	1.48

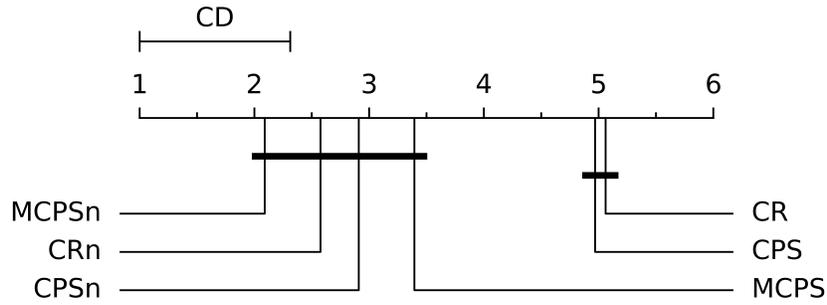


Figure 4: Median efficiency ranks

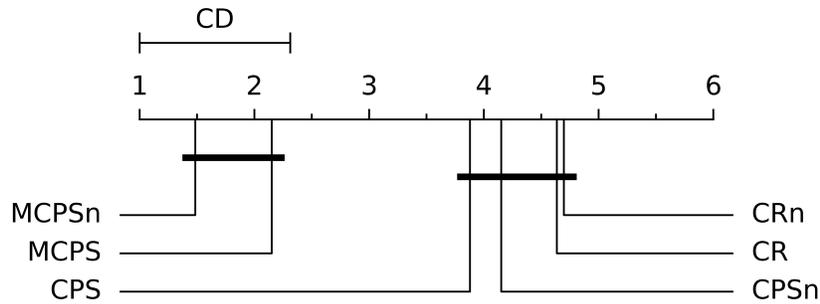


Figure 5: Mean absolute error ranks

5. Concluding remarks

We have introduced a new approach to forming conformal predictive distributions, called Mondrian conformal predictive systems, which generate one predictive distribution for each available Mondrian category. We have shown that by using the predictions of an underlying regression forest to form the categories through binning, the use of the resulting Mondrian predictive distributions will not only improve upon using standard and normalized conformal predictive distributions, but also result in as tight prediction intervals as produced by normalized conformal regressors, and even improve upon the point predictions of the underlying forest with respect to the mean absolute error.

Directions for future work include investigating more sophisticated approaches to forming the categories, e.g., using clustering as an alternative to binning as proposed in (Zhou et al., 2014). Such approaches may not only use the predictions of the underlying model but also the quality estimates. Currently, these estimates have only been considered for normalizing (rescaling) the predictive distribution of each category. The proposed framework also needs to be applied and evaluated in conjunction with other types of underlying model. Alternative ways of forming the categories may be required to allow for similar performance improvements as for the considered regression forests. Another important direction for future research includes investigating the use of the distributions for decision making, possibly

Table 6: Mean squared errors

	CR	CRn	CPS	CPSn	MCPS	MCPSn		CR	CRn	CPS	CPSn	MCPS	MCPSn
abalone	.006	.006	.006	.006	.006	.006	kin8nh	.016	.016	.016	.016	.015	.015
anacalt	.001	.001	.001	.001	.001	.001	kin8nm	.011	.011	.011	.011	.009	.009
bank8fh	.009	.009	.010	.010	.009	.009	laser	.001	.001	.001	.001	.001	.001
bank8fm	.003	.003	.003	.004	.002	.002	mg	.006	.006	.006	.006	.006	.006
bank8nh	.013	.013	.014	.015	.013	.013	mortgage	.000	.000	.000	.000	.000	.000
bank8nm	.003	.003	.003	.004	.002	.002	plastic	.029	.029	.031	.029	.031	.030
boston	.005	.005	.005	.006	.005	.005	puma8fh	.021	.021	.021	.021	.020	.021
comp	.001	.001	.001	.001	.001	.001	puma8fm	.005	.005	.005	.005	.004	.004
concrete	.005	.005	.005	.005	.004	.004	puma8nh	.021	.021	.021	.021	.019	.019
cooling	.002	.002	.002	.002	.002	.002	puma8nm	.008	.008	.008	.008	.005	.005
deltaA	.001	.001	.001	.001	.002	.002	quakes	.031	.031	.032	.032	.032	.032
deltaE	.003	.003	.003	.003	.003	.003	stock	.001	.001	.001	.001	.001	.001
friedm	.005	.006	.006	.006	.003	.003	treasury	.000	.000	.000	.000	.000	.000
heating	.000	.000	.000	.000	.000	.000	wineRed	.013	.013	.013	.013	.013	.013
istanbul	.006	.006	.006	.006	.007	.006	wineWhite	.010	.010	.010	.010	.010	.010
kin8fh	.006	.006	.006	.006	.005	.005	wizmir	.000	.000	.000	.000	.000	.000
kin8fm	.002	.002	.002	.002	.001	.001	Mean	.007	.007	.008	.008	.007	.007
							Ranks	3.27	3.21	4.58	4.45	3.00	2.48

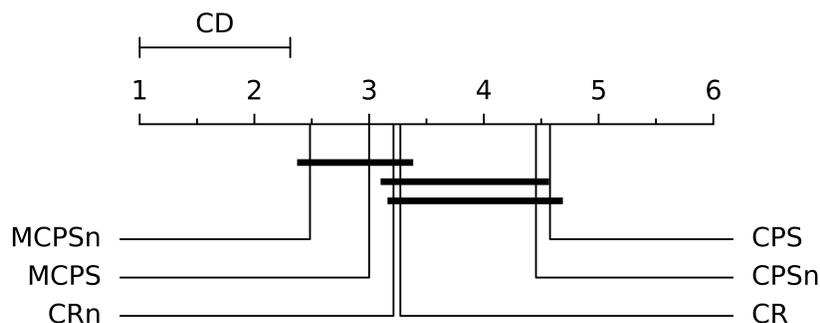


Figure 6: Mean squared error ranks

including performance metrics that are more directly connected to the utility, compared to the metrics employed in this study, e.g., CRPS.

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