# Scaling Integer Arithmetic in Probabilistic Programs (Supplementary Material) 

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## A EXPERIMENTAL DETAILS

All experiments were run on a server with a 2.20 GHz CPU and 504 GB RAM. WebPPL experiments were run on WebPPL v0.9.15; Psi experiments were run on Psi version ec2cfc14a62a168afe7ce1d7269b92cf2882b830. Dice.jl experiments were run using the code available at/https://github.com/Juice-jl/Dice.jl/tree/ arithmetic. The implementations of each model run are available at the same repository.

## B BENCHMARK MODELS

We provide a small description of our benchmarks in this section with details of their sources. The code implementations of all models are available at the repository https://github.com/Juice-jl/Dice.jl/tree/arithmetic.

1. book: A model of flipping towards a target page in a book adapted from the Psi test directory [Gehr et al., 2022].
2. tugofwar: Adapted from a traditional tug-of-war example [Huang et al. 2021], with values made discrete.
3. caesar: The caesar-cipher example from Dice [Holtzen et al. 2020]. with different number of characters being observed.
4. ranking: A model for learning a ranking system, adapted from Kisa et al. [2014].
5. radar1: A model of radar reception, adapted from a continuous model from Psi’s [Gehr et al. 2016] benchmark suite.
6. floydwarshall: An implementation of the Floyd-Warshall algorithm [Floyd, 1962] on a graph with edges of random weight.
7. linear-extensions: A model counting linear extensions [Dittmer and Pak, 2018] where we observe a partial order and get an output distribution over all matching total orders.
8. triangle: A model categorizing a triangle of random side lengths, adapted from the Psi test directory [Gehr et al., 2022].
9. gcd: A model checking if two random numbers are coprime implementing Euclid's algorithm [Lehmer, 1938].
10. disease: A discrete disease model taken from existing works [Laurel and Misailovic, 2020].
11. luhn: A probabilistic model of student IDs leveraging the Luhn algorithm [Luhn, 1960].

## C ADDITIONAL EXPERIMENTAL RESULTS

We provide additional experimental results supplementing those in the main paper.
BDD size serves as a proxy for how well knowledge compilation can exploit structure: the more compact the BDD, the smaller the representation of our function, and the faster weighted model counting can be executed. Figure 1 compares the resultant BDD size for various models when compiled in Dice.jl using a binary or one-hot encoding. We can see that in almost all cases the binary encoding results in a smaller BDD, in some cases much smaller. Note that these models are those for which the one-hot encoding did not timeout; it is likely for those models that timed out that the true compiled BDD size will end up being much larger than the binary encoded BDD size.

Table 1: BDD sizes for probabilistic models using a binary vs one-hot encoding

| Benchmarks | binary | one-hot |
| :--- | ---: | ---: |
| tugofwar | $\mathbf{2 4 0 0}$ | 2821 |
| caesar-small | $\mathbf{1 3 0 4}$ | 3879 |
| caesar-medium | $\mathbf{6 3 4 4}$ | 17879 |
| caesar-large | $\mathbf{1 2 6 4 4}$ | 35379 |
| ranking-small | $\mathbf{6 9 1}$ | 1146 |
| ranking-medium | $\mathbf{6 2 1 8}$ | 8297 |
| ranking-large | $\mathbf{1 1 4 9 1}$ | 15680 |
| radar1 | $\mathbf{1 8 1}$ | 332 |
| floydwarshall-small | $\mathbf{1 0}$ | $\mathbf{1 0}$ |
| floydwarshall-medium | 341 | $\mathbf{2 3 7}$ |
| linear extensions-small | $\mathbf{2 9}$ | 53 |
| linear extensions-medium | $\mathbf{1 3 3}$ | 257 |
| linear extensions-large | $\mathbf{9 9 7}$ | 1538 |
| triangle-small | $\mathbf{1 0 2 7 3}$ | 150089 |
| triangle-medium | $\mathbf{4 0 4 1 9}$ | 1156785 |
| luhn-small | $\mathbf{5 1 8}$ | 1010 |
| luhn-medium | $\mathbf{2 8 9 9}$ | 7361 |

## D BDD SIZES OF INTEGER DISTRIBUTION ENCODINGS (PROPOSITION 1)

## D. 1 NOTATION AND DEFINITIONS

A $b$-rooted BDD $B$ with $m$ decision variables computes some function $\{0,1\}^{m} \rightarrow\{0,1\}^{b}$ at its roots.
Let the values of the roots of $B$, given truth assignments to decision variables $\vec{x}$, be denoted $B_{1}(\vec{x}), \ldots, B_{b}(\vec{x})$. Let the variable order of $B$ match the order of the components of $\vec{x}$. A BDD can thus be specified by a tuple of functions $\left(\{0,1\}^{m} \rightarrow\{0,1\}\right)^{b}$. Let $B(\vec{x})$ denote this tuple, i.e. $\left(B_{1}(\vec{x}), \ldots, B_{b}(\vec{x})\right)$.

As $b$-length bitvectors are isomorphic to $b$-bit unsigned machine integers, we can also specify a BDD by a function from assignments to $b$-bit unsigned integers. Given $f:\{0,1\}^{m} \rightarrow\left\{0, \ldots, 2^{b}-1\right\}$, let $\llbracket f \rrbracket$ denote the BDD defined such that $\llbracket f \rrbracket_{i}(\vec{x})=\operatorname{LSB}_{i}(f(\vec{x}))$, where $\operatorname{LSB}_{i}(j)$ denotes the $i$ th least significant bit of $j$.

Let $\vec{p} \in[0,1]^{m}$ assign a probability to each decision variable. For each $p_{i}$, let $X_{i}$ denote an independent $\operatorname{Bernoulli}\left(p_{i}\right)$. We can construct a random variable for the result of $B$ with $B(\vec{X})$ (passing in $X_{i}$ for each $x_{i}$ ). Note the indexing of the decision variable $x_{i}$, the corresponding probability $p_{i}$, and the corresponding random variable $X_{i}$ may vary to make notation clearer: for CATEG_INT, these will be zero-indexed; for BITWISE_INT, these will be indexed by bit strings.

Let $\mathcal{D}_{b}$ denote a distribution over $b$-length bit vectors. Let $v^{\star}$ be an arbitrary random variable sampled from $\mathcal{D}_{b}$, interpreted as an unsigned integer. Denote the bits of $v^{\star}$, from least to most significant, as $v_{1}^{\star}, \ldots, v_{b}^{\star}$.

Definition 1 (Integer distribution encoding). A BDD and probability assignments $(B, \vec{p})$ encode a distribution $\mathcal{D}_{b}$ if $B(\vec{X}) \sim \mathcal{D}_{b}$.

## D. 2 CATEGORICAL ENCODING OF INTEGER DISTRIBUTIONS

Definition 2. Define encode $\operatorname{CATEG}\left(\mathcal{D}_{b}\right)=(\llbracket g \rrbracket, \vec{p})$ such that $p_{i}=\operatorname{Pr}\left(v^{\star}=i \mid v^{\star} \geq i\right), 0 \leq i \leq 2^{b}-2$, and that $g$ is defined as follows.

$$
g(\vec{x})= \begin{cases}0 & x_{0} \\ 1 & \neg x_{0} \wedge x_{1} \\ 2 & \neg x_{0} \wedge \neg x_{1} \wedge x_{2} \\ \cdots & \\ j & x_{j} \wedge \bigwedge_{i=0}^{j-1} \neg x_{i} \text { and } 0 \leq j \leq 2^{b}-2 \\ \cdots & \\ 2^{b}-1 & \bigwedge_{i=0}^{2^{b}-2} \neg x_{i}\end{cases}
$$

We could also write this as the following, which is also how it would likely be implemented in a probabilistic programming language.

$$
g(\vec{x})= \begin{cases}0 & \text { if } x_{0} \\ 1 & \text { else if } x_{1} \\ 2 & \text { else if } x_{2} \\ \cdots & \\ 2^{b}-2 & \text { else if } x_{2^{b}-2} \\ 2^{b}-1 & \text { else }\end{cases}
$$

Lemma 1. If $\operatorname{Img}(v) \subseteq \mathbb{N}_{0}$, then $\prod_{i=0}^{x-1} \operatorname{Pr}(v \neq i \mid v \geq i)=\operatorname{Pr}(v \geq x)$. Proof omitted; proceed by induction.

Lemma 2. For any distribution $\mathcal{D}_{b}$ over unsigned integers, encode CATEG $^{\left(\mathcal{D}_{b}\right) \text { encodes } \mathcal{D}_{b} .}$

Proof. Let encode ${ }_{\text {CATEG }}\left(\mathcal{D}_{b}\right)=(B, \vec{p})$. We prove that for all $0 \leq j \leq 2^{b}-1, \operatorname{Pr}(g(\vec{X})=j)=\operatorname{Pr}\left(v^{\star}=j\right)$.
Case 1: $0 \leq j \leq 2^{b}-2$.

$$
\begin{array}{rlrl}
\operatorname{Pr}(g(\vec{X})=j) & =\operatorname{Pr}\left(X_{j} \wedge \bigwedge_{i=0}^{j-1} \neg X_{i}\right) & & \\
& =\operatorname{Pr}\left(X_{j}\right) \prod_{i=0}^{j-1} \operatorname{Pr}\left(\neg X_{i}\right) & & \text { (Independence) } \\
& =\operatorname{Pr}\left(v^{\star}=j \mid v^{\star} \geq j\right) \prod_{i=0}^{j-1} \operatorname{Pr}\left(v^{\star} \neq i \mid v^{\star} \geq i\right) & & \text { (As } \left.\operatorname{Pr}\left(X_{i}\right)=p_{i}\right) \\
& =\operatorname{Pr}\left(v^{\star}=j \mid v^{\star} \geq j\right) \operatorname{Pr}\left(v^{\star} \geq j\right) & & \text { (Lemma@1) } \\
& =\operatorname{Pr}\left(v^{\star}=j\right) &
\end{array}
$$

Case 2: $j=2^{b}-1$.

$$
\begin{aligned}
\operatorname{Pr}\left(g(\vec{X})=2^{b}-1\right) & =\operatorname{Pr}\left(\bigwedge_{i=0}^{2^{b}-2} \neg X_{i}\right) & & \\
& =\prod_{i=0}^{2^{b}-2} \operatorname{Pr}\left(\neg X_{i}\right) & & \text { (Independence) } \\
& =\prod_{i=0}^{2^{b}-2} \operatorname{Pr}\left(v^{\star} \neq i \mid v^{\star} \geq i\right) & & \text { (As } \left.\operatorname{Pr}\left(X_{i}\right)=p_{i}\right) \\
& =\operatorname{Pr}\left(v^{\star} \geq 2^{b}-1\right) & & \text { (Lemma【1 } \\
& =\operatorname{Pr}\left(v^{\star}=2^{b}-1\right) & &
\end{aligned}
$$

## D. 3 BDD SIZE OF THE CATEG_INT ENCODING

We prove the first half of Proposition 1.
Proposition 1 (first half). A discrete distribution over the integers $\left\{0,1 \ldots, 2^{b}-1\right\}$ compiles to a BDD of size $\Theta\left(b 2^{b}\right)$ when represented using CATEG_INT (Algorithm 1).

Lemma 3. Let encode $\operatorname{CATEG}\left(\mathcal{D}_{b}\right)=(\llbracket g \rrbracket, P) . \llbracket g \rrbracket$ exactly matches a $B D D C$ with the following reduced structure.
For $1 \leq i \leq b, C$ has nodes $N_{i, 0}, N_{i, 1}, \ldots, N_{i, 2^{b}-2^{i-1}-1}$, at the levels of decision variables $x_{0}, x_{1}, \ldots, x_{2^{b}-2^{i-1}-1}$, respectively.

$$
\begin{array}{rlrlrl}
\operatorname{low}\left(N_{i, j}\right) & =\text { if } \quad j<2^{b}-2^{i-1}-1 \quad \text { then } \quad N_{i, j+1} \quad \text { else } & 1 \\
\operatorname{high}\left(N_{i, j}\right) & =\operatorname{LSB}_{i}(j) & &
\end{array}
$$

Let the roots of $C$ be placed such that $C_{i}(\vec{x})$ matches $N_{i, 0}$ given assignments to decision variables $\vec{x}$.

As an example, the BDD for $b=2$ follows (terminal nodes visually duplicated for clarity).


Proof. We will show for all $\vec{x}, C_{i}(\vec{x})=\operatorname{LSB}_{i}(g(\vec{x}))$.
Let $S_{i, j}(\vec{x})$ denote the state of $C_{i}$ at the level of $x_{j}$, given an assignment to $\vec{x}$. This is either a node at level $x_{j}$, or possibly one at a lower level if there was a low/high edge that skipped levels. See the following BDD and corresponding states as an example.

$$
\begin{array}{ll}
S_{1,0}\left(x_{0}, x_{1}, x_{2}\right) & =N_{1,1} \\
S_{1,1}\left(0, x_{1}, x_{2}\right) & =N_{1,2} \\
S_{1,1}\left(1, x_{1}, x_{2}\right) & =0 \\
S_{1,2}\left(0,0, x_{2}\right) & =N_{1,2} \\
S_{1,2}\left(0,1, x_{2}\right) & =1 \\
S_{1,2}\left(1, x_{1}, x_{2}\right) & =0
\end{array}
$$



Let $P(t)$ be the statement, $S_{i, t}=$ if $\quad g(\vec{x})<t \vee t>2^{b}-2^{i-1}-1 \quad$ then $\quad \operatorname{LSB}_{i}(g(\vec{x})) \quad$ else $\quad N_{i, t}$. $P(1)$ holds as the initial node of all $C_{i}(\vec{x})$ is always $N_{i, 0}$.
Assume $P(t)$ for an arbitrary $t \leq 2^{b}-1$.
Consider going to the next node. If we are at a terminal node, then the state stays the same, otherwise we go to the low or high edge based on $x_{t}$. Thus, to advance state, we replace $N_{i, t}$ with if $\quad x_{t}$ then $\operatorname{high}\left(N_{i, t}\right) \quad$ else $\quad \operatorname{low}\left(N_{i, t}\right)$.

$$
\begin{aligned}
S_{i, t+1}= & \text { if } g(\vec{x})<t \vee t>2^{b}-2^{i-1}-1 \\
& \text { then } \operatorname{LSB}_{i}(g(\vec{x})) \\
& \text { else } \quad\left(\text { if } \quad x_{t} \quad \text { then } \quad \operatorname{high}\left(N_{i, t}\right) \quad \text { else } \quad \operatorname{low}\left(N_{i, t}\right)\right)
\end{aligned}
$$

We substitute for $\operatorname{high}\left(N_{i, t}\right)$ and $\operatorname{low}\left(N_{i, t}\right)$.

$$
\begin{aligned}
S_{i, t+1}= & \text { if } g(\vec{x})<t \vee t>2^{b}-2^{i-1}-1 \\
& \text { then } \operatorname{LSB}_{i}(g(\vec{x})) \\
& \text { else } \quad\left(\text { if } \quad x_{t} \quad \text { then } \quad \operatorname{LSB}_{i}(t) \quad \text { else } \quad\left(\text { if } \quad t<2^{b}-2^{i-1}-1 \quad \text { then } \quad N_{i, t+1} \text { else } 1\right)\right)
\end{aligned}
$$ $x_{t} \wedge \neg(g(\vec{x})<t)$ implies $g(\vec{x})=t$.

$$
S_{i, t+1}=\text { if } \quad g(\vec{x})<t \vee t>2^{b}-2^{i-1}-1
$$

then $\operatorname{LSB}_{i}(g(\vec{x}))$
else (if $x_{t}$ then $\operatorname{LSB}_{i}(g(\vec{x}))$ else (if $t<2^{b}-2^{i-1}-1$ then $N_{i, t+1}$ else 1$)$ )
Two branches are now equivalent $\left(\operatorname{LSB}_{i}(g(\vec{x}))\right)$ and can be combined.

$$
\begin{aligned}
S_{i, t+1}= & \text { if } \quad g(\vec{x})<t \vee t>2^{b}-2^{i-1}-1 \vee x_{t} \\
& \text { then } \operatorname{LSB}_{i}(g(\vec{x})) \\
& \text { else } \quad\left(\text { if } \quad t<2^{b}-2^{i-1}-1 \quad \text { then } \quad N_{i, t+1} \quad \text { else } \quad 1\right)
\end{aligned}
$$

$g(\vec{x})<t \vee x_{t}$ is equivalent to $g(\vec{x})<t+1$.

$$
\begin{aligned}
S_{i, t+1}= & \text { if } g(\vec{x})<t+1 \vee t>2^{b}-2^{i-1}-1 \\
& \text { then } \operatorname{LSB}_{i}(g(\vec{x})) \\
& \text { else } \quad\left(\text { if } \quad t<2^{b}-2^{i-1}-1 \quad \text { then } \quad N_{i, t+1} \quad \text { else } 1\right)
\end{aligned}
$$

In the outer else, $t \leq 2^{b}-2^{i-1}-1$ (by the condition). In the inner else, $t$ is also $\geq 2^{b}-2^{i-1}-1$.

$$
\left.\begin{array}{rl}
S_{i, t+1}= & \text { if } g(\vec{x})<t+1 \vee t>2^{b}-2^{i-1}-1 \\
& \text { then } \operatorname{LSB}_{i}(g(\vec{x})) \\
& \text { else } \quad\left(\text { if } \quad t=2^{b}-2^{i-1}-1 \quad \text { then } \quad N_{i, t+1}\right.
\end{array} \text { else } \quad 1\right) ~ \$ ~ \$
$$

In the inner else, $g(\vec{x})>t$ by the outer condition and $t=2^{b}-2^{i-1}-1$ by the inner condition; together, $g(\vec{x})>2^{b}-2^{i-1}-1$. For all such $g(\vec{x}), \operatorname{LSB}_{i}(g(\vec{x}))=1$ (as $\left.2^{b}-1-2^{i-1}\right)$ is the largest number at most $2^{b}-1$ with the $i^{\text {th }}$ bit set to 0 ). Thus, we can merge two more branches.

$$
\begin{aligned}
S_{i, t+1}= & \text { if } \quad g(\vec{x})<t+1 \vee t+1>2^{b}-2^{i-1}-1 \\
& \text { then } \operatorname{LSB}_{i}(g(\vec{x})) \\
& \text { else } \quad N_{i, t+1}
\end{aligned}
$$

Thus $P(t) \rightarrow P(t+1)$ and thus $P(t)$ holds for all $1 \leq t \leq 2^{b}-1$.
Consider $P\left(2^{b}-1\right)$.

$$
\begin{aligned}
& S_{i, 2^{b}-1}=\text { if } \quad g(\vec{x})<2^{b}-1 \vee 2^{b}-1>2^{b}-2^{i-1}-1 \quad \text { then } \quad \operatorname{LSB}_{i}(g(\vec{x})) \quad \text { else } \quad N_{i, 2^{b}} \\
& S_{i, 2^{b}-1}=\operatorname{LSB}_{i}(g(\vec{x}))
\end{aligned}
$$

Therefore $\llbracket g \rrbracket$ and $C$ are equivalent.

## Proposition 1 (first half).

Proof. Let encode $\operatorname{CATEG}\left(\mathcal{D}_{b}\right)=(B, \vec{p}) . B$ has $b 2^{b}-2^{b}+1$ decision nodes.
It follows directly from Lemma 3 that $\sum_{i=1}^{n} 2^{b}-2^{i-1}=b 2^{b}-2^{b}+1$ nodes are needed for $B$ as well.

## D. 4 BITWISE ENCODING OF INTEGER DISTRIBUTIONS

Let $s_{i}$ denote the $i^{\text {th }}$ bit of a 1-indexed bit string; for example, $0010_{3}=1$. Let $|s|$ denote the length of a bit string. Let $\frown$ denote concatenation for bits and bit strings.

Definition 3. Define encode BITwISE_INT $\left(\mathcal{D}_{b}\right)=(B, \vec{p})$ such that:
For each bit string $s, 0 \leq|s|<b$, there is a decision variable $x_{s}$ with truth probability $p_{s}=\operatorname{Pr}\left(v_{|s|+1}^{\star} \mid \bigwedge_{i=1}^{|s|} v_{i}^{\star}=s_{i}\right)$. Intuitively, the bits are chosen left-to-right, and each decision variable chooses the next bit given a bit string of past choices. Consider the following examples, where the empty bit string is denoted as $\varepsilon$.

$$
\begin{aligned}
\operatorname{Pr}\left(x_{\varepsilon}\right) & =\operatorname{Pr}\left(v_{1}^{\star}\right) \\
\operatorname{Pr}\left(x_{0}\right) & =\operatorname{Pr}\left(v_{2}^{\star} \mid \neg v_{1}^{\star}\right) \\
\operatorname{Pr}\left(x_{1}\right) & =\operatorname{Pr}\left(v_{2}^{\star} \mid v_{1}^{\star}\right) \\
\operatorname{Pr}\left(x_{0100}\right) & =\operatorname{Pr}\left(v_{5}^{\star} \mid \neg v_{1}^{\star} \wedge v_{2}^{\star} \wedge \neg v_{3}^{\star} \wedge \neg v_{4}^{\star}\right)
\end{aligned}
$$

We specify $B_{i}(\vec{x})$ for $1 \leq i \leq b$ :

$$
B_{i}(\vec{x})=x_{B_{1}(\vec{x}) \frown B_{2}(\vec{x}) \frown \ldots \frown B_{i-1}(\vec{x})}
$$

For example, $B_{1}(\vec{x})=x_{\varepsilon}, B_{2}(\vec{x})=x_{B_{1}(\vec{x})}, B_{3}(\vec{x})=x_{B_{1}(\vec{x})-B_{2}(\vec{x})}$, and so on. Note that the specification above is equivalent to the following, which is a more natural representation in probabilistic programming, and an "unrolled" version of Algorithm 2.

$$
\begin{aligned}
& B_{1}(\vec{x})=x_{\varepsilon} \\
& B_{2}(\vec{x})=\text { if } x_{\varepsilon} \text { then } x_{1} \text { else } x_{0} \\
& B_{3}(\vec{x})\left.\left.=\text { if } x_{\varepsilon} \text { then (if } x_{1} \text { then } x_{11} \text { else } x_{10}\right) \text { else (if } x_{0} \text { then } x_{01} \text { else } x_{00}\right) \\
& \vdots \\
& B_{n}(\vec{x})=\ldots
\end{aligned}
$$

Lemma 4. For any distribution $\mathcal{D}_{b}$ over unsigned integers, encode $\operatorname{BITWISE}^{\text {INT }}\left(\mathcal{D}_{b}\right)$ encodes $\mathcal{D}_{b}$.

Proof. We prove that for any possible assignment to bits $\vec{a} \in\{0,1\}^{b}, \operatorname{Pr}(B(\vec{X})=\vec{a})=\operatorname{Pr}\left(\overrightarrow{v^{\star}}=\vec{a}\right)$.

$$
\begin{aligned}
& \left.\operatorname{Pr}(B(\vec{X})=\vec{a})=\prod_{i=1}^{b} \operatorname{Pr}\left(B_{i}(\vec{X})\right)=a_{i} \mid \bigwedge_{j=1}^{i-1} B_{j}(\vec{X})=a_{j}\right) \quad \quad \text { (Chain rule of probability) } \\
& =\prod_{i=1}^{b} \operatorname{Pr}\left(X_{B_{1}(\vec{X}) \frown \ldots \frown B_{i-1}(\vec{X})}=a_{i} \mid \bigwedge_{j=1}^{i-1} B_{j}(\vec{X})=a_{j}\right) \quad \text { (Bit specification) } \\
& =\prod_{i=1}^{b} \operatorname{Pr}\left(X_{a_{1} \ldots \frown a_{i-1}}=a_{i} \mid \bigwedge_{j=1}^{i-1} B_{j}(\vec{X})=a_{j}\right) \quad \text { (Condition) } \\
& =\prod_{i=1}^{b} \operatorname{Pr}\left(X_{a_{1} \ldots \frown a_{i-1}}=a_{i} \mid \bigwedge_{j=1}^{i-1} X_{B(\vec{X})_{1} \frown \cap_{j-1}(\vec{X})}=a_{j}\right) \quad \text { (Bit specification) } \\
& =\prod_{i=1}^{b} \operatorname{Pr}\left(X_{a_{1} \ldots} \ldots a_{i-1}=a_{i}\right) \\
& =\prod_{i=1}^{b} \operatorname{Pr}\left(b_{i}^{\star}=a_{i} \mid \bigwedge_{j=1}^{i-1} b_{j}^{\star}=a_{j}\right) \quad\left(\text { As } \operatorname{Pr}\left(X_{s}\right)=p_{s}\right) \\
& =\operatorname{Pr}\left(\overrightarrow{b^{\star}}=\vec{a}\right) \quad \text { (Chain rule of probability) }
\end{aligned}
$$

## D. 5 BDD SIZE OF THE BITWISE_INT ENCODING

We prove the second half of Proposition 1.
Proposition 1 (second half). A discrete distribution over the integers $\left\{0,1 \ldots, 2^{b}-1\right\}$ compiles to a BDD of size $\Theta\left(2^{b}\right)$ when represented using BITWISE_INT(Algorithm 2).

Lemma 5. Let encode $\operatorname{BitwiSE\_ INT}^{\left(\mathcal{D}_{b}\right)}=(B, P)$.
$B$ exactly matches a BDD $C$ with the following reduced structure.
For all $1 \leq i \leq b$, for all bit strings $s$ of length less than $i$, let $C$ have a node $N_{i, s}$ corresponding to decision variable $x_{s}$.

$$
\begin{aligned}
& \operatorname{low}\left(N_{i, s}\right)=\text { if } \quad|s|=i-1 \quad \text { then } \quad 0 \quad \text { else } \quad N_{i, s \prec 0} \\
& \operatorname{high}\left(N_{i, s}\right)=\text { if } \quad|s|=i-1 \quad \text { then } \quad 1 \quad \text { else } \quad N_{i, s-1}
\end{aligned}
$$

Let the roots of $C$ be placed such that $C_{i}(\vec{x})$ matches $N_{i, \varepsilon}$ given assignments to decision variables $\vec{x}$.
Any ordering of decision variables is valid as long is they are ordered by increasing bit string lengths. We arbitrarily choose to order bit strings of the same length by their lexicographical order.

As an example, the BDD for $b=2$ follows (terminal nodes visually duplicated for clarity).


Proof. Let $S_{i, j}(\vec{x})$ denote the state of $C_{i}$ at the level of the first decision variable whose bit string is of length $j$. See the following BDD and corresponding states as an example.

$$
\begin{aligned}
S_{2,0}\left(x_{\varepsilon}, x_{0}, x_{1}\right) & =N_{2, \varepsilon} \\
S_{2,1}\left(0, x_{0}, x_{1}\right) & =N_{2,0} \\
S_{2,1}\left(1, x_{0}, x_{1}\right) & =N_{2,1} \\
S_{2,2}\left(0, x_{0}, x_{1}\right) & =x_{0} \\
S_{2,2}\left(1, x_{0}, x_{1}\right) & =x_{1}
\end{aligned}
$$



Let $P(t)$ be the statement, $S_{i, t}(\vec{x})=$ if $\quad t \geq i \quad$ then $\quad B_{i}(\vec{x}) \quad$ else $\quad N_{i, B_{1}(\vec{x}) \frown \ldots \frown B_{t}(\vec{x})}$.
$P(0)$ holds as there is no bit index greater than or equal to 0 and $S_{i, 0}(\vec{x})=N_{i, \varepsilon}$ by the placement of the roots.
Assume $P(t)$, which specifies $S_{i, t}$. Consider advancing to the next node in the BDD based on the assignment to decision variables (which does nothing if we are already at a terminal node):

```
\(\operatorname{next}\left(\mathrm{S}_{\mathrm{i}, \mathrm{t}}\right)=\mathrm{if} \quad t \geq i\)
    then \(\quad B_{i}(\vec{x})\)
    else \(\quad\left(\right.\) if \(\quad x_{B_{1}(\vec{x}) \subset \ldots \frown B_{t}(\vec{x})}\) then \(\operatorname{high}\left(N_{i, B_{1}(\vec{x}) \subset \ldots \frown B_{t}(\vec{x})}\right) \quad\) else \(\left.\quad \operatorname{low}\left(N_{i, B_{1}(\vec{x}) \frown \ldots \frown B_{t}(\vec{x})}\right)\right)\)
```

We replace $x_{B_{1}(\vec{x}) \frown \ldots \frown B_{t}(\vec{x})}$ with $B_{t+1}(\vec{x})$.

$$
\operatorname{next}\left(\mathrm{S}_{\mathrm{i}, \mathrm{t}}\right)=\text { if } \quad t \geq i
$$

$$
\text { then } \quad B_{i}(\vec{x})
$$

$$
\text { else } \quad\left(\text { if } \quad B_{t+1}(\vec{x}) \text { then } \operatorname{high}\left(N_{i, B_{1}(\vec{x}) \frown \ldots \frown B_{t}(\vec{x})}\right) \quad \text { else } \quad \operatorname{low}\left(N_{i, B_{1}(\vec{x})} \frown \ldots \frown B_{t}(\vec{x})\right)\right)
$$

We replace the high and low edges by the definition:

$$
\begin{aligned}
\operatorname{next}\left(\mathrm{S}_{\mathrm{i}, \mathrm{t}}\right)= & \text { if } t \geq i \\
& \text { then } B_{i}(\vec{x}) \\
& \text { else }(\quad
\end{aligned} \begin{aligned}
& \text { if } B_{t+1}(\vec{x}) \\
& \text { then } \quad(\text { if } t=i-1 \\
& \\
& \\
& \text { else } \quad \text { (if } \quad t=i-1 \quad \text { then } \quad 1 \\
&
\end{aligned}
$$

We rearrange the if conditions:

$$
\text { then } \left.\quad \text { (if } \quad B_{t+1}(\vec{x}) \text { then } 1 \text { else } 0\right)
$$

$$
\text { else } \quad\left(\text { if } \quad B_{t+1}(\vec{x}) \text { then } \quad N_{i, B_{1}(\vec{x}) \frown \ldots \frown B_{t}(\vec{x}) \frown 1} \text { else } \quad N_{\left.\left.i, B_{1}(\vec{x}) \frown \ldots \frown B_{t}(\vec{x}) \frown 0\right)\right)}\right.
$$

The first two then branches collapse:

$$
\begin{aligned}
\operatorname{next}\left(\mathrm{S}_{\mathrm{i}, \mathrm{t}}\right)= & \text { if } \quad t+1 \geq i \\
& \text { then } \quad B_{i}(\vec{x}) \\
& \text { else } \quad N_{i, B_{1}(\vec{x}) \frown \ldots \frown B_{t+1}(\vec{x})}
\end{aligned}
$$

$$
\begin{aligned}
& \operatorname{next}\left(\mathrm{S}_{\mathrm{i}, \mathrm{t}}\right)=\text { if } \quad t \geq i \\
& \text { then } \quad B_{i}(\vec{x}) \\
& \text { else } \quad(\quad \text { if } \quad t=i-1 \\
& \operatorname{next}\left(\mathrm{S}_{\mathrm{i}, \mathrm{t}}\right)=\mathrm{if} \quad t \geq i \\
& \text { then } \quad B_{i}(\vec{x}) \\
& \text { else } \quad \text { if } \quad t=i-1 \\
& \text { then } \quad B_{t+1}(\vec{x}) \\
& \text { else } \left.\quad N_{i, B_{1}(\vec{x}) \frown \ldots} \frown_{B_{t+1}(\vec{x})}\right)
\end{aligned}
$$

As the only node we now reach has bit string length $t+1$, next $\left(S_{i, t}\right)=S_{i, t+1}$. Therefore $P(t) \rightarrow P(t+1)$.
Consider $P(b-1): S_{i, b-1}=$ if $\quad b-1 \geq i \quad$ then $\quad B_{i}(\vec{x}) \quad$ else $\quad N_{i, B_{1}(\vec{x}) \subset \ldots-B_{b-1}(\vec{x})}$.
For all $i \in\{1, \ldots, b-1\}$, we see that the root labeled $C_{i}(\vec{x})$ reaches the value $B_{i}(\vec{x})$. For $i=b$, the
 (if $\quad x_{b_{1}}^{-\ldots} \prec_{b_{b-1}}$ then 1 else 0$)=B_{b}(\vec{x})$. Therefore $B$ and $C$ are equivalent.

## Proposition 1 (second half).

Proof. Let encode $\operatorname{Bitwise\_ int~}^{\left(\mathcal{D}_{b}\right)}=(B, \vec{p})$. $B$ has $2^{b+1}-b-2$ decision nodes.
It directly follows from Lemma 5 that $B$ requires the following number of nodes.

$$
\sum_{i=1}^{b}(\# \text { of bit strings of length less than } i)=2^{n+1}-n-2
$$

## D. 6 ENCODING UNIFORM INTEGER DISTRIBUTIONS

Let $U_{n}$ denote the discrete uniform distribution over the integers $\{0,1, \ldots, n-1\}$

$$
U_{n}(i)= \begin{cases}\frac{1}{n} & i \in 0,1, \ldots, n-1 \\ 0 & \text { otherwise }\end{cases}
$$

Lemma 6. For all non-negative integers $b$ and $i$ such that $0<i<b, x_{i} \in\{0,1\}$

$$
U_{2^{b}}\left(<x_{b}, x_{b-1}, \ldots, x_{1}>\right)=\frac{1}{2} \cdot U_{2^{b-1}}\left(<x_{b-1}, \ldots, x_{1}>\right)
$$

Proof. Let $a=<x_{b}, x_{b-1}, \ldots, x_{1}>$, then

$$
U_{2^{b}}(a)=\frac{1}{2^{b}}=\frac{1}{2} \cdot \frac{1}{2^{b-1}}=\frac{1}{2} \cdot U_{2^{b-1}}\left(a-x_{b} 2^{b-1}\right)=\frac{1}{2} \cdot U_{2^{b-1}}\left(<x_{b-1}, \ldots, x_{1}>\right)
$$

Lemma 7. For integers $b$ and $i$ such that $0<i<b, x_{i} \in\{0,1\}$ and $b>1$

$$
[\operatorname{UNIFORM}(b)]\left(x_{b}, x_{b-1}, \ldots, x_{1}\right)=\frac{1}{2}[\operatorname{UNIFORM}(b-1)]\left(x_{b-1}, \ldots, x_{1}\right)
$$

Proof. Consider the case when $x_{b}=0$. The proof for the other case when $x_{b}=1$ would go in a similar fashion. Let $\pi$ be the set of executions of $\operatorname{UNIFORM}(b)$ resulting in $\left[0, x_{b-1}, \ldots, x_{1}\right]$. Let $\pi^{\prime}$ be the set of executions of $\operatorname{UNIFORM}(b-1)$ resulting in $\left[x_{b-1}, \ldots, x_{1}\right]$. Since $b>1$ is false, $X_{b}$ gets assigned flip ${ }_{1}$ evaluating to false; we return $0:: \operatorname{UNIFORM}(b-1)$.

$$
\begin{aligned}
& {[\operatorname{UNIFORM}(b)]\left(0, x_{b-1}, \ldots, x_{1}\right)} \\
& =\sum_{p \in \pi} w(p)=\frac{1}{2} \sum_{p \in \pi^{\prime}} w(p)=\frac{1}{2}[\operatorname{UNIFORM}(b-1)]\left(x_{b-1}, \ldots, x_{1}\right)
\end{aligned}
$$

Proposition 2. $\forall b>0,[\operatorname{UNIFORM}(b)]=U_{2^{b}}$

Proof. The proof is by induction on $b$
Base Case: For $b=1$,

$$
U_{2}(a)= \begin{cases}\frac{1}{2} & a=0 \\ \frac{1}{2} & a=1 \\ 0 & \text { otherwise }\end{cases}
$$

Executing UNIFORM(1) would result in $X_{1}=$ flip $\frac{1}{2}$. So, $\operatorname{Pr}\left(X_{1}=1\right)=\frac{1}{2}$ and $\operatorname{Pr}\left(X_{1}=0\right)=\frac{1}{2}$
Induction Hypothesis: $\forall b^{\prime}<b\left[U N I F O R M\left(b^{\prime}\right)\right]=U_{2^{b^{\prime}}}$
Inductive Step: To prove $[U N I F O R M(b)]=U_{2^{b}}$
Let the output of $\operatorname{UNIFORM}(b)$ be $a$ where $a_{b}=<x_{b}, x_{b-1}, \ldots, x_{1}>$
Case 1: $\forall 0 \leq a<2^{b-1}$
Since $a<2^{b-1}, x_{b}=0$
$[\mathrm{UNIFORM}(b)]\left(0, x_{b-1}, \ldots, x_{1}\right)$
$=\frac{1}{2}[\operatorname{UNIFORM}(b-1)]\left(x_{b-1}, \ldots, x_{1}\right)=\frac{1}{2} \cdot U_{2^{b-1}}\left(<x_{b-1}, \ldots, x_{1}>\right)$
(by Lemma 7 and I.H.)
$=U_{2^{b}}\left(<0, x_{b-1}, \ldots, x_{1}>\right)$
(by Lemma 6)

Case 2: $\forall 2^{b-1} \leq a<2^{b}$
$\overline{\text { Since } a} \geq 2^{b-1}, x_{b}=1$, the proof is similar to that for Case 1.

## E BETA PRIOR APPLICATION: BAYESIAN NETWORK PARAMETER LEARNING

One natural application of a Beta prior is as a prior distribution for learning the parameters of a (binary) Bayesian network. The task of Bayesian network parameter learning can be described as follows: given a set of data consisting of instantiations of network variables, we want to find the network parameters maximizing the probability of this data. One interesting case is when our data is incomplete; that is, there are some variables not given a value. In this setting, a Bayesian approach to parameter learning must consider all (exponentially many) possible instantiations of these missing values [Darwiche, 2009.

Using our Beta prior implementation described in the main paper, this setting can naturally be modeled within Dice.jl.. A Bayesian network can be expressed in the probabilistic program with Beta priors on each network parameter; a dataset can then be observed. By returning the distribution over our Beta parameters $\alpha$ and $\beta$, we obtain our posterior, a mixture over Beta distributions. These can then be manually combined to obtain the exact posterior density function.
We provide a brief example of this on the survey Bayesian network The structure of this network is shown below; we simplify the network by making all variables binary so that the Beta a suitable prior. We note that we can actually generalize to the non-binary case with a Dirichlet prior using an approach similar to that used for the Beta.


Suppose we want to learn parameters from the dataset given below; the entry ? indicates a missing value. We focus on the specific parameter $\theta_{o \mid e}=\operatorname{Pr}(O=1 \mid E=1)$, which we give a uniform prior $\operatorname{Beta}(1,1)$.

[^0]| A | S | E | O | R | T |
| :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | $?$ | $?$ | 1 | 1 | 1 |
| 1 | 0 | $?$ | 1 | 0 | 1 |
| 1 | 0 | 1 | $?$ | 0 | 1 |
| 1 | $?$ | 1 | 0 | $?$ | 1 |
| 0 | 0 | 1 | 1 | $?$ | 1 |
| 0 | 1 | $?$ | 1 | 1 | $?$ |
| 1 | $?$ | 0 | 0 | 1 | 0 |
| 0 | $?$ | 1 | $?$ | $?$ | $?$ |
| 1 | 1 | 1 | $?$ | 1 | $?$ |
| 1 | 0 | 0 | $?$ | 1 | 1 |

By running the program representing this task, we get the following large output distribution over Beta parameters - a mixture over Beta distributions. Note that it does not contain as many entries as possible instantiations, as some complete variable instantiations result in the same posterior Beta.

| $\alpha$ | $\beta$ | $\operatorname{Pr}()$. |
| :---: | :---: | :---: |
| 9 | 3 | 0.226 |
| 8 | 4 | 0.207 |
| 10 | 2 | 0.177 |
| 7 | 5 | 0.160 |
| 6 | 6 | 0.109 |
| 5 | 7 | 0.066 |
| 4 | 8 | 0.035 |
| 3 | 9 | 0.016 |
| 2 | 10 | 0.005 |

If we plot the corresponding mixture of Betas, we can get the following posterior PDF - note that this is an exact posterior recovered from the Beta parameters, in constrast to the approximate result one would get from sampling-based inference methods.

The code used in this example is available in the same Dice.jl repository (https://github.com/Juice-jl/ Dice.jl/tree/arithmetic).


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[^0]:    ${ }^{1}$ https://www.bnlearn.com/bnrepository/

